# On the Development of Block-Ciphers and Pseudo-Random Function Generators Using the Composition and XOR Operators 

## by

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A thesis submitted in conformity with the requirements for the degree of Master's Of Science Graduate Department of Computer Science University of Toronto

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# Abstract <br> On the Development of Block-Ciphers and Pseudo-Random Function Generators Using the Composition and XOR Operators 

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We attempt to provide evidence for the security of block-ciphers which are constructed by taking the composition and exclusive-or of non-secure function generators. We provide this evidence by showing that such construction can be used to combine partially secure pseudo-random function generators into generators with stronger security properties than any of their constituents. We extend results of Luby and Rackoff, and show that there are constructions based on the composition and $\square$ operators which can be used to combine $1-\delta$ secure pseudo-random permutation and function generators, where $0<\delta<1$, to achieve a $1-\frac{1}{\log ^{c} n}$ secure generator, for any $c>0$. We then give the first proof that there is a natural construction which combines together $1-\delta$ secure function generators to form a pseudo-random function generator.

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## Chapter 1

## Introduction

The field of cryptography has a long and rich history, yet it is only in the past few decades that work in the field has become more of a science then an art. Previously, cryptographic systems were developed based on heuristics, and the cleverness of their designers. However, the designers gave no evidence that the techniques or operators they used in the development of their systems were actually increasing security.

Two operators which have consistently been used in the development of block ciphers are the composition and exclusive-or (XOR) operators, for it seems to have been taken for granted in the cryptographic community that these operators increase security. This belief seems to stem from empirical evidence that cryptographic systems based on these operators have remained fairly secure, under the scrutiny of very intense and extensive cryptanalysis. One such example is the Data Encryption Standard (DES), in which 16 permutations generators, which are "completely insecure" individually, are composed and combined together with the composition and XOR operators, to give what appears to be a remarkably secure block cipher.

Since the development of DES, in the early 1970s, it has undergone extensive public cryptanalysis, and during the past 30 years only two forms of cryptanalysis have suggested attacks which require significantly fewer than the $2^{55}$ calls to the DES algorithm
which are necessary to break it by a brute-force attack on its 56 bit key. The two forms of cryptanalysis which have performed reasonably well are linear, and differential cryptanalysis developed by M. Matsui in [16]; and E. Biham and A. Shamir in [3], respectively. Using the techniques of linear cryptanalysis, $2^{47}$ executions of DES on known plain texts are required to break it; where as in differential cryptanalysis, $2^{47}$ executions on chosen plain texts are required. In either case, the attacks are not significantly stronger than a brute force attack on the keys of DES, and the empirical evidence seems to suggest that DES is a very secure cryptographic system.

Further evidence of the pervasiveness of the use of composition and XOR as operators for augmenting the security of block ciphers is the fact that MARS, TWOFISH, RC6 and most of the other Advanced Encryption Standard (AES) candidate ciphers seem to rely on these operators to increase their security. All of the candidate ciphers are currently in a competition to become the Advanced Encryption Standard. The AES will replace DES as the standard block cipher used by the U.S. Government, and therefore the candidate algorithms are expected to be very secure, given the current state of knowledge in cryptology.

The purpose of this thesis is to provide evidence which supports the use of composition and XOR operators in the development of block ciphers. We will not be able to demonstrate that their use in the development of a cryptographic system can result in a provably secure cryptographic system, as such a proof would imply $\mathcal{P} \neq \mathcal{N} \mathcal{P}$, and thus solve one of the most important open questions in computer science. Instead, we have much more modest goals, and will demonstrate that there are constructions based on composition and XOR which combine function and permutation generators which are already partially secure, and which result in function or permutation generators which are more secure. This will be used to give evidence to suggest that composition improves security, and partially justifies their use in the design of block ciphers.

### 1.1 Overview of the Thesis

We will now provide a brief outline of the thesis. We begin, in Chapter 2, by introducing some standard definitions of cryptographic primitives. We also explain the relationships that exist between these primitives. Then in Chapter 3 we look at some previous work which has been done on constructions which increase the security of different types of cryptographic primitives. We will look at work which has been done in both complexity and information theoretic models.

In Chapter 4, we will then develop a list of possible constructions which use composition and XOR operators to combine partially secure function generators into another function or permutation generator. In the model we will work in, our definition of security will be based on complexity theoretic notions. We will show that most of our proposed constructions result in a function or permutation generator which is at best as secure as its constituent generators. Therefore these constructions are not useful for increasing security in the construction of block ciphers, but based on some work done by Luby and Rackoff in [12], in Chapters 5, 6 and 7 we will prove that three of the constructions result in generators which have stronger security than any of their constituent generators.

The first security increasing construction is based on the composition operator. Previously in [12], Luby and Rackoff have shown that composing partially secure permutation generators a constant number of times results in generator with stronger security properties than any of the constituents. By making some modifications to their argument, by making it less dependent on sampling different distributions of functions, we will extend their result to show that we can compose a $\mathcal{O}(\log \log n)$ number of partially secure permutation generators. This will result in a construction which is more secure than the constructions which are provably secure from the initial result by Luby and Rackoff. These results are presented in Chapter 5.

The second security increasing construction is based on the $\square$ operator, proposed by Luby and Rackoff in [13], which is a combination of the composition and XOR operators.

Given two functions $f_{1}$ and $f_{2}$ such that $f_{i}:\{0,1\}^{n} \rightarrow\{0,1\}^{n}$, for $i \in\{1,2\}$, define the operator $\square$ (read as box) as:

$$
\left(f_{1} \square f_{2}\right)(x)=f_{1}\left(f_{1}(x) \oplus f_{2}(x)\right) \oplus f_{2}\left(f_{1}(x) \oplus f_{2}(x)\right)
$$

By proving two technical arguments about this operator, we will show that we can use the arguments previously presented for the composition operator acting on permutation generators, and apply them to the case of the operator acting on function generators. Therefore, it will follow that we can combine $\mathcal{O}(\log \log n)$ function generators using the $\square$ operator, and achieve the same security as was achieved by composing together permutation generators. These results are presented in Chapter 6.

The third security increasing construction is based on the $\diamond$ operator generator. We define the $\diamond$ operator generator (read as Diamond) as $\delta=\left\{\diamond_{r_{1}, r_{2}}^{n} \mid n \in \mathbb{N} \wedge r_{1}, r_{2} \in\right.$ $\left.\{0,1\}^{n}\right\}$. Let $f_{1}$ and $f_{2}$ be two functions such that $f_{i}:\{0,1\}^{n} \rightarrow\{0,1\}^{n}$, for $i \in\{1,2\}$. For each $r_{1}, r_{2} \in\{0,1\}^{n}$ we define the operator $\nabla_{r_{1} \bullet r_{2}}^{n}$, which acts on the functions $f_{1}$ and $f_{2}$ as:

$$
\left(f_{1} \diamond_{r_{1} \bullet r_{2}}^{n} f_{2}\right)(x)=f_{1}\left(x \oplus r_{1}\right) \oplus f_{2}\left(x \oplus r_{2}\right)
$$

We will show that by combining function generators using randomly chosen operators from the $\diamond$ operator generator, we can combine any polynomial number of function generators, and the result will be a completely secure ${ }^{1}$ function generator. This proof will be based on the proof by Luby and Rackoff, but we will observe that this construction allows us to short-circuit a key section of the proof. This in turn is what allows us to combine together many more function generators than in the previous cases. We will then observe that while the third construction improves security significantly, there is no reason to suspect that it will be useful in the design of block ciphers. These results are presented in Chapter 7.

Finally, in Chapter 8 we will suggest future research directions and goals which are

[^1]similar but more ambitious than those presented in this document. In particular, we suggest the development of a theory for the development of block-ciphers. In this theory we would conjecture that completely insecure function generators could be combined in different constructions -such as the ones mentioned above- to form secure generators. We would say that a construction was secure if the initial, insecure generators, and the constructed generator had certain combinatorial properties. The motivation for such a theory is that it would allow for the comparison of different resource and design trade-offs which are made during the development of different ciphers.

### 1.2 Notation \& Terminology

Below we introduce some notation and terminology which will be used throughout the rest of the thesis.

Terminology 1.1 For $\mu, \nu \in\{0,1\}^{*}$, let $\mu \bullet \nu$ denote their concatenation.
Terminology 1.2 Let $\mathcal{P}$ denote the set of all permutations, and let $\mathcal{P}^{n}$ denote the set of all permutations $\sigma:\{0,1\}^{n} \rightarrow\{0,1\}^{n}$.

Terminology 1.3 Let $\mathcal{F}$ denote the set off all functions, $\mathcal{F}^{\boldsymbol{t} \rightarrow p}$ denote the set of all functions $f:\{0,1\}^{l} \rightarrow\{0,1\}^{p}$, and let $\mathcal{F}^{n}$ be the set $\mathcal{F}^{n \rightarrow n}$.

Terminology 1.4 For $\alpha, \beta \in\{0,1\}^{n}$, let $\alpha \oplus \beta$ denote the bit-by-bit exclusive-or of $\alpha$ and $\beta$. For $f, g \in \mathcal{F}^{n}$, let $(f \oplus g)(\alpha)$ denote $f(\alpha) \oplus g(\alpha)$.

Terminology 1.5 For any two functions $g$ and $h$, let $g \circ h(x)=g(h(x))$. When we refer to the composition $f \circ g(x)=f(g(x))$ of two functions $f(x)$ and $g(x)$, we refer to $f$ as the outer function, and $g$ as the inner function.

Terminology 1.6 For any set $A$, let $x \in A$ be the action of uniformly at random choosing an element $x$ from $A$. It will be clear from context when $\in$ is used to refer to an element in a set, and when it refers to choosing uniformly from a set.

Terminology 1.7 Let $\mathcal{D}_{1}, \mathcal{D}_{2}, \ldots$ be a sequence of distributions, and let e represent a series of events $e_{1}, e_{2}, \ldots$. such that for all $i, e_{i}$ is an event of $D_{i}$. We say that e occurs with significant probability if for some constant c>0 and for infinitely many $n$ the $\operatorname{Pr}_{\mathcal{D}_{n}} e_{n} \geq \frac{1}{n^{c}}$. We say that an event $e$ occurs with negligible probability if, for all constants $c>0$ and for all sufficiently large $n, \operatorname{Pr}_{D_{n}} e_{n}<\frac{1}{n^{c}}$.

Terminology 1.8 When we refer to a function poly $(n)$ we are referring to any function in the set $\bigcup_{i=1}^{\infty} \mathcal{O}\left(n^{i}\right)$.

Terminology 1.9 When we refer to a function $p o l y-l o g(n)$ we are referring to any function in the set $\bigcup_{i=1}^{\infty} \mathcal{O}\left(\log ^{i} n\right)$.

Terminology 1.10 Let $A$ be an algorithm or Turing Machine, and let $\pi$ be an orcale. We denote by $A^{\pi}$ the algorithm or Turing Machine $A$, with access to the oracle $\pi$.

Terminology 1.11 We call a circuit $C$ a probabilistic circuit, if it requires as inputs a string of random bits.

Terminology 1.12 Let $C$ be a circuit whose outputs are in the range $\{0,1\}$. Then we say $C$ is a decision circuit. Let $x$ be an input to $C$. Then we say $C$ accepts $x$ if $C(x)=1$, and we say that $C$ rejects $x$ if $C(x)=0$.

Terminology 1.13 Let $\mathcal{D}$ be a distribution over the inputs of a decision circuit $C$. Then we define

$$
\operatorname{Pr}_{C}(\mathcal{D})=\operatorname{Pr}_{d \in \mathcal{D}}[C(d)=1]
$$

and we say that $C$ accepts $a$ fraction $\operatorname{Pr}_{C}(\mathcal{D})$ of its inputs, and rejects $a$ fraction 1 $\operatorname{Pr}_{C}(\mathcal{D})$ of its inputs.

Terminology 1.14 Let $\alpha \in\left\{0, \ldots, 2^{n}-1\right\}$, then let $\bar{\alpha}$ be the bit-wise binary representation of $\alpha$ in $\{0,1\}^{n}$.

Terminology 1.15 Given two pairs $(a, b)$ and $(c, d)$, where $a, b, c, d \in \mathbb{N}$, we say that $(a, b)<(c, d)$ iff $a<c$ or $a=c$ and $b<d$.

Terminology 1.16 We call $G:\{0,1\}^{\kappa} \times\{0,1\}^{n} \rightarrow\{0,1\}^{m}$ a function generator instance. We say that $k \in\{0,1\}^{\kappa}$ is a key of $G$, and we write $G(k, \cdot)$ as $g_{k}(\cdot)$, and say that key $k$ chooses the function $g_{k}$. Let $g \in G$ represent the act of uniformly at random choosing a key $k$ from $\{0,1\}^{\kappa}$, and then using the key $k$ to choose the function $g_{k}$.

Let $m$ and $\ell$ be polynomials, and let $\mathcal{N} \subset \mathbb{N}$ be an infinitely large set. For each $n \in \mathcal{N}$, let $G^{n}:\{0,1\}^{\ell(n)} \times\{0,1\}^{n} \rightarrow\{0,1\}^{m(n)}$ be a function generator. We call $G=\left\{G^{n} \mid n \in \mathcal{N}\right\}$ a function generator.

In an abuse of notation, we will often refer to both specific function generator instances and function generators as function generators. We hope it will be clear from the context which term is actually being referred to.

## Chapter 2

## Cryptographic Primitives

Cryptographers have developed a series of conjectured primitives which can be used in the development of cryptographic protocols. Due to a lack of progress in complexity theory, no one has been able to prove the existence of these primitives. In fact, as will be shown in the sequel, such a proof would imply that $\mathcal{P} \neq \mathcal{N} \mathcal{P}$, and would close the largest open question in computer science. In this chapter we present the definitions for a number of cryptographic primitives and some theorems which represent the relationships between them. However, in the definition of every cryptographic primitive there is the notion of an adversary. Therefore, we begin the chapter with a discussion of the two standard adversarial models, and the relationships that exist between them.

### 2.1 Uniform vs. Non-Uniform Adversaries

In the definition of each cryptographic primitive there exists the notion of an adversary. Abstractly, its purpose is to break an effect that a primitive is trying to achieve. Resource bounds are imposed on the adversaries, so that they model the computational power "real world" adversaries might feasibly have access to, were they trying to break existential instances of primitives. There are two standard computational models which are used to define resource bounded adversaries.

## Uniform Adversaries

A uniform adversary is a Turing Machine which runs in time bounded by a polynomial in the size of its input. In practice we know of algorithms which run more quickly using randomization then the best known non-randomized versions, so we allow the Turing machine to be probabilistic and have access to an infinite string of random bits. In order to model the adversaries of certain primitives, we allow that the Turing Machines have access to one or more oracles, which answer queries. Unless otherwise stated we assume that the oracles respond to a query in unit time.

## Non-Uniform Adversary

A non-uniform adversary is an infinite set of circuits $\left\{C_{1}, C_{2}, \ldots\right\}$, where circuit $C_{i}$ is used on inputs of size $i$. We wish to model efficient computation on part of the adversary, so we assume that the size of each circuit $C_{i}$ is bounded by $p(i)$, for some polynomial $p$. The size of a circuit is defined to be the number of gates, and the number of connections between gates in the circuit.

Unlike the uniform adversary we do not need to allow our circuits to be probabilistic. Given a probabilistic circuit which takes a string of random bits as part of its input, we can use the non-uniformity of the circuits to fix a specific set of bits as the random-input of the circuit. Normally, a simple averaging argument will show that such a string exists. We refer the interested reader to [7] for a further discussion of this topic.

In correspondence with the oracles in the uniform model, we assume that circuits can have oracle-gates, which respond to inputs in the same manner an oracle would. We assume that the gates are of unit size, and are otherwise treated like any other gate. We stress that the description of the circuit family need not be efficiently computable even though each circuit is of small size, relative to the size of its input.

Terminology 2.1 We write $C(f)$ to represent a circuit $C$ which has access to the oracle $f$.

Notice that in an abuse of notation, circuit inputs and circuit oracles are represented in the same manner. This is done because in this thesis, in most cases of interest, our circuits will not have inputs in the standard sense. Rather, they will be constructed to differentiate between different oracles, and therefore, in some sense, we are considering the oracle as an input. Notice that when we treat oracles as inputs to circuits we consider the size of the oracle query to be the size of input of the circuit. We hope that this abuse of notation will not cause confusion, and that we have made it clear from context whether or not the input to a circuit is an oracle or a traditional input bit string.

Terminology 2.2 Let $\mathcal{D}$ be a distribution over oracles of an oracle-decision-circuit $C$. Then we define

$$
\operatorname{Pr}_{C}(\mathcal{D})=\operatorname{Pr}_{d \in \mathcal{D}}[C(d)=1]
$$

and we say that $C$ accepts $a$ fraction $\operatorname{Pr}_{C}(\mathcal{D})$ of its inputs, and rejects $a$ fraction $1-$ $\operatorname{Pr}_{C}(\mathcal{D})$ of its inputs.

In the sequel we present the primitives in the uniform model, but note that it is easy to determine the corresponding definitions for the non-uniform model.

### 2.2 Primitives

The first primitive we consider is the pseudo-random number generator, which takes a random sequence of bits and extends them into a much longer sequence of random looking bits.

[^2]$$
\left|\operatorname{Pr}_{X \in\{0,1\}^{n}}\left[A\left(G^{n}(X)\right)=1\right]-\operatorname{Pr}_{Y \in\{0,1\}^{(n)}}[A(Y)=1]\right|<\frac{1}{q(n)} .
$$

Theorem 2.1 If there exists a PRNG $G$, then $\mathcal{P} \neq \mathcal{N} \mathcal{P}$.

Proof: Let $L=\left\{y \mid \exists n, \exists x \in\{0,1\}^{n}\right.$ s.t. $\left.G^{n}(x)=y\right\}$. Observe, there exists a nondeterministic machine $M$ which decides $L$; given an input $y$, of size $p(n), M$ guesses a string $x \in\{0,1\}^{n}$ and accepts if $G^{n}(x)=y$. However, by the definition of $G$, there exists no polynomial time bounded machine which decides $L$, and therefore $\mathcal{P} \neq \mathcal{N} \mathcal{P}$.

An other natural primitive to consider is the pseudo-random function generator, which given a random string of bits generates a random looking function $f$.

Definition 2.2 (Pseudo-Random Function Generator - PRFG) Let $m$ and $\ell$ be polynomials. For each $n$ let $G^{n}:\{0,1\}^{\ell(n)} \times\{0,1\}^{n} \rightarrow\{0,1\}^{m(n)}$ be a function generator instance, computable in time bounded by a polynomial in $n$. Define $G=\left\{G^{n} \mid n \in \mathbb{N}\right\}$ to be a function generator. For $k \in\{0,1\}^{\ell(n)}$, let $G^{n}(k, \cdot)$ be written as $g_{k}^{n}(\cdot)$ or $g_{k}(\cdot)$ when the value of $n$ is clear, and referred to as the function $g_{k}^{n}$ or $g_{k}$ respectively. If $A^{h}$ is an adversary which queries a function $h$, then we say it is capable of $\epsilon$ distinguishing $g$ from random if, for $\epsilon: \mathbb{N} \rightarrow[0,1]$, some polynomial $p$ and infinitely many $n$ :

$$
\left|\operatorname{Pr}_{k \in\{0,1\}^{n}}\left[A^{g_{k}^{n}}=1\right]-\operatorname{Pr}_{f \in \mathcal{F}(n) \rightarrow m(n)}\left[A^{f}=1\right]\right| \geq \epsilon(n)+\frac{1}{p(n)}
$$

We say that $G$ is $(1-\epsilon(n))$ secure if there exists no function querying adversary $A$, bound to run in time polynomial in $n$, which can $\epsilon$ distinguish $G$ from random.

We say that $G$ is a pseudo-random function generator if it is 1 secure.

If $G$ is a $(1-\epsilon(n))$ secure PRFG, and if it is the case that for all $n$, and for all $k \in\{0,1\}^{\ell(n)}$ that $g_{k}^{n}$ is a permutation, then we call $g$ a $(1-\epsilon(n))$ secure pseudo-random permutation generator. We say that $G$ is a pseudo-random permutation generator (PRPG), if $G$ is 1 -secure.

Terminology 2.3 If $G$ is a 1 -secure generator, we say it is completely secure. If $G$ is a 0 -secure generator, we say it is insecure, and if $G$ is neither completely secure nor insecure, then we say that it is partially secure.

It was shown by Goldreich et. al., in [5], that the existence of a PRNG is a necessary and sufficient condition to ensure the existence of a PRFG.

## Theorem 2.2 PRNGs exist iff PRFGs exist.

We will demonstrate the constructions, and refer the reader to the original paper for the proof of correctness.

Construction: To construct the PRFG, we consider a PRNG $G=\left\{g^{n}:\{0,1\}^{n} \rightarrow\right.$ $\left.\{0,1\}^{2 n} \mid n \in \mathbb{N}\right\}$. For any $x \in\{0,1\}^{n}$ we will consider $g^{n}(x)$ as $g(x)_{0}^{n} \bullet g(x)_{1}^{n}$, where $\left|g(x)_{0}\right|=\left|g(x)_{1}\right|=n$. We will construct a PRFG, $F=\left\{F^{n}:\{0,1\}^{n} \times\{0,1\}^{n} \rightarrow\right.$ $\left.\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$. For each key $x \in\{0,1\}^{n}$ will define the function $F^{n}(x, \cdot)=f_{x}^{n}$.

In order to compute $f_{x}^{n}(y)$, for $y \in\{0,1\}^{n}$, we define a complete binary tree of depth $n$. We label the root with the empty string $\lambda$, and assign it the value $x$. We then recursively apply the following rule to assign a label and value to all other nodes, $v$, in the tree. Let $L_{p_{v}}$ and $V_{p_{v}}$ be the respective label and value of $v$ 's parent.

- If $v$ is a left child set $i=0$, otherwise set $i=1$.
- Let the label for node $v$ be $i \bullet L_{p_{0}}$.
- Let the value for node $v$ be $g^{n}\left(V_{p_{v}}\right)_{i}$.

We define the value of $f_{x}^{n}(y)$ to be equal to the value of the leaf node labeled $y$.
The construction in the other direction is much simpler. Let $H=\left\{H^{n}:\{0,1\}^{n} \times\right.$ $\left.\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ be a PRFG. We construct a PRNG $G=\left\{g^{n}:\{0,1\}^{n} \rightarrow\right.$ $\left.\{0,1\}^{2 n} \mid n \in \mathbb{N}\right\}$ as follows. For each $n$ and each $x \in\{0,1\}^{n}$ let $g^{n}(x)=H_{x}^{n}(\overline{0}) \bullet H_{x}^{n}(\overline{1})$.

It was shown by Luby and Rackoff in [14] that the existence of PRFGs is a necessary and sufficient condition to ensure the existence of PRPGs. Their construction is based on a generalization of DES which uses independent keys in each of the sixteen rounds of the DES protocol, rather than using a key scheduling algorithm to generate a new key in each round. This generalization of DES is commonly referred to as MDES.

Additionally, Luby and Rackoff demonstrate how to construct a PRPG which is secure against an adversary which not only has access to an oracle which computes a permutation $\sigma$, but which also has access to an oracle which computes the inverse of the permutation $\sigma^{-1}$. We call a generator a super pseudo-random permutation generator (SPRPG) if, it is considered a PRPG under a modified version of the PRPG definition, in which each adversary $A^{\sigma}$, is replaced by a stronger adversary $A^{\sigma, \sigma^{-1}}$.

## Theorem 2.3 PRFGs exist iff PRPGs exist.

Construction: Let $F=\left\{F^{n}:\{0,1\}^{n} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ be a PRFG. We will demonstrate the Luby-Rackoff construction, which produces a PRPG $H=\left\{H^{n}\right.$ : $\left.\{0,1\}^{3 n} \times\{0,1\}^{2 n} \rightarrow\{0,1\}^{2 n} \mid n \in \mathbb{N}\right\}$. For each $n$ we perform the following construction. For any strings $x \in\{0,1\}^{2 n}$ and $k \in\{0,1\}^{n}$, let $L_{x}^{n} \bullet R_{x}^{n}=x$, where $\left|L_{x}^{n}\right|=\left|R_{x}^{n}\right|=n$ and let $g_{k}^{n}(x)=R_{x}^{n} \bullet\left(L_{x}^{n} \oplus f_{k}^{n}\left(R_{x}^{n}\right)\right)$. We will define $h_{k_{1} \odot k_{2} \bullet k_{3}}^{n}(x)=g_{k_{1}}^{n} \circ g_{k_{2}}^{n} \circ g_{k_{3}}^{n}(x)$. Further, to construct a SPRPG define $\widehat{H}=\left\{\widehat{H^{n}}:\{0,1\}^{4 n} \times\{0,1\}^{2 n} \rightarrow\{0,1\}^{2 n} \mid n \in \mathbb{N}\right\}$, we let $\hat{h}_{k_{1} \odot k_{2} \odot k_{3} \circ k_{4}}^{n}(x)=g_{k_{1}}^{n} \circ g_{k_{2}}^{n} \circ g_{k_{3}}^{n} \circ g_{k_{1}}^{n}(x)$.

Recent work by Naor and Reingold [18] has shown that these constructions can be modified to be less dependent on calls to the PRFG. This is desirable because calls to a PRFG can consume a considerable amount of resources. Their construction requires pair-wise independent permutation generators, and while these generators are not cryptographic primitives, they are used extensively in cryptography. Therefore, we will present a definition of pair-wise independent permutation generators, and then present the construction of Naor and Reingold.

Definition 2.3 (k-wise Independent Permutation Generator) Let $\ell$ be a polynomial. For each $n$ let $G^{n}:\{0,1\}^{\ell(n)} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n}$ be a function, computable in time bounded by a polynomial in $n$, where for each $n$ and each $k \in\{0,1\}^{\ell(n)}$ the function $G^{n}(k, \cdot)$ is a permutation. We write $G^{n}(k, \cdot)$ as $g_{k}^{n}(\cdot)$, and the function is referred to as the permutation $g_{k}^{n}$, or $g_{k}$ if the value of $n$ is clear from the context. We define $G^{n}=\left\{G^{n} \mid n \in \mathbb{N}\right\}$, and say that $G$ is a $k$-wise independent permutation generator, if for each $n$ and for all $x_{1}, \ldots, x_{k} \in\{0,1\}^{n}$, where $x_{i} \neq x_{j}$ for $i \neq j$; and for all $y_{1}, \ldots, y_{k} \in\{0,1\}^{n}$, where $y_{i} \neq y_{j}$ for $i \neq j$ :

$$
\operatorname{Pr}_{q \in\{0,1\}^{\ell(n)}}\left[g_{q}^{n}\left(x_{1}\right)=y_{1} \wedge \cdots \wedge g_{q}^{n}\left(x_{k}\right)=y_{k}\right]=\prod_{i=0}^{k-1} \frac{1}{2^{n}-i}
$$

If $G$ is a 2 -wise independent permutation generator then we say that it is pair-wise independent.

We now show how Naor and Reingold simplify the construction of a PRPG from a PRFG originally presented by Luby and Rackoff. Naor and Reingold show that if we consider the construction of the functions $h^{n}$ and $\hat{h}^{n}$ proposed by Luby and Rackoff, then the first permutation $g_{k_{3}}^{n}$ in $h^{n}$, and the first and last permutations, $g_{k_{1}}^{n}$ and $g_{k_{4}}^{n}$, in $\hat{h}^{n}$ can be replaced by permutations chosen from a pair-wise independent permutation generator ${ }^{1}$. Thus if $P=\left\{P^{n}:\{0,1\}^{2 n} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ is pair-wise independent permutation generator, then they replace the definition of $h_{k_{1} \circ k_{2} \circ k_{3}}^{n}(x)=g_{k_{1}}^{n} \circ g_{k_{2}}^{n} \circ g_{k_{3}}^{n}(x)$ with $h_{k_{1} \bullet k_{2} \circ \kappa_{3}}^{n}=g_{k_{1}}^{n} \circ g_{k_{2}}^{n} \circ p_{\kappa_{3}}^{n}$, where $\kappa_{3} \in\{0,1\}^{2 n}$, to construct a PRPG. Similarly, they replace the definition of $\hat{h}_{k_{1} \circ k_{2} \circ k_{3} \circ k_{4}}^{n}(x)=g_{k_{1}}^{n} \circ g_{k_{2}}^{n} \circ g_{k_{3}}^{n} \circ g_{k_{4}}^{n}(x)$ with $\hat{h}_{\kappa_{1} \circ k_{2} \circ k_{3} \circ \kappa_{4}}^{n}=$ $p_{\kappa_{1}}^{n} \circ g_{k_{2}}^{n} \circ g_{k_{3}}^{n} \circ p_{\kappa_{4}}^{n^{-1}}$, where $\kappa_{1}, \kappa_{4} \in\{0,1\}^{2 n}$, to construct a SPRPG.

Note that no construction was given to produce a PRFG from a PRPG, this is because any PRPG is a PRFG. This is due to the fact that in order to differentiate a random function $f \in \mathcal{F}^{n}$ from a random permutation $g \in \mathcal{P}^{n}$, it would be necessary

[^3]to demonstrate that $f$ is either not onto or one-to-one. It can easily be shown, by the birthday attack argument ${ }^{2}$, that the probability of finding a pair $(\alpha, \beta) \in\{0,1\}^{n \times 2}$ s.t. $f(\alpha)=f(\beta)$, with at most $\mathrm{p}(\mathrm{n})$ queries to $f$ is exponentially small, and therefore no adversary, bounded to run in polynomial time, would be able to distinguish a permutation from a function with a significant probability.

We now consider a primitive which at first glance seems unrelated to the primitives which have previously been seen.

Definition 2.4 (One-Way Functions) For each $n$, let $f^{n}:\{0,1\}^{n} \rightarrow\{0,1\}^{l(n)}$ be a function computable in time bounded by a polynomial in $n$. Define $F=\left\{f^{n} \mid n \in \mathbb{N}\right\}$. We say that an adversary is capable of $\epsilon(n)$ inverting $F$ if for some some polynomial $p$, and all sufficiently large $n$ :

$$
\left|\operatorname{Pr}_{x \in\{0,1\}^{n}}\left(f^{n}\left(A\left(f^{n}(x)\right)\right)=f^{n}(x)\right)-\frac{1}{2^{i(n)}}\right| \geq \epsilon(n)+\frac{1}{p(n)} .
$$

We say that $F$ is a $\delta(n)$-weak one-way function, if no adversary $A$ is capable of $1-\delta(n)$ inverting $F$. We say that $F$ is a one-way function if it is 1-weak. Finally, if $F$ is a permutation, then it may be referred to as a $\delta(n)$-weak one-way permutation, or oneway permutation respectively.

Intuitively, a $\epsilon(n)$-weak one-way function is a function which is easy to compute, but which is hard to invert on a $1-\epsilon(n)$ fraction of its range. Similarly, a one-way function is easy to compute, but hard to invert. It may seem that assuming the existence of a weak one-way function is a weaker assumption than assuming the existence of a one-way function. This however is not the case, given a weak one-way function $f$, it is possible to construct a one-way function $g$.

Theorem 2.4 There exists a $\delta(n)$-weak one-way function iff there exits a one-way function.

[^4]We refer the interested reader to $[7,15]$ for a proof of the above theorem.
Although one-way functions appear to have little in common with the pseudo-random primitives, it can be shown that the existence of one-way functions is a necessary and sufficient condition for the existence of PRNGs.

## Theorem 2.5 There exist PRNGs iff there exist one-way functions.

Although producing a one-way function from a PRNG is a relatively trivial task, producing a PRNG from a one-way function is much more difficult. We refer the interested reader to $[7,15]$ for a proof of the above theorem.

## Chapter 3

## Related Work

In this thesis we are trying to show that there are constructions based on composition and XOR which can be used to take function and permutation generators with relatively weak security properties and use them to construct function and permutation generators with strengthened security properties. In this chapter we will review some of the work which has been done which has a relation to constructions based on XOR and composition, and which increase the security of weak primitives.

### 3.1 Complexity Theoretic Model

We will first consider the work which is based in complexity theory. This means that the security definitions in this section will rely on the fact that our adversaries are computationally limited in performing their tasks.

### 3.1.1 Yao's XOR Lemma

If we are given a predicate which is weakly unpredictable in some resource bounded computational model; then a natural question to ask is whether it is possible to construct, by some "natural" method, a new predicate which amplifies the unpredictability of the
original function. Ultimately, the constructed predicate should be intractable to predict, better than at random, under a resource bounded computational model which is similar -if not identical- to the one given for the weakly unpredictable predicate. In [21], Yao postulated that the answer to this question was positive and suggested a solution that is both simple and intuitive: if $p(x)$ is weakly unpredictable then surely $P\left(x_{1}, \ldots, x_{t}\right)=$ $\oplus_{i=1}^{t} p\left(x_{i}\right)$ is almost completely unpredictable for large enough $t$. Yao did not present a proof, and the first proof presented was by Levin in [11]. Since then alternative proofs have appeared from Impagliazzo [8] and from Goldreich, Nisan and Wigderson [6]. In addition to giving their own proof of the lemma, Goldreich, Nisan and Wigderson [6] provide a survey of Levin and Impagliazzo's results.

Before the XOR Lemma there had been many proofs which showed that certain constructions maintained the security of their components. However, the true importance of this result is that it was the ground breaking work which shows that there exist constructions based on partially secure components, but which are strictly more secure then any of their components.

### 3.1.2 PRPGs and Cryptographic Composition

Cryptographers have noted that DES is effectively the composition of 16 completely insecure permutation generators. Because DES has withstood much cryptanalysis it is often both considered to be secure and conjectured to be a PRPG. This led Luby and Rackoff to define the notion of a partially secure PRPG and conjecture that the composition of several partially secure PRPGs resulted in a PRPG with stronger security then any of its components. They proved this conjecture to be true in [12]. Later we will use their notion of partially secure generators to study other constructions which appear to be security increasing. Additionally, many of the proofs in this thesis will be based on the proof that Luby and Rackoff presented in [12].

### 3.2 Perfect Cipher Model

Previously we have seen work where the composition and the XOR operators have been used to increase the security of partially secure cryptographic primitives. In this section we assume that block-ciphers (or just cipher for short) exist. A cipher is simply a function generator $F:\{0,1\}^{\kappa} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n}$, which generates functions which intuitively are supposed to look random to computationally limited adversaries. Note that there is no notion of asymptotic security, as opposed to the definitions of PRFGs and PRPGs. The notion of ciphers corresponds to real world cryptographic functions such as DES, and AES which are just specific function generators, and not defined asymptotically for every $n$. In this section we present some work which asks the question of how secure ciphers can be combined so that the minimum number of resources needed to break the construction is more than the number of resources necessary to break the individual cipher.

### 3.2.1 DES Transforms \& Generic Attacks

When DES was first introduced it came under criticism for its small 56-bit key length. People feared that it would be feasible to mount an attack on DES by performing a brute force search on the key space of DES. As the speed of computers has greatly increased since the release of DES, in 1975, this attack has become a serious threat, and in 1993 Weiner provided an estimate that for $\$ 1,000,000$ someone could build a machine which broke DES in an average time of 3.5 hours. In 1999 the cost of such a machine has surely decreased, and the average amount of time needed to break DES is probably much smaller, given the faster processors currently available.

Cryptographers questioned if it was possible to increase the key-length of DES while maintaining the time-tested security of the DES cipher. One solution which was proposed by Rivest, in 1984, is called DESX. We define $\operatorname{DESX}$ as $\operatorname{DESX}_{k_{1}, k_{2}, k}(x)=k_{1} \oplus \operatorname{DES}_{k}\left(k_{2} \oplus\right.$ $x$ ), where $k_{1}, k_{2} \in\{0,1\}^{64}$ and $k \in\{0,1\}^{56}$. Of course this is easily generalized to
arbitrary function and permutation generators. If $F:\{0,1\}^{\kappa} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n}$ is a function generator then define $\mathrm{FX}_{k_{1}, k_{2}, k}(x)=k_{1} \oplus \mathrm{~F}_{k}\left(k_{2} \oplus x\right)$, where now $k_{1}, k_{2} \in\{0,1\}^{n}$ and $k \in\{0,1\}^{\kappa}$.

In [10],[20] and [9] Kilian and Rogaway show that FX does in fact increase the effective key length of a cipher $F$. However, it is not easy to define a model in which bruteforce key-searches on a cipher and a transformation of the cipher can be quantified and compared, and so the model they have developed is as important of a contribution as their result of the effective extended key-length of the FX cipher, as compared to the F cipher. A final observation made by Aiello et. al in [1] is that the model is not only useful in measuring the effect of a cipher transform on key-search attacks, but rather it is useful in modeling the effect of any cipher-transform on any generic attack. A generic attack is one which does not take into account the combinatorial structure of the initial cipher, and therefore a generic attack can only exploit the structure resulting from the transform itself, and information learned from the function in a black-box model.

## The Kilian-Rogaway Model

For presentation purposes we present a simplified version of the model originally presented by Kilian and Rogaway in [10].

Let $F:\{0,1\}^{\kappa} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n}$ be a block-cipher. Let $p$ be a polynomial, and let $T F:\{0,1\}^{p(\kappa)} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n}$ be the cipher which results from applying the transformation $T$ to $F$.

Definition 3.1 A generic-attack adversary $A$ is an algorithm which has unlimited computational power and has access to two oracles:
$F(k, x)$ Oracle - The adversary supplies $a k \in\{0,1\}^{\kappa}$ and an $x \in\{0,1\}^{n}$, and the oracle returns $F(k, x)$.
$E(x)$ Oracle - The adversary supplies an $x \in\{0,1\}^{n}$ and the oracle consistently responds
as if it were in one of the following manners:

World $1-T F_{\bar{k}}(x)$, for $a \bar{k} \in\{0,1\}^{p(\kappa)}$
World $2-\pi(x)$, for $a \pi \in \mathcal{P}^{\boldsymbol{n}}$.

The adversary will be placed in one of the two worlds uniformly at random. Later we will define the security of a function transform as an upper bound on the ability of all adversaries to distinguish between which of the two worlds they were placed. However, an adversary which can query $F$ and $E$ at all possible locations will never have any difficulty distinguishing between worlds. Further, one might argue that in the "real world" no adversary truly has the ability to query each oracle at every point of its respective domain. In practice there are limitations on the number of chosen cipher texts that an adversary has the ability to request. This selection of chosen cipher texts corresponds to queries to the $E$ oracle. Additionally, time restrictions would limit the effective number of queries to $F$ which could be made.

Definition 3.2 An $(m, t)$ generic attack adversary, $A(m, t)$, is a generic-attack adversary, $A$, which makes exactly $m$ queries of the $E$ oracle, and exactly $t$ queries to the $F$ oracle.

Kilian and Rogaway point out that when considering brute-force key searches we must disallow the adversary from considering any internal combinatorial structure of the primitive cipher $F$. This is because we are interested in the effect of the specific transformation of a cipher $F$ on the key length, and not the weaknesses of any particular cipher to combinatorial attack. Therefore, rather than using a specific cipher we draw $F$ uniformly at random from $B$, the set of all function generators. Specifically, if we are interested in a function generator $F:\{0,1\}^{\kappa} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n}$ chosen uniformly at random from set $\mathcal{B}^{\kappa, n}$ of all generators of the same form, then we can choose a random function generator $F$ in the following manner. For each $k \in\{0,1\}^{n}$ we associate $f_{k} \in \mathcal{F}^{n}$,
and we define $F=\left\{\left(k, f_{k}\right) \mid k \in\{0,1\}^{\kappa}\right\}$. Observe that by choosing $F$ in this fashion we prevent an adversary from making use of the combinatorial structure of $F$ by ensuring that it has no structure.

We can now formally define what it means to break the cipher $T F$, which is the result of applying the transform $T$ on a cipher $F$.

Definition 3.3 A generic attack adversary $A(m, t)$ is said to $\epsilon$-break the $T F$ function transform on parameters $\kappa, n$ if:

$$
\left|\operatorname{Pr}_{F \in \mathcal{B}^{\kappa, n}, k \in\{0,1\} p(\kappa)}\left[A^{T F_{k}, F}=1\right]-\operatorname{Pr}_{F \in \mathcal{B}^{\alpha}, n, \pi \in \mathcal{P}^{n}}\left[A^{\pi, F}=1\right]\right| \geq \epsilon
$$

We refer to the above difference as the advantage of $A$, written $A D V_{A}$.

Observe that this notion of breaking $T F$ is very weak. There is no requirement of determining the key, $k$, nor a requirement of producing a previously unknown pair ( $x, T F_{k}(x)$ ), and therefore any upper bound results in this model are very strong.

## The Security of DESX

Using the model defined above Kilian and Rogaway, it is possible to show that the effective key length of a cipher FX: $\{0,1\}^{2 n+\kappa} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n}$ resulting from the application of the DESX transform to a cipher $F:\{0,1\}^{*} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n}$ is longer than $\kappa$.

We first note that, for all $t$, there exists an adversary $A(m, t)$ which which can $\epsilon$ break the cipher $F$, for $\epsilon=\frac{t}{2^{k}}$ (note that breaking $F$ corresponds to breaking $T F$, where $T$ is the identity transformation). In order to $\frac{t}{2^{k}}$ break $F$, the adversary simply performs a brute force key-search. Kilian and Rogaway show the same attack cannot be as effective for the DESX transform, by proving an upper bound on the advantage of any $A(m, t)$ adversary against FX.

Theorem 3.1 (Kilian-Rogaway) Let $A(m, t)$ be a generic attack adversary which $\epsilon$ breaks the DESX transform applied to a cipher $F$ with parameters $\kappa$ and $n$, then $\epsilon \leq$ $\frac{t}{2^{x+n-1-\log m}}$.

The theorem implies that, with the exception of adversaries which have the ability to query the $E$ oracle on close to its entire domain, adversaries perform significantly worse on the FX then they do on F.

## The Security of Composition and Triple-DES

Aiello et. al were able to show in [1] that under the Kilian-Rogaway model, composition increases security against generic attacks.

They define the m-fold cascade transform which given the cipher $F:\{0,1\}^{\kappa} \times$ $\{0,1\}^{n} \rightarrow\{0,1\}^{n}$, produces the cipher $m F:\{0,1\}^{m \kappa} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n}$, where $m F_{k_{1} \bullet \ldots \bullet k_{m}}(x)=F_{k_{1}} \circ \cdots \circ F_{k_{m}}(x)$ and $\left|k_{i}\right|=\kappa$.

Theorem 3.2 (Aiello et. al) Let $A(m, t)$ be a generic attack adversary which $\epsilon$-breaks the m-fold cascade transform applied to a cipher $F$ with parameters $\kappa$ and $n$, then $\epsilon \leq$ $\left(\frac{t}{2^{\kappa}}\right)^{m}$.

The result is tight for $m=2$, as a variant on the meet in the middle attack achieves the bound. For the the cases where $m \geq 3$, Aiello et. al were not able to show that the bound is tight, and they conjecture that the actual bound is much smaller.

For historical reasons Aiello et. al also give an upper bound for the advantage of adversaries on the triple-DES transform, which given $F$ as a cipher outputs a cipher 3DES-F, where 3DES-F $\mathrm{F}_{k_{1} \circ k_{2}}(x)=F_{k_{1}} \circ F_{k_{2}}^{-1} \circ F_{k_{1}}(x)$ and $\left|k_{i}\right|=\kappa$. Note that in the transform the inverse of the second function is done for historical reasons, and has no security related purpose.

Theorem 3.3 (Aiello et. al) Let $A(m, t)$ be a generic attack adversary which $\epsilon$-breaks
the triple-DES transform applied to the cipher $F$ with parameters $\kappa$ and $n$, then $\epsilon \leq\left(\frac{t}{2^{\kappa}}\right)^{2}$.

This is the same bound as for the 2-fold cascade transform and effectively follows from the same proof, but the authors were not able to prove that this bound was tight, and they conjecture that smaller upper bounds do exist.

## Chapter 4

## Proposed Operators Relating to Composition and XOR

Our primary goal is to consider constructions which combine function and permutation generators into a resulting function or permutation generator which has increased security over its component generators. In particular we will be interested in constructions which use partially secure PRPGs and PRFGs to produce new PRPGs and PRFGs. Further, we want to consider constructions where the security of the resulting generator is stronger than that of its component generators, assuming none of the component generators were perfectly secure to begin with. Finally, we want the constructions to be scalable, so that as we scale the construction to be larger, we improve the security of the resulting generators.

In this chapter we will present a list of possible constructions which might be used to combine function and permutation generators in a security increasing fashion. We will consider some of the more obvious constructions which are based on the composition and XOR operators. Again, we consider these operators, as historically they have been two of the most widely used operators in the development of cryptographic systems. We will show that some constructions can be discarded as they either do not increase the security
of the resulting generator, or in some cases the security of the component generators is not even maintained. In subsequent chapters we will show that three of the proposed constructions do increase the security of the resulting generator, but unfortunately we are only able to prove that one of them can be used to achieve 1 -security by scaling the size of the generator.

Before describing the constructions, we will first formally describe how to combine two function generators using a generic operator, and what it means for the security to be necessarily increasing.

Definition 4.1 Let $G=\left\{G^{n}:\{0,1\}^{\ell(n)} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ be a $1-\delta(n)$ secure PRFG. Let $H=\left\{H^{n}:\{0,1\}^{\kappa(n)} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ be a $1-\epsilon(n)$ secure PRFG. Let op $=\left\{o p^{n}:\{0,1\}^{r(n)} \times \mathcal{F}^{n} \times \mathcal{F}^{n} \rightarrow \mathcal{F}^{n} \mid n \in \mathbb{N}\right\}$ be an operator generator, and write $o p^{n}(s, a, b)$ as "a op $p_{s}^{n} b^{\prime}$. Then let $F=\left\{F^{n}:\{0,1\}^{\ell(n)+\kappa(n)+r(n)} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ be the function generator defined by $\mathrm{F}^{n}\left(k_{1} \bullet k_{2} \bullet k_{3}, x\right)=\left(g_{k_{1}}^{n} \circ p_{k_{3}}^{n} h_{k_{2}}^{n}\right)(x)$, where $\left|k_{1}\right|=$ $\ell(n),\left|k_{2}\right|=\kappa(n)$ and $\left|k_{3}\right|=r(n)$. This is written in short-hand as $F=G$ op H . We say that $o p$ is security increasing if, for all polynomial time computable function generators $G$ and H , the generator G op H is a $1-\theta(n)$ secure PRFG; where for some constant $c>0$, and for all sufficiently large $n: \min \{\epsilon(n), \delta(n)\}-\theta(n)>\frac{1}{n^{c}}$. Similarly, we say that op is security preserving if for all sufficiently large $n$ we know that $\min \{\epsilon(n), \delta(n)\}-\theta(n) \leq \frac{1}{n^{c}}$, for all $c>0$. Finally, if op is neither security increasing or preserving then we say that it is possibly security diminishing.

We will now present an operator and an operator generator which are are used in the constructions below.

We define the $\square$ operator (read as box) on two functions $f_{1}$ and $f_{2}$, where $f_{i}$ : $\{0,1\}^{n} \rightarrow\{0,1\}^{n}$, for $i \in\{1,2\}$, as:

$$
\left(f_{1} \square f_{2}\right)(x)=f_{1}\left(f_{1}(x) \oplus f_{2}(x)\right) \oplus f_{2}\left(f_{1}(x) \oplus f_{2}(x)\right)
$$

The operator was originally defined in some working notes by Luby and Rackoff [13].

We next define the $\diamond$ operator generator (read as Diamond), as $\diamond=\left\{\nabla_{r_{1}, r_{2}}^{n} \mid \boldsymbol{n} \in\right.$ $\left.\mathbb{N} \wedge r_{1}, r_{2} \in\{0,1\}^{n}\right\}$. For each $r_{1}, r_{2} \in\{0,1\}^{n}$ we define the operator $\nabla_{r_{1} \bullet r_{2}}^{n}$, which acts on the functions $f_{1}$ and $f_{2}$ as:

$$
\left(f_{1} \diamond_{r_{1} \bullet r_{2}}^{n} f_{2}\right)(x)=f_{1}\left(x \oplus r_{1}\right) \oplus f_{2}\left(x \oplus r_{2}\right)
$$

### 4.1 Constructions with Known Properties

In this section we present the constructions which we were able to classify as either security increasing, preserving or possibly diminishing. We present each construction followed by an explanation of the security properties it has. In some cases the properties are stated, but the explanations are put off for a later, more in depth discussion.

In the constructions below we will assume that $G=G_{i}=G_{j}$ for $i \neq j$. We assume that G a partially secure PRPG or PRFG, depending on the context. The indices are only there to help in describing to constructions in a clear and precise manner.

### 4.1.1 $\quad F(\alpha)=\left(G_{m(n)} \circ \cdots \circ \mathrm{G}_{1}\right)(\alpha)$

The composition of permutation generators results in a permutation generator. Therefore, we will consider the cases when the generators $G_{i}$ are PRPGs and PRFGs separately. We will see that the outcomes are very different.

## $G$ is a PRPG

It was shown by Luby and Rackoff in [12] that the composition of partially secure PRPGs results in a PRPG of increased security when $m(n)=c$ for some constant $c$. In Chapter 5 we will discuss this result and show how to extend the result to permit any $m \in$ $\mathcal{O}(\log \log n)$.

## $G$ is a PRFG

The composition operator is possibly security diminishing when it acts on PRFGs. The composition of a constant number of $1-\delta$ secure PRFGs, for some constant $\delta$, can be less secure than any of its component generators. Let $G=\left\{G^{n} \mid n \in \mathbb{N}\right\}$ be a $1-\delta$ secure PRFG ${ }^{1}$, where for each $n$ the probability $\operatorname{Pr}_{g_{k}^{n} \in G^{n}}\left[\forall x \quad g_{k}^{n}(x)=\overline{0}\right]=\delta$.

For each polynomial in $n$ sized circuit family $\left\{C_{n}\right\}$, for all $c>0$ and for all sufficiently large $n$ let

$$
\left|\operatorname{Pr}_{g_{k}^{n} \in G^{n}}\left[C_{n}\left(g^{n}\right)=1 \mid \exists x g^{n}(x) \neq \overline{0}\right]-\operatorname{Pr}_{f^{n} \in \mathcal{F}-n}\left[C_{n}\left(f^{n}\right)=1\right]\right| \leq \frac{1}{n^{c}} .
$$

Clearly, for any function $f_{k_{1} \cdots \cdots k_{m(n)}^{n}}^{n}=g_{k_{m(n)}} \circ \cdots \circ g_{k_{1}}$ if there exists an $i$ s.t. $g_{k_{i}}$ is the zero function, then $h$ is a constant function. However, for each $i, g_{k_{i}}$ is the zero function with probability $\delta$, and indistinguishable from random with probability $1-\delta$. Therefore, if we randomly choose functions from $\mathrm{G}^{\boldsymbol{n}}$ we expect to choose the first zero function on draw $\frac{1}{\delta}$. Thus, by using a simple Markov Bound we can show that for as few as $c \geq\left\lceil\frac{1}{\delta^{2}}\right\rceil$ composition of $1-\delta$ secure PRFGs, the construction is less than $1-\delta$ secure. Therefore, the construction is possibly security diminishing.

### 4.1.2 $\quad F(\alpha)=\left(\mathrm{G}_{m(n)} \oplus \cdots \oplus \mathrm{G}_{1}\right)(\alpha)$

The XOR of two permutations does not result in a permutation, and therefore we will only consider the construction when $G$ is a partially secure PRFGs. We show that the XOR operator is at best security preserving.

Let $\widehat{\mathrm{G}}=\left\{\widehat{\mathrm{G}}^{n}:\{0,1\}^{k(n)} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n}\right\}$ be a PRFG. We modify $\widehat{\mathrm{G}}$ to form G as follows: for each $n$ and each $\kappa \in\{0,1\}^{k(n)}$ we set $g_{\kappa}^{n}=\hat{g}_{\kappa}^{n}$ and then we set the first bit of $g_{\kappa}^{n}(\overline{0})$ to be 0 . Notice that $G$ is $\frac{1}{2}$ secure, as the $\operatorname{Pr}_{g_{n} \in G^{n}}\left(g_{\kappa}^{n}(\overline{0})^{\prime} s\right.$ first bit is 0$)=1$, while

[^5]$\operatorname{Pr}_{f^{n} \in \mathcal{F}^{n}}\left(f^{n}(\overline{0})^{\prime} s\right.$ first bit is 0$)=\frac{1}{2}$. However, for all $\tilde{g} \in G \oplus G$ it is the case $\tilde{g}(\overline{0})$ 's first bit is 0 . Therefore, it is clear that if we let $G_{i}=G$, for all $i$, then for all $c$ the construction $G_{c} \oplus \cdots \oplus G_{1}$ is still $\frac{1}{2}$ secure.

Unfortunately this example is not very robust. In particular, the example clearly generalizes for PRFGs which are $\delta$ secure, where $\delta=2^{-i}$ for some positive integer $i$. However, we are not aware of any examples of PRFGs which are $1-\delta$ secure when $\delta=2^{-i}$, which when used in the construction above are security preserving. All examples we have tried have been security increasing. This leaves us with the following question.

Open Question 4.1 Is the construction $F(\alpha)=\left(G_{m(n)} \oplus \cdots \oplus G_{1}\right)(\alpha)$ security increasing, when the generators $G_{i}$ are restricted to being $1-\delta$ secure for $\delta<\frac{1}{2}$.

### 4.1.3 $\quad F(\alpha)=\left(G_{m(n)} \square\left(G_{m(n)-1} \square\left(\cdots G_{3} \square\left(G_{2} \square G_{1}\right) \cdots\right)\right)(\alpha)\right.$

When the operator is applied to two permutations the resulting generator is not a permutation, and so we consider only the case in which G is a PRFG. The operator is a security increasing operator, when applied to PRFGs. The proof of this is shown in Chapter 6 , and so we do not discuss it further here.

### 4.1.4 $\quad F(\alpha)=\left(\mathrm{G}_{2 m(n)} \square \mathrm{G}_{2 m(n)-1}\right) \circ \cdots \circ\left(\mathrm{G}_{2} \square \mathrm{G}_{1}\right)(\alpha)$

For the same reason as in section 4.1.3, we consider only the case in which G is a PRFG. This construction is not even security preserving for PRFGs, for the same reasons as those presented in section 4.1.1.

Let $G=\left\{G^{n} \mid n \in \mathbb{N}\right\}$ be a $1-\delta$ secure PRFG, where for each $n$ the probability $\operatorname{Pr}_{g_{k}^{n} \in G^{n}}\left[\forall x g_{k}^{n}(x)=\overline{0}\right]=\delta$. For each polynomial in $n$ sized circuit family $\left\{C_{n}\right\}$, for all $c>0$, and for all sufficiently large $n$ let

$$
\left|\operatorname{Pr}_{g_{n}^{n} \in G^{n}}\left[C_{n}\left(g^{n}\right)=1 \mid \exists x g^{n}(x) \neq \overline{0}\right]-\operatorname{Pr}_{f^{n} \in \mathcal{F}_{n}}\left[C_{n}\left(f^{n}\right)=1\right]\right| \leq \frac{1}{n^{c}} .
$$

Notice that with probability $\delta^{2}$ the generator $\left(G_{i+1} \square G_{i}\right)$ is the zero function. Now following the same argument made in section 4.1.1, we see that the security of the construction for size $m(n) \geq\left\lceil\frac{1}{\delta^{4}}\right\rceil$ is less than $1-\delta$ secure. Therefore, the construction is possibly security diminishing.

### 4.1.5 $\quad F(\alpha)=G_{m(n)} \diamond \cdots \diamond G_{1}(\alpha)$

When the $\diamond$ operator generator is applied to two permutations the result is not a permutation, and so we consider only the case in which G is a PRFG. The $\diamond$ operator generator is a security increasing operator generator, when applied to PRFGs. The proof of this is shown in Chapter 7, and so we do not discuss it further here.

### 4.2 Constructions with Unknown Properties

Below we give a list of constructions for which we are unaware of their status as either security increasing, preserving or possibly diminishing. However, for all but one of the constructions, we will give reasons to suggest that these operators are not interesting, and therefore not worth future study. We list the constructions below, and then discuss them.

1. $F(\alpha)=\left(\left(G_{2 m(n)} \circ G_{2 m(n)-1}\right) \oplus \cdots \oplus\left(G_{2} \circ G_{1}\right)\right)(\alpha)$
2. $F(\alpha)=\left(\left(G_{2 m(n)} \square G_{2 m(n)-1}\right) \oplus \cdots \oplus\left(G_{2} \square G_{1}\right)\right)(\alpha)$
3. $F(\alpha)=\left(\left(G_{2 m(n)} \oplus G_{2 m(n)-1}\right) \circ\left(G_{2 m(n)-2} \oplus G_{2 m(n)-3}\right) \circ \cdots \circ\left(G_{4} \oplus G_{3}\right) \circ\left(G_{2} \oplus G_{1}\right)\right)(\alpha)$

We suspect that construction 1 in the list above is probably security increasing, for it is similar in concept to the construction in section 4.1 .5 which uses the $\diamond$ operator. This is because we can view the expression $H \diamond G$ as $(H \circ R) \oplus(G \circ R)$ where $R^{n}:\{0,1\}^{n} \times$ $\{0,1\}^{n} \rightarrow\{0,1\}^{n}$ is a generator, where for all $k \in\{0,1\}^{n}$ we define $r_{k} \in R^{n}$ as $r_{k}(x)=$
$x \oplus k$. However, the effect of construction 1 on security is probably weaker than that of the construction presented in section 4.1.5. Further, the constructions would have slower running times, and be harder to implement. Finally, due to the structure of the constructions, we believe that any proof that the construction is security increasing would be much more technically complicated then the proof which will be presented for construction 4.1.5. Therefore, we suspect that further research into these construction will fail to provide any insight which has not previously been seen in considering other constructions, and thus we will not mention them further in this work.

Progress in determining the security related properties of construction 2 from the list has been stifled for many of the same reasons we were stifled from making progress in answering Open Question 4.1. We suggest that the solution to this problem is closely related to the solution of previously mentioned open question.

Open Question 4.2 Is the construction $F(\alpha)=\left(\left(G_{2 m(n)} \square G_{2 m(n)-1}\right) \oplus \cdots \oplus\left(G_{2} \square G_{1}\right)\right)(\alpha)$ security increasing, security preserving or security decreasing.

Finally, construction 3 may be slightly security increasing. However, we will show that, even if the construction is security increasing, there is maximum security of $1-\delta^{2}$ which can be achieved if $G$ is a $1-\delta$ secure generator. Therefore, we cannot use this construction to try and create 1 secure generators, and thus it is of limited use.

To observe the upper bound of $1-\delta^{2}$ security, we define $G=\left\{G^{n} \mid n \in \mathbb{N}\right\}$ to be a $1-\delta$ secure PRFG, where for each $n$ the probability $\operatorname{Pr}_{g_{k}^{n} \in G^{n}}\left[\forall x g_{k}^{n}(x)=\overline{0}\right]=\delta$. For each polynomial in $n$ sized circuit family $\left\{C_{n}\right\}$, for all $c>0$, and for all sufficiently large $n$ let

$$
\left\lvert\, \operatorname{Pr}_{g_{k}^{n} \in G^{n}}\left[C_{n}\left(g^{n}\right)=1\left[\exists x g^{n}(x) \neq \overline{0}\right]-\operatorname{Prr}_{f_{n} \in \mathcal{F}^{n}}\left[C_{n}\left(f^{n}\right)=1\right] \left\lvert\, \leq \frac{1}{n^{c}} .\right.\right.\right.
$$

We observe that with probability $\delta^{2}$ the generator $G \oplus G$ is the zero function. Therefore with probability $\delta^{2}$ the generator $\left(G_{2 m(n)} \oplus G_{2 m(n)-1}\right)$ is the zero function, and therefore with probability at least $\delta^{2}$ construction 3 is the zero function. Therefore, the security of
the construction can never be higher than $1-\delta^{2}$, and thus we cannot use it to develop a 1-secure generator.

## Chapter 5

## The Composition Operator

Luby and Rackoff [12] were the first to prove that the composition of two partially secure PRPGs results in a PRPG which has stronger security than either of the initial primitives. In [12] they gave an outline of the proof, and later Akcoglu and Rackoff in [2] filled in the missing details. It would be convenient if, by inductively repeating the argument many times, it could shown that a partially secure PRPG could be composed with itself many times, and result in a completely secure PRPG. Unfortunately due to technical restrictions, which will be explained in the sequel, the given proof only works for a constant number of compositions.

In this chapter we will demonstrate that it is possible to modify the argument presented by Luby and Rackoff to permit a non-constant number of compositions. We will first give the framework for the original result by Luby and Rackoff and explain its limitation to a constant number of compositions. Next, we will extend this result to allow a $\mathcal{O}(\log \log n)$ number of compositions, and in doing so we basically outline the proof given by Luby and Rackoff, but make it less dependent on sampling. This will result in a more secure construction then was possible under the Luby-Rackoff result.

### 5.1 The Luby-Rackoff Result

The main result of Luby and Rackoff is to show that the composition of two partially secure PRPGs results in a PRPG which is provably more secure then either of the component generators. We present their theorem in a slightly different manner than it was originally presented in [12] or in the unpublished manuscript [2]. This presentation makes the limitations of their results more immediate, and is comparable to the presentation of Levin's proof of Yao's XOR lemma in [6].

## Lemma 5.1 (Composition Isolation Lemma - Luby \& Rackoff's Version) -

 There exist fixed polynomials $p_{1}$ and $p_{2}$ such that for all $0 \leq \epsilon, \delta \leq 1$; polynomials $c_{G}, c_{H}$, and $s_{F}$; and permutation generators $H$ and $G$, where $c_{G}(n)$ and $c_{H}(n)$ bound from above the size of the circuits which compute $\mathrm{G}^{n}$ and $\mathrm{H}^{n}$ respectively. Define $\mathrm{F}=\mathrm{G} \circ \mathrm{H}$. Hypothesis: If there exists a family of decision-circuits $\left\{C_{n}\right\}$, where for each $n$ the circuit $C_{n}$ is of size $s_{F}(n)$, and for some $c>0$ and infinitely many $n$ :$$
\left|{\underset{C}{n}}^{\operatorname{Pr}}\left(F^{n}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \epsilon \delta(2-\max \{\epsilon, \delta\})+\frac{1}{n^{c}} .
$$

Conclusion: Then for infinitely many $n$ there exists either a decision-circuit $\Lambda_{n}$ of size $p_{1}\left(n^{c} \cdot c_{G}(n)\right) s_{F}(n)$ for which:

$$
\left\lvert\, \underset{\Lambda_{n}}{\operatorname{Pr}\left(G^{n}\right)-\underset{\Lambda_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) \left\lvert\, \geq \delta+\frac{1}{n^{24 c}}\right. ;, ~ ; ~}\right.
$$

or a decision-circuit $\Upsilon_{n}$ of size $p_{2}\left(n^{c} \cdot c_{G}(n)\right) s_{F}(n)$ for which:

$$
\left|\operatorname{Pr}_{\mathrm{Pr}_{n}}\left(\mathrm{H}^{n}\right)-\underset{\mathrm{Y}_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \epsilon+\frac{1}{n^{3 c}}
$$

or a decision-circuit $\Xi_{n}$ of size $c_{H}(n)+s_{F}(n)$ for which:

$$
\left|\operatorname{Pr}_{\overline{\Xi r}_{n}}\left(\mathrm{G}^{n}\right)-\underset{\bar{\Xi}_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \delta+\frac{1}{n^{6 c}} .
$$

We can now show that by composing a partially secure PRPG with itself a constant number of times we get a significantly more secure PRPG. This result follows from
repetitively applying the Isolation Composition Lemma. We state this formally in the corollary below.

Theorem 5.1 (Composition Theorem - Luby \& Rackoff's Version) Let G be a $1-\epsilon$ secure PRPG. Then for each positive integer $c$, the generator $F=\underbrace{G \circ \ldots \circ G}_{c \text { timea }}$ is $1-\theta$ secure, where $\theta=\epsilon^{c}(2-\epsilon)^{c-1}$.

Proof(Sketch): Let $\left\{C_{n}\right\}$ be a family of circuits. For each $n$ let the size of $C_{n}$ be less than $s_{F}(n)$, for a polynomial $s_{F}$. Let $C_{G}$ and $C_{H}$ be polynomials which bound the size of the circuits needed to compute $G^{n}$ and $H^{n}$ respectively. Assume that, for infinitely many $n, C_{n}$ can distinguish between a random function in $\mathcal{F}^{n}$ and a pseudo-random permutation in $F^{n}$ with probability at least $\epsilon^{c}(2-\epsilon)^{c-1}+\frac{1}{n^{d}}$, for some constant $d$.

Since $F=\underbrace{G \circ(G \circ \cdots \circ G)}_{c \text { times }}$, we know by the Composition-Isolation Theorem that there exits polynomials $p_{2}$ and $p_{1} ;$ and either a circuit of size $\max \left\{s_{F}(n) p_{1}\left(c_{G}(n) n^{d}\right), s_{F}(n)-c_{H}(n)\right\}$ which $\epsilon$ distinguishes $G$; or a circuit of size $\left(s_{F}(n) p_{2}\left(c_{G}(n) n^{d}\right)\right)$ which $\epsilon^{c-1}(2-\epsilon)^{c-2}$ distinguishes $\underbrace{(G \circ \cdots \circ G)}_{c-1 \text { imes }}$. Since the former would contradict the fact that $G$ is $(1-\epsilon)$ secure, it must be the case that there is a circuit which $\epsilon^{c-1}(2-\epsilon)^{c-2}$ distinguishes $\underbrace{(\mathrm{G} \circ \ldots-0 \mathrm{G})}_{c-1 \text { times }}$. We can now re-apply this argument inductively $c-2$ more times to show that there exists a polynomial $p_{2}^{\prime}$ and a circuit of size no larger than $s_{F}(n)\left(p_{2}^{\prime}\left(c_{G}(n) n^{d^{\prime}}\right)\right)^{c-1}$, for some constant $d^{\prime}$, which $\epsilon$ distinguishes $G$ from random functions. This contradicts the claim that $G$ is $1-\epsilon$ secure, and completes the proof.

Notice that the proof fails when the number of compositions, $\boldsymbol{c}$, becomes non-constant. This is because the circuit which $\epsilon$ distinguishes $G_{c}$ is of size $s_{F}(n)\left(p_{2}^{\prime}\left(c_{G}(n) n^{d^{\prime}}\right)\right)^{c-1}$, but when $c$ is constant this functions is asymptotically larger than any polynomial. It is this limitation which prevents us from attaining a PRPG which is completely secure using the Isolation Lemma. In the sequel we will show that this limitation can be partially
overcome, but while more security is attained the goal of complete security is still elusive.
A natural question to contemplate is whether the upper bound of distinguishing $F$ from $\mathcal{F}^{\boldsymbol{n}}$ is tight in the Isolation Lemma. Perhaps there are smaller bounds which could replace the bound of $\epsilon \delta(2-\max \{\epsilon, \delta\}$ ), and otherwise leave the lemma unchanged. In the next section we show that the Isolation Lemma is tight with respect to this bound.

### 5.1.1 The Isolation Lemma is Tight

We show that there exist PRPGs $H$ and $G$ which are respectively $(1-\epsilon)$ and $(1-\delta)$ secure ${ }^{1}$, but when composed together the result is exactly ( $1-\theta$ ) secure, where $\theta=$ $\epsilon \delta(2-\max \{\epsilon, \delta\})$.

## The Construction of H and G

To simplify the presentation we assume that $\delta$ and $\epsilon$ are of the form $\frac{1}{2^{c}}$ or $1-\frac{1}{2^{c}}$, for some constant $c$. Let $\widehat{G}=\left\{\widehat{G}^{n}:\{0,1\}^{\ell(n)} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ be a PRPG. We construct $G=\left\{G^{n}:\{0,1\}^{\ell(n)+c} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ to be $1-\delta$ secure in two steps. We will describe the construction of $\mathrm{G}^{\boldsymbol{n}}$, and note that the construction of $\mathrm{H}^{\mathbf{n}}$ is similar.

First we set a fraction $\delta$ of the keys to correspond to the identity permutation, and the remainder to correspond to permutations chosen from $\widehat{\mathbf{G}}^{\mathbf{n}}$. This is done in two different fashions dependent on the form of $\delta$ as described below:

Case $1\left(\delta=1-\frac{1}{2^{c}}\right)$ For each $k \in\{0,1\}^{\ell(n)+c}$ we set the permutation $g_{k}^{n} \in G^{n}$ to be the identity permutation if the first $c$ bits of $k$ are not all 0 ; otherwise let $\bar{k}$ be the last $\ell(n)$ bits of $k$, and set $g_{k}^{n}=\hat{g}_{\hat{k}}^{n}$, for $\hat{g}_{\hat{k}}^{n} \in \widehat{\mathbf{G}}^{n}$.

Case $2\left(\delta=\frac{1}{2^{c}}\right)$ For each $k \in\{0,1\}^{\ell(n)+c}$ we set the permutation $g_{k}^{n} \in G^{n}$ to be the

[^6]identity permutation if the first $c$ bits of $k$ are all 0 ; otherwise let $\bar{k}$ be the last $\ell(n)$ bits of $k$, and set $g_{k}^{n}=\hat{g}_{\hat{k}}^{n}$, for $\hat{g}_{\bar{k}}^{n} \in \widehat{\mathrm{G}}^{n}$.

The second step in the construction of $G^{n}$ is to ensure that for each $k \in\{0,1\}^{\ell(n)+c}$ we set the value of $g_{k}^{n}(\overline{0})$ in one of two manners, depending on the form of $\delta$. We describe this transformation below:

Case $1\left(\delta=1-\frac{1}{2^{c}}\right)$ For all $k$ set the first $c$ bits of $g_{k}^{n}(\overline{0})$ to 0 ; the last $n-c$ bits remain as they were in $g_{k}^{n}(\overline{0})$.

Case $2\left(\delta=\frac{1}{2^{c}}\right)$ For all $k$ ensure that not all of the first $c$ bits of $g_{k}^{n}(\overline{0})$ are 1 . If they are, set them to be a member of the set $\{0,1\}^{c} \backslash\left\{1^{c}\right\}$ chosen uniformly at random ${ }^{2}$; the last $n-c$ bits remain as they were in $g_{k}^{n}(\overline{0})$.

Notice that in order to maintain the permutation property of $\mathbf{G}^{\mathbf{n}}$ we can simply store both the initial and modified value of $g_{k}^{n}(\overline{0})$. Should a query ever be made to the initial preimage of the modified value of $g_{k}^{n}(\overline{0})$, then we respond with the initial value of $g_{k}^{n}(\overline{0})$.

It remains to show that $G$ and $H$ are in fact $1-\delta$ and $1-\epsilon$ secure, and we refer the interested reader to Appendix A for a proof of this claim.

## The Adversary

We now show that there exists an adversary which can $\epsilon \delta(2-\max \{\epsilon, \delta\})$ distinguish $F=G \circ H$, and thereby show that the Isolation Lemma is tight. We assume that $\delta \geq \epsilon$. A similar argument to the one presented handles the case when $\epsilon>\delta$.

Given a function $w \in \mathcal{F}^{n}$ the adversary, $A$, accepts in one of the following two conditions, dependent on the form of $\delta$ :

[^7]Case $1\left(\delta=1-\frac{1}{2^{c}}\right)$ The adversary $A$ accepts iff the first $c$ bits of $w(\overline{0})$ are 0 ;

Case $2\left(\delta=\frac{1}{2^{c}}\right)$ The adversary $A$ accepts iff the first $c$ bits of $w(\overline{0})$ are not 1 .

We first consider the accepting probability of $A$ if $\boldsymbol{w}$ was chosen uniformly from $\mathrm{F}^{\boldsymbol{n}}$. In this case $w$ is the composition of two permutations chosen randomly from each of the respective generators $\mathrm{G}^{\boldsymbol{n}}$ and $\mathrm{H}^{n}$. We partition the permutations of $\mathrm{G}^{\boldsymbol{n}}$ into two sets. Let $I_{\mathrm{G}^{n}}$ represent the set of all identity permutations in $\mathrm{G}^{\boldsymbol{n}}$, and let $P_{\mathrm{G}^{n}}$ represent the remaining permutations in $\mathrm{G}^{n}$. Similarly, partition $\mathrm{H}^{n}$ into two sets $I_{\mathbf{H}^{n}}$ and $\boldsymbol{P}_{\mathrm{H}^{n}}$. We now consider the probability that $w$ is accepted based on it being the composition of two permutations from the above mentioned partitions.

Case $1\left(w \in I_{G^{n}} \circ I_{\mathrm{H}^{n}}\right)$ The adversary will accept $w$ with probability 1 , and the case occurs with probability $\delta \varepsilon$.

Case $2\left(w \in I_{G^{n}} \circ P_{\mathrm{H}^{n}}\right)$ The adversary will accept $w$ with probability $\frac{1-\delta}{1-\epsilon}$. This is the probability that $A$ would accept a random function conditioned on the fact that $w$ was chosen from $P_{H^{n}}$. The case occurs with probability $\delta(1-\epsilon)$.

Case 3 ( $w \in P_{\mathrm{G}^{n}} \circ I_{\mathrm{H}^{n}}$ ) The adversary will accept $w$ with probability 1 , and the case occurs with probability $(1-\delta) \epsilon$.

Case $4\left(w \in P_{\mathrm{G}^{n}} \circ P_{\mathrm{H}^{n}}\right)$ The adversary will accept $w$ with probability $1-\delta$. This is the probability that the result of a random function evaluated at a random point in its range would meet the acceptance criteria of $A$. The case occurs with probability $(1-\delta)(1-\epsilon)$.

Clearly a random permutation $\boldsymbol{w}$ will be accepted by $A$ with probability $1-\delta$. Allowing $X$ to be $\max \{\epsilon, \delta\}$, and $Y$ to be $\min \{\epsilon, \delta\}$, we see that $X=\delta$ and $Y=\epsilon$. We now see that the distinguishing probability of $A$ on $F$ is:
$\epsilon \delta+\frac{(1-\delta)}{(1-\epsilon)} \delta(1-\epsilon)+(1-\delta) \epsilon+(1-\delta)^{2}(1-\epsilon)-(1-\delta)$

$$
\begin{aligned}
& =\left[Y X+(1-X) X+(1-X) Y+(1-X)^{2}(1-Y)\right]-(1-X) \\
& =[(1-X)+X Y(2-X)]-(1-X) \\
& =\epsilon \delta(2-\max \{\epsilon, \delta\})
\end{aligned}
$$

### 5.2 The Improved Composition Lemma

As we've just seen, we cannot improve the composition lemma by improving the security parameter of the resulting PRPG, in the Isolation Lemma. But, in this section we will prove that if we permit a slightly more liberal notion of security for our partially secure PRPGs, then we can prove a lemma similar to the Isolation Lemma. Further, this lemma permits more than a constant number of compositions of PRPGs. The result is that a non-constant number of compositions of a partially secure PRPG, under the liberal notion of security, gives a PRPG which is provably more secure then the previously stated results.

### 5.2.1 Semi-Secure PRPGs

For the remainder of this chapter we will call a PRPG, H, 1- $\epsilon$ semi-secure, if there exists no family of polynomial-sized circuits $\left\{C_{n}\right\}$ which can $\epsilon+\frac{1}{\log ^{c} n}$ distinguish H from random, for some constant $c$ and infinitely many $n$.

Note that a generator $H$ which is $1-\epsilon$ secure is also $1-\epsilon$ semi-secure. However, the converse need not be true. The reason for the relaxation of the definition of security is due to the observation that the blow-up in circuit size which occurs in the Luby-Rackoff Isolation Lemma is due to a large amount of sampling which must be done on various distributions which occur in the proof. However, the amount of sampling that must be done is inversely proportional to the leniency we permit the adversary in $\epsilon$-distinguishing a PRPG from the random set of functions in order for the generator to still be called $1-\epsilon$ secure. Therefore, by making the security definition slightly weaker we reduced the
amount of sampling required for the proof to hold; which reduces the size of the circuits constructed in the proof of the isolation lemma; which permits a greater number of compositions in the composition lemma; which results in a PRPG with stronger security then was possible under the original isolation lemma.

Finally observe that if, for any constant $0<\epsilon<1$, we are given a $1-\epsilon$ semi-secure generator $F$, then the same generator is definitely $1-\sqrt{\epsilon}$ secure. However, by composing the generator with itself we get a $1-(2 \cdot \epsilon-\sqrt{\epsilon})$ secure generator, which is strictly more secure than the a $1-\epsilon$ secure generator. This result follows from the original LubyRackoff Isolation Lemma, and allows us to transform many -but definitely not all- of the security results under the semi-secure definition to corresponding results under the normal security definition.

Now that we have motivated the semi-secure definition we present the new version of the Isolation Lemma which is less dependent on sampling.

### 5.2.2 A Stronger Isolation Lemma

## Lemma 5.2 (Composition Isolation Lemma - Stronger Version) -

There exists a fixed polynomial $p_{1}$ and a fixed poly-logarithmic function $p_{2}$ such that for all $0<\delta<1 ; \epsilon: \mathbb{N} \rightarrow[0,1]$, where for all sufficiently large $n, \epsilon(n)<\delta ;$ polynomials $c_{G}, c_{H}$, and $s_{F}$; and permutation generators H and G , where $c_{G}(n)$ and $c_{\mathrm{H}}(n)$ bound from above the size of the circuits which compute $\mathrm{G}^{\boldsymbol{n}}$ and $\mathrm{H}^{n}$ respectively. Define $\mathrm{F}=\mathrm{G} \circ \mathrm{H}$. Hypothesis: If there exists a family of decision-circuits $\left\{C_{n}\right\}$, where for each $n$ the circuit $C_{n}$ is of size $s_{F}(n)$, and for some $c>0$ and infinitely many $n$ :

$$
\left|\operatorname{Pr}_{C_{n}}\left(F^{n}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \epsilon(n) \delta(2-\max \{\epsilon, \delta\})+\frac{1}{\log ^{c} n}
$$

Conclusion: Then for infinitely many $n$ :
there exists either a decision-circuit $\Lambda_{n}$ of size $p_{1}\left(\log ^{c} n \cdot c_{G}(n)\right) s_{F}(n)$ for which,

$$
\left|\operatorname{Pr}_{\Lambda_{n}}\left(G^{n}\right)-\underset{\Lambda_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \delta+\frac{1}{n}
$$

or a decision-circuit $\mathrm{Y}_{n}$ of size $p_{2}\left(\log ^{c} n \cdot c_{G}(n)\right) s_{F}(n)$ for which,

$$
\left|\operatorname{Pr}_{\mathrm{P}_{n}}\left(H^{n}\right)-\underset{\mathrm{P}_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \epsilon(n)+\frac{1}{\log ^{2 c} n} ;
$$

or a decision-circuit $\Xi_{n}$ of size $c_{H}(n)+s_{F}(n)$ for which,

$$
\left|\operatorname{Pr}_{\Xi_{n}}\left(G^{n}\right)-\underset{\Xi_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \delta+\frac{1}{n} .
$$

The interesting difference between the two Isolation Lemmas occurs in the differing definitions of $p_{2}$. The new definition of $p_{2}$ increases the upper bound for distinguishing between $\mathrm{H}^{n}$ and random functions. This is what necessitates the definition of semisecurity. Further, the new definition of $p_{2}$ significantly affects the size of the circuit which can distinguish $H^{n}$ from $\mathcal{F}^{n}$. The differing circuit size permits a stronger version of the Composition Theorem, which we will present later.

Another difference to observe is that, in the original Isolation Lemma the results were symmetric. By this we mean that given generators $\widehat{H}$ and $\widehat{G}$, the Lemma applied equally to both of the constructions $F_{1}=\widehat{H} \circ \widehat{G}$ and $F_{2}=\widehat{G} \circ \widehat{H}$. This can be observed by simply renaming the generators $\widehat{G}$ and $\widehat{H}$. Notice that a similar result does not follow from the statement of the strong version of the Isolation Lemma. We note that it is possible to prove a symmetric version of this Lemma by using symmetric versions of the arguments presented in this chapter. Unfortunately, there is one exception, for it seems that in order to prove a symmetric version of one of the lemmas used in the proof, one must first prove the lemma as it is presented in this chapter, and then use this result to prove the lemma's symmetric version. We will discuss this further in the sequel.

## Theorem 5.2 (Composition Theorem - Stronger Version) -

Let $0<\epsilon<1$, and let $G$ be a $1-\epsilon$ secure PRPG. Then for each $f \in \mathcal{O}(\log \log n)$ the generator $F=\underbrace{G \circ \ldots-0 G}_{f(n) \text { timet }}$ is $1-\theta(n)$ semi-secure, where $\theta(n)=\epsilon^{f(n)}(2-\epsilon)^{f(n)-1}$.

Proof: Similar to the proof of Theorem 5.1.

Observe that this theorem shows that given some generator which is $1-\epsilon$ secure, where $\epsilon=\frac{1}{c}$ for some $c>0$, then by performing $\mathcal{O}(\log \log n)$ compositions of the generator we can achieve a $1-\frac{1}{\log ^{d} n}$ semi-secure generator, for any constant $d>0$. We then observe the fact that a $1-\frac{1}{\log ^{d} n}$ semi-secure secure generator is $1-\frac{1}{\log ^{d-1} n}$ secure. Therefore by performing $\mathcal{O}(\log \log n)$ compositions of the generator we can achieve a $1-\frac{1}{\log ^{d r} n}$ secure generator, for any constant $d^{\prime}>0$.

Before presenting the proof of Theorem 5.2 we cite some lemmas used in the proof which are either commonly seen in the literature or which have previously appeared in the manuscript of Akcoglu and Rackoff [2].

Lemma 5.3 (Chernoff bound -(Weak)) Let $x_{1}, \ldots, x_{\log _{n} n}$ be i.i.d.r.v. which take the values 0 or 1 with probabilities $q$ or $p=1-q$ respectively. Let $X_{\log ^{\prime} n}=\frac{1}{\log ^{t} n} \sum_{i=1}^{\log ^{2} n} x_{i}$. Then for any $k$ and $l$, there exists a $t$ such that:

$$
\operatorname{Pr}\left[\left|X_{\log ^{t} n}-p\right| \geq \frac{1}{\log ^{k} n}\right] \leq \frac{1}{n^{l}}
$$

Lemma 5.4 (Chernoff Bound-(Strong)) Let $x_{1}, . ., x_{n^{t}}$ be i.i.d.r.v. which take the values 0 or 1 with probabilities $q$ or $p=1-p$ respectively. Let $X_{n^{\varepsilon}}=\frac{1}{n^{t}} \sum_{i=1}^{n^{t}} x_{i}$. Then for any $k$ and $l$, there exists a $t$ s.t.:

$$
\operatorname{Pr}\left[\left|X_{n^{t}}-p\right| \geq \frac{1}{n^{k}}\right] \leq \frac{1}{2^{n^{i}}}
$$

Both versions of the Chernoff Bound follow immediately from the standard proof of the Chernoff bound. For a proof of the Chernoff bound see [17] or [19].

As mentioned in section 2.1, there is no need to have probabilistic circuits as adversaries, as we can always derandomize them. This is done by fixing the random input bits of the circuit to be a specific string of bits which "does well" on the average input, where "does well" is defined relative to how we are attempting to use the circuit. Next we will present a lemma which is a formalization of a similar but more complicated idea.

In this case we have a probabilistic circuit which "does well" on two different distributions of inputs (note these are distributions over the inputs of the circuit, and not the distributions over the strings of random bits which are used by the circuit). We will show that that there exists a specific string which can be fixed as the random input bits of the circuit, so that the circuit is now guaranteed to do "almost as well" on both distributions of inputs, as it originally did on each distribution individually, when the circuit was still probabilistic.

Lemma 5.5 (Derandomization Lemma) Let $C_{n}(w, r)$ be a probabilistic oracle-circuit, where $w \in \mathcal{F}^{n}$ represents the oracle function, and $r$ represents $C_{n}$ 's random input bits. Let $D_{1}^{n}$ and $D_{2}^{n}$ be two distributions over $\mathcal{F}^{n}$, and let $R_{n}$ be the distribution over $C_{n}$ 's random bits. Let $P: \mathcal{F}^{n} \times \mathbb{R} \rightarrow\{0,1\}$ be a predicate. Then, If

$$
\operatorname{Pr}_{w \in D_{1}^{2}, r \in R_{n}}^{\operatorname{Pr}}\left[P\left(w, C_{n}(w, r)\right)=1\right] \geq 1-p \quad \text { and } \quad \operatorname{Pr}_{w \in D_{2}^{2}, r \in R_{n}}\left[P\left(w, C_{n}(w, r)\right)=1\right] \geq 1-p
$$

then there exists an $\tilde{r} \in R_{n}$ such that $\operatorname{Pr}_{r_{w \in D_{i}}}\left[P\left(w, C_{n}(w, \tilde{r})\right)=1\right] \geq 1-2 p$, for $i \in\{1,2\}$.

Intuitively, in the lemma above, the circuit $C$ is trying to compute a value for the oracle $w$. However, since $C$ is probabilistic, the value of the circuit on input $w$ may be different for different strings of random bits. Therefore, we measure whether or not the circuit has computed an acceptable value, given the oracle $w$, by using the predicate $P$, which indicates, on all pairs $(w, v)$, whether $v$ is an acceptable value for $C$ to output, given the oracle $w$. For an example of this lemma's use see section 5.3.1.

Proof: We build a third distribution $D_{3}^{n}$ by first choosing a random random bit $i \in$ $\{1,2\}$, and then choosing a random function $w \in D_{i}$. Note that

$$
\operatorname{Pr}_{w \in D_{3}^{n}, r \in R_{n}}\left[P\left(w, C_{n}(w, r)\right)=1\right] \geq \frac{1}{2}(1-p)+\frac{1}{2}(1-p)=1-p
$$

and so there exists an $\tilde{r} \in R_{n}$, for which

$$
\begin{equation*}
\operatorname{Pr}_{w \in D_{3}^{n}}\left[P\left(w, C_{n}(w, \tilde{r})\right)=1\right] \geq 1-p \tag{5.1}
\end{equation*}
$$

Now it must be the case that $\operatorname{Pr}_{w \in D_{i}^{n}}\left[P\left(w, C_{n}(w, \tilde{r})\right)=1\right] \geq 1-2 p$, for $i \in\{1,2\}$; otherwise $D_{3}^{n}\left[P\left(w, C_{n}(w, \tilde{r})\right)=1\right]<\frac{1}{2}(1-2 p)+\frac{1}{2}=1-p$, and this contradicts (5.1).

Lemma 5.6 For any $S \subset \mathcal{P}^{n}$ and any decision-circuit $C_{n}$ the $\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n} \circ S\right)=\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)$.

Proof: Follows from the fact that each element of $S$ is a permutation, and therefore the resulting distribution on the queries to $\mathcal{F}^{\boldsymbol{n}}$ is invariant under when it is composed with the set $S$.

Lemma 5.7 (Luby-Rackoff) For each $n$, let $S^{n} \subset \mathcal{F}^{n}$ such that $\frac{\left|S^{n}\right|}{\left|\mathcal{F}^{n}\right|} \geq \frac{1}{n^{d}}$, for some constant d. Then, for any constant $k$, for any polynomial-sized decision-circuit family $\left\{C_{n}\right\}$, and for all but a $\frac{1}{2^{n / 3}}$ fraction of $\nu \in \mathcal{F}^{n}$ :

$$
\left|\underset{C_{\mathbf{n}}}{\operatorname{Pr}}\left(\nu \circ S^{n}\right)-\underset{C_{\mathbf{n}}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right|<\frac{1}{n^{k}},
$$

for sufficiently large $n$.
This lemma follows directly from the lemma stated bellow, and the strong version of the Chernoff Bound.

Lemma 5.8 Let $S^{n}$ and $\left\{C_{n}\right\}$ be as in Lemma 5.7. Then, given $k$ there is an $r$ such that for all sufficiently large $n$ :

$$
\operatorname{Pr}_{\nu, s_{1}, \ldots, s_{n} r \in \mathcal{F}^{n} \times\left(S^{n}\right)^{n^{r}}}^{\operatorname{Pr}}\left(\left|\frac{1}{n^{r}} \sum_{i=1}^{n^{r}} C_{n}\left(\nu \circ s_{i}\right)-\underset{C_{n}}{\operatorname{Pr}\left(\mathcal{F}^{n}\right)}\right|>\frac{1}{n^{k}}\right)<\frac{1}{2^{n / 3}} .
$$

This lemma is proven in [2], and a complete proof of similar lemmas will be presented in Chapters 6 and 7 (See the proofs of Lemmas 6.3 and 7.3). Therefore, we will only give a sketch of the lemma's proof.

Proof(Sketch): Observe that whenever we have a circuit $C_{n}$ which has oracle-gates which correspond to a function $f$, it can be easily modified, by adding at most a polynomial number of gates and connections between gates, to render a new circuit $\widehat{C_{n}}$, where
$\widehat{C_{n}}$ never repeats queries to $f$ on the same input. Further, any polynomial sized family of circuits $\left\{C_{n}\right\}$, in which each circuit $C_{n}$ has at most $m(n)$ oracle gates, can be easily be modified to a equivalent family of circuits $\left\{\overline{C_{n}}\right\}$, where each circuit $\bar{C}_{n}$ has exactly $m(n)$ oracle gates. We shall assume that all circuits in this proof sketch are of the modified forms described above.

We perform two experiments. In the first we randomly choose $f_{1}, \ldots, f_{n^{r}} \in\left(\mathcal{F}^{\boldsymbol{n}}\right)^{\boldsymbol{n}^{\boldsymbol{r}}}$, and compute the average value, $E_{1}$, of the circuit on these function. Specifically we let $E_{1}=\frac{1}{n^{r}} \sum_{i=1}^{n^{r}} C_{n}\left(f_{i}\right)$. In the second experiment we randomly choose $\nu, s_{1}, \ldots, s_{n^{r}} \in$ $\mathcal{F}^{n} \times\left(S^{n}\right)^{n^{r}}$, and compute the average value, $E_{2}$, of the circuit on the functions $\nu \circ s_{i}$. Specifically we let $E_{2}=\frac{1}{n^{r}} \sum_{i=1}^{n^{r}} C_{n}\left(\nu \circ s_{i}\right)$. Using the strong version of the Chernoff Bound we see that the probability that $\left|E_{1}-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right|>\frac{1}{n^{k}}$ is less than $2^{n / 2}$. We then show that the difference between $E_{1}$ and $E_{2}$ is bounded to be less than $2^{n / 2}$, and the result follows.

To bound the difference between $E_{1}$ and $E_{2}$ we notice that from the perspective of the circuit $C_{n}$ that the functions $\left(f \circ s_{1}\right), \ldots,\left(f \circ s_{n r}\right)$ will appear to be random functions chosen uniformly from $\mathcal{F}^{n}$, so long as $f$ is never queried on the same input throughout experiment two. Since each circuit never performs the same oracle query twice, we know that for each $i$, the function $f$ will not be queried twice on the same input during the evaluation of $C_{n}\left(f \circ s_{i}\right)$. But for $i \neq j$, it may be the case that during the evaluations of $C_{n}\left(f \circ s_{i}\right)$ and $C_{n}\left(f \circ s_{j}\right)$, that the function $f$ is queried on the same input twice. However, we can bound this probability. WLOG, assume that $i<j$, and let $\alpha$ be a query made during the evaluation of $C_{n}\left(f \circ s_{i}\right)$. Let $\beta_{k}$ to be the $k^{\text {th }}$ query which is made during the evaluation of $C_{n}\left(f \circ s_{j}\right)$. If we fix $s_{i}$ and $\alpha$, then the probability that $s_{i}(\alpha)=s_{j}\left(\beta_{k}\right)$ is:

$$
\operatorname{Pr}_{s_{j} \in S^{n}}\left[s_{i}(\alpha)=s_{j}\left(\beta_{k}\right)\right]=\operatorname{Pr}_{s_{j} \in \mathcal{F} n}\left[s_{i}(\alpha)=s_{j}\left(\beta_{k}\right) \mid s_{j} \in S^{n}\right] \leq \frac{\frac{1}{2^{n}}}{\frac{1}{n^{d}}}=\frac{n^{d}}{2^{n}}
$$

where the $\frac{1}{n^{d}}$ in the last inequality is due to the fact that $\frac{\left|S^{n}\right|}{\left|\mathcal{F}^{n}\right|}>\frac{1}{n^{d}}$. Since each circuit can make at most a polynomial in $n$ number of queries, and since there are a polynomial in $n$ number of circuit evaluations, we can bound the probability of this event occurring
to be less than $2^{n / 2}$. The argument is made by performing simple counting arguments.

As mentioned earlier, there is one lemma for which the symmetric version of the lemma does not follow immediately from a symmetric version of the argument presented in this chapter. The lemma in question is Lemma 5.7. Note that Lemma 5.7 follows immediately from Lemma 5.8, but observe that in the proof of Lemma 5.8 we need to consider functions of the form $f \circ s_{i}$, and further observe that an argument similar to the one presented would not work for functions of the form $s_{i} \circ f$. Therefore, we cannot directly prove a symmetric version of Lemma 5.8 by presenting a proof similar to the one just given for Lemma 5.8, and therefore we cannot conclude a symmetric version of Lemma 5.7. We note that a symmetrical version of Lemma 5.7 may be proven by using an averaging argument, and a slightly stronger version of Lemma 5.7 then is stated. We refer the interested reader to [2] for further details.

### 5.3 Proof of Lemma 5.2

Assume that there exists a polynomial-sized decision-circuit family $\left\{C_{n}\right\}$, which for some constant $c>0$ and infinitely many $n$ distinguishes $F^{n}$ from $\mathcal{F}^{n}$ with probability at least $\epsilon(n) \delta(2-\delta)+\frac{1}{\log ^{\varepsilon} n}$.

Lemma 5.9 (Luby-Rackoff) Either there exists a family of decision-circuits $\left\{\Lambda_{n}\right\}$, where for each $n$ the circuit $\Lambda_{n}$ is of size $p_{\mathbf{l}}\left(\log ^{c} n \cdot c_{G}(n)\right) s_{F}(n)$, and for infinitely many $n$ :

$$
\left|\underset{\Lambda_{n}}{\operatorname{Pr}\left(G^{n}\right)-\operatorname{Pr}} \operatorname{Pr}\left(\mathcal{F}^{n}\right)\right| \geq \delta+\frac{1}{n} ;
$$

or for each $n$ let

$$
K_{\mathrm{n}}=\operatorname{Pr} \operatorname{Pr}_{\mathrm{n}}\left(\mathcal{F}^{n}\right)+\delta\left(1-\operatorname{Pr}_{\mathrm{C}_{\mathrm{n}}}\left(\mathcal{F}^{n}\right)\right)+\frac{1}{n^{1 / 3}} \quad \text { and } \quad L_{n}=\operatorname{Pr}_{C_{\mathrm{n}}}\left(\mathcal{F}^{n}\right)-\delta \operatorname{Pr}_{C_{\mathrm{n}}}\left(\mathcal{F}^{n}\right)-\frac{1}{n^{1 / 3}} ;
$$

and let

$$
S^{n}=\left\{w \in \mathcal{F}^{n}| |_{C_{n}}^{\operatorname{Pr}}\left(G^{n} \circ w\right) \geq K_{n}\right\} \quad \text { and } \quad T^{n}=\left\{w \in \mathcal{F}^{n}| |_{C_{n}}^{\left.\operatorname{Pr}\left(G^{n} \circ w\right) \leq L_{n}\right\}}\right.
$$

Then:

$$
\operatorname{Pr}_{w \in \mathcal{F} n}\left(w \in S^{n}\right) \leq \frac{1}{n} \quad \text { and } \quad \operatorname{Pr}_{w \in \mathcal{F} n}\left(w \in T^{n}\right) \leq \frac{1}{n}
$$

for all sufficiently large $n$.

Proof(Sketch): If there exists a family of circuits $\left\{\Lambda_{n}\right\}$ with the above mentioned properties then we are finished. Otherwise, no such family of circuits exist, and we need to prove the second case of the lemma.

We introduce the following notation. Let

$$
K_{n}(i)=\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)+\delta\left(1-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right)+\frac{1}{n^{i}} \quad \text { and } \quad L_{n}(i)=\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)-\delta \underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)-\frac{1}{n^{i}}
$$

and let

Notice that $K_{n}\left(\frac{1}{3}\right)=K_{n}$, and that $S_{n}\left(\frac{1}{3}\right)=S_{n}$, and similarly that $L_{n}\left(\frac{1}{3}\right)=L_{n}$ and that $T_{n}\left(\frac{1}{3}\right)=T_{n}$.

We will assume for contradiction that for all sufficiently large $n$ the $\operatorname{Pr}_{w^{n} \in \mathcal{F}^{n}}$ ( $w^{n} \in$ $\left.S^{n}\left(\frac{1}{3}\right)\right) \geq \frac{1}{n}$. We will show that there exists a family of decision-circuits $\left\{\widehat{\Lambda}_{n}\right\}$, where for each $n$ the circuit $\widehat{\Lambda}_{n}$ is of size $p_{1}\left(\log ^{c} n \cdot c_{G}(n)\right) s_{F}(n)$; and for all sufficiently large $n$ : $\left|\operatorname{Pr}_{\bar{\Lambda}_{n}}\left(G^{n}\right)-\operatorname{Pr}_{\bar{\Lambda}_{n}}\left(\mathcal{F}^{n}\right)\right| \geq \delta+\frac{1}{n}$; contradicting the fact that no such family of circuits exist. First we assume that we can randomly and uniformly sample from $S^{n}\left(\frac{1}{3}\right)$ (we will deal with the issue of uniform sampling from $S^{n}\left(\frac{1}{3}\right)$ later). Then, we can uniformly sample $f_{1}, \ldots, f_{n^{r}} \in\left(S^{n}\left(\frac{1}{3}\right)\right)^{n^{r}}$, and construct a probabilistic circuit $A_{n}$, which computes $A_{n}(\nu)=\frac{1}{n^{r}} \sum_{i=1}^{n^{r}} C_{n}\left(\nu \circ f_{i}\right)$. We choose $r$ so that for a randomly chosen function, $\nu \in \mathcal{F}^{\boldsymbol{n}}$, the probability that $\left|A_{n}(\nu)-\operatorname{Pr}_{C_{n}}\left(\nu \circ S^{n}\left(\frac{1}{3}\right)\right)\right| \geq \frac{1}{n}$ is less than $2^{n / 2} ;$ and for a randomly
chosen $\nu$ the probability that $\left|A_{n}(\nu)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right| \geq \frac{1}{n}$ is less then $2^{n / 3}$. Both of these properties are possible due to the Chernoff Bound and Lemma 5.7 respectively.

We will now construct a probabilistic circuit $A_{n}^{\prime}(\nu)$ which will accept $\nu$ iff $\boldsymbol{A}_{n}(\nu)>$ $\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)+\frac{1}{n^{2 i}}$. Notice that a random function $\nu \in \mathcal{F}^{n}$ will be accepted by $A_{n}^{\prime}$ with negligible probability. Also notice that $E_{g \in G^{n}}\left[A_{n}(g)\right]>K_{n}\left(\frac{1}{3}\right)$. Using a simply averaging argument and the two facts just mentioned we can show that at least a fraction $\delta+\frac{1}{n}$ of the $g \in \mathrm{G}^{n}$ will be accepted by $A_{n}^{\prime}$.

Allowing $\widehat{\Lambda}_{n}$ to be a derandomized version of $A_{n}^{\prime}$, we have a family of circuits $\left\{\widehat{\Lambda}_{n}\right\}$ for which all sufficiently large $n$ it is the case that $\left|\operatorname{Pr}_{\boldsymbol{\Lambda}_{n}}\left(G^{n}\right)-\operatorname{Pr}_{\boldsymbol{\Lambda}_{n}}\left(\mathcal{F}^{n}\right)\right| \geqslant \delta+\frac{1}{n}$. This provides a contradiction. We can now perform a similar argument by assuming that for all sufficiently large $n$ the $\operatorname{Pr}_{w^{n} \in \mathcal{F} n}\left(w^{n} \in T^{n}\left(\frac{1}{3}\right)\right) \geq \frac{1}{n}$. We can then derive a similar contradiction. This then proves the lemma.

Unfortunately, in the above argument we assumed that we could sample uniformly from $S^{n}\left(\frac{1}{3}\right)$, and in practice this is not possible, so we must show how to get around this. Notice that in the above argument our circuits are significantly smaller than is required by the lemma, and the reason is that many extra gates are needed for properly sampling $S^{n}\left(\frac{1}{3}\right)$. We construct a probabilistic circuit $B_{n}(w)=\frac{1}{n^{s}} \sum_{i=1}^{n^{s}} C_{n}\left(g_{i} \circ w\right)$, where $g_{1}, \ldots, g_{n}$ are randomly chosen from $\left(G^{n}\right)^{n^{2}}$. Using the Chernoff bound, the value $s$ is chosen so that the probability that $\left|B_{n}(w)-\operatorname{Pr}_{C_{n}}\left(G^{n} \circ w\right)\right|>\frac{1}{n}$ is negligible. We now create a probabilistic circuit $B_{n}^{\prime}$ which accepts a function $w$ iff $B_{n}(w) \geq K_{n}\left(\frac{2}{3}\right)$. This algorithm will accept all but a negligible number of the functions $w \in S^{n}\left(\frac{1}{3}\right)$, and it will reject all but a negligible number of functions $w \notin S^{n}(1)$. We now consider the set $R^{n}=\left\{w \left\lvert\, w \in S^{n}(1) \wedge A_{n}(w) \geq K_{n}\left(\frac{2}{3}\right)\right.\right\}$, and note that since $\frac{\left|S_{n}\left(\frac{1}{3}\right)\right|}{\left|\mathcal{F}^{n}\right|} \geq \frac{1}{n^{d}}$, for some $d$, then it is the case that $\frac{\left|R^{r}\right|}{\left|F^{n}\right|} \geq \frac{1}{n^{d}}$, for some $d^{\prime}$.

We now simulate the process of choosing random functions in $R^{n}$ as follows. We run $B_{n}^{\prime}$, but for each function-oracle query, $\alpha_{i}$, that the circuit $B_{n}^{\prime}$ makes, we substitute the output of the oracle with a random string $r_{i} \in\{0,1\}^{n}$. We assume that $B_{n}^{\prime}$ has
$m(n)$ oracle gates, and has been constructed so that it never makes the same oracle query twice. Now if $B_{n}^{\prime}$ accepts, we know that the set $P F=\bigcup_{i=1}^{m(n)}\left\{\left(\alpha_{i}, r_{i}\right)\right\}$ defines a partial function. All random extensions of $P F$ will be accepted by $B_{n}^{\prime}$. Therefore, except with a negligibly small probability of error, the random extensions of $P F$ are all in $R_{n}$. Therefore, we choose a random function in $R_{n}$ by running $B_{n}^{\prime}$ as described above until it accepts. We then save the partial function, and randomly extend it as needed. Since $R_{n}$ represents a fraction greater than $\frac{1}{n^{1-\lambda}}$ of $\mathcal{F}^{n}$, for any $0<\lambda<1$ and sufficiently large $n$, we can simulate $B_{n}^{r}$ a polynomial number of times, and with very high probability we will find a random $w \in R_{n}$. Therefore, while we cannot randomly sample from $S^{n}\left(\frac{1}{3}\right)$, we can sample from $R_{n} \subset S^{n}(1)$. Combining this sampling method with small technical modifications to the original argument will allow us to prove the lemma.

Lemma 5.10 Either there exists a family of decision circuits $\left\{\Xi_{n}\right\}$, where for each $n$ the circuit $\Xi_{n}$ is of size $c_{H}(n)+s_{F}(n)$, and for infinitely many $n$ :

$$
\left|\operatorname{Pr}_{\underline{I}}^{\operatorname{Pr}}\left(\mathrm{G}^{n}\right)-\underset{\Xi_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \delta+\frac{1}{n}
$$

or for all sufficiently large $n$ and all $h^{n} \in H^{n}$ :

$$
\left|\operatorname{Pr}_{C_{n}}\left(G \circ h^{n}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right|<\delta+\frac{1}{n}
$$

Proof: If a family of circuit $\left\{\Xi_{n}\right\}$ with the above properties exists, we are done. Otherwise, there exists no such family of circuits. Assume that for the circuit family $\left\{C_{n}\right\}$ there exists a set $\left\{h^{n} \in \mathrm{H}^{n} \mid n \in \mathbb{N}\right\}$ such that for all sufficiently large $n$ it is the case that $\left|\operatorname{Pr}_{C_{n}}\left(G^{n} \circ h^{n}\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right| \geq \delta+\frac{1}{n}$. We create the circuit family $\left\{D_{n}\right\}$, where $\left.D_{n}(\nu)=C_{n}\left(\nu \circ h^{n}\right)\right\}$. We note that $\left\{D_{n}\right\}$ is capable of $\delta$ distinguishing $G$ from random:

$$
\begin{aligned}
\operatorname{Pr}_{D_{n}}\left(G^{n}\right)-\operatorname{Pr}_{D_{n}}\left(\mathcal{F}^{n}\right) & =\operatorname{Pr}_{C_{n}}\left(G^{n} \circ h^{n}\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n} \circ h^{n}\right) \\
& =\operatorname{Pr}_{C_{n}}\left(G^{n} \circ h^{n}\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right) \\
& \geq \delta+\frac{1}{n}
\end{aligned}
$$

Letting $D_{n}=\Xi_{n}$ we have a constructed a family of circuits $\left\{\Xi_{n}\right\}$. This contradicts the fact that there is no circuit family of size $c_{H}(n)+s_{F}(n)$, such that for infinitely many $n$ :

$$
\left\lvert\, \underset{\Xi_{n}}{\operatorname{Pr}\left(G^{n}\right)-\underset{\Xi_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) \left\lvert\, \geq \delta+\frac{1}{n}\right., ~}\right.
$$

proving the lemma.

### 5.3.1 Main Argument

Let $\bar{\epsilon}_{n}=\epsilon(n)+\frac{1}{\log ^{\varepsilon} n}$ and similarly let $\bar{\delta}_{n}=\delta+\frac{1}{\log ^{c} n}$. Then the family of decision circuits, $\left\{C_{n}\right\}$, is capable of distinguishing $\mathrm{G}^{n} \circ \mathrm{H}^{n}$ from $\mathcal{F}^{n}$ with probability at least $\bar{\epsilon}_{n} \bar{\delta}_{n}\left(2-\bar{\delta}_{n}\right)$, for infinitely many $n$, as this value is strictly smaller than $\epsilon \delta(2-\delta)+\frac{1}{\log ^{\varepsilon} n}$, for sufficiently large $n$. We assume that there exists no family of decision circuits $\left\{\Lambda_{n}\right\}$, where each circuit is of size $p_{1}\left(\log ^{c} n \cdot c_{G}(n)\right) s_{F}(n)$, such that for infinitely many $n$ :

$$
\left|\operatorname{Pr}_{\Lambda_{n}}\left(G^{n}\right)-\underset{\Lambda_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \delta+\frac{1}{n}
$$

From the above assumption and Lemma 5.9 we know that for

$$
\begin{equation*}
\left.K_{n}=\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)+\delta\left(1-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right)+\frac{1}{n^{1 / 3}} \quad \text { and } \quad S^{n}=\left\{w \in \mathcal{F}^{n}| |_{C_{n}}^{\operatorname{Pr}\left(G^{n}\right.} \circ w\right) \geq K_{n}\right\}, \tag{5.2}
\end{equation*}
$$

that $\operatorname{Pr}_{w \in \mathcal{F} n}\left(w \in S^{n}\right) \leq \frac{1}{n}$, for all sufficiently large $n$. Similarly we know that for
that $\operatorname{Pr}_{w \in \mathcal{F n}}\left(w \in T^{n}\right) \leq \frac{1}{n}$, for all sufficiently large $n$.
Next, we assume that there exists no family of decision circuits $\left\{\Xi_{n}\right\}$, where each circuit $\Xi_{n}$ is of size $c_{H}(n)+s_{F}(n)$, such that for infinitely many $n$ :

$$
\left|\frac{\operatorname{Pr}}{\Xi_{n}}\left(G^{n}\right)-\underset{\Xi_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \delta+\frac{1}{n}
$$

From the above assumption and Lemma 5.10 , we know that for all sufficiently large $n$ and all $h^{n} \in H^{n}$ :

$$
\begin{equation*}
\left|\operatorname{Pr}_{C_{n}}\left(G \circ h^{n}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right|<\delta+\frac{1}{n} . \tag{5.4}
\end{equation*}
$$

We will construct a family of decision circuits $\left\{\Upsilon_{n}\right\}$, where each circuit $\Upsilon_{n}$ is of size $p_{2}\left(\log ^{c} n \cdot c_{G}(n)\right) s_{F}(n)$, such that for infinitely many $n$ :

$$
\left|\operatorname{Pr}_{\mathrm{r}_{n}}\left(\mathrm{H}^{n}\right)-\underset{\mathrm{P}_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \epsilon(n)+\frac{1}{\log ^{2 c} n}
$$

proving the lemma. In the sequel we show how to construct $\Upsilon_{n}$ for each $\boldsymbol{n}$ such that both $\left|\operatorname{Pr}_{C_{n}}\left(F^{n}\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right| \geq \epsilon(n) \delta(2-\max \{\epsilon(n), \delta\})+\frac{1}{n^{c}}$, and $n$ is sufficiently large for all inequalities in the following proof to hold true.

We first give an overview of the proof. We will construct a circuit $B_{n}$ (via an intermediate circuit $A_{n}$ ) which, given as input a function $w$, almost surely approximates the value $\operatorname{Pr}_{C_{n}}\left(G^{n} \circ w\right)$. From Equations 5.2 and 5.3, we know that for almost all random functions $w \in \mathcal{F}^{n}$ that $\operatorname{Pr}_{C_{n}}\left(G^{n} \circ w\right)$ will be bounded from above by the value $K_{n}$, and below by the value $L_{n}$. We break the proof into two cases.

In the first case, if for a large fraction of the $h \in H_{n}$ (a fraction greater than $\epsilon(n)+\frac{1}{\log ^{c} n}$ ) it is the case that $\operatorname{Pr}_{C_{n}}\left(\mathrm{G}^{n} \circ h\right)$ falls outside of the range $\left[L_{n}, K_{n}\right.$ ], then we can distinguish between $\mathrm{H}^{n}$ and $\mathcal{F}^{n}$ by computing $B(w)$, and accepting if $B(w)$ is greater than $K_{n}$ or less than $L_{n}$. Because the sets $S^{n}$ and $T^{n}$ contain all of the functions $w$ in $\mathcal{F}^{n}$, for which $\operatorname{Pr}_{C_{n}}\left(G^{n} \circ w\right)$ is less than or greater than $L_{n}$ and $K_{n}$ respectively, and since $\frac{\left|S^{n}\right|}{\left|F^{n}\right|}<\frac{1}{n}$ and $\frac{\left|T^{n}\right|}{\left|\mathcal{F}^{n}\right|}<\frac{1}{n}$; we know that about a fraction $\frac{2}{n}$ of the $w$ will be accepted by $B(n)$.

In the second case there is not a large enough fraction of $h \in G^{n}$ for which $\operatorname{Pr}_{C_{n}}\left(G^{n} \circ h\right)$ falls outside of the range $\left[L_{n}, K_{n}\right.$ ]. In this case, to distinguish $H^{n}$ from $\mathcal{F}^{n}$ we use a technique of Levin's, and construct a circuit $D_{n}$ which on input $\boldsymbol{w}$ tosses a biased coin which is heads with probability $B_{n}(w)$, and tails otherwise. The circuit $D_{n}$ accepts the input $w$ if the result of the flip is heads. We will now present the technical details of the proof which was just described.

We construct a probabilistic circuit $A_{n}$ such that

$$
A_{n}(w)=\frac{1}{\log ^{b} n} \sum_{i=1}^{\log ^{b} n} C_{n}\left(g_{i} \circ w\right)
$$

where $g_{1}, \ldots, g_{n^{r}}$ are randomly chosen from $\mathrm{G}^{n}$. Using the Chernoff Bound $b$ is chosen large enough such that:

$$
\operatorname{Pr}_{w \in \mathcal{F}^{n}}\left[\left\lvert\, A_{n}(w)-\underset{C_{n}}{\left.\operatorname{Pr}(G \circ w) \left\lvert\, \geq \frac{1}{\log ^{4 c} n}\right.\right] \leq \frac{1}{n^{2}}, ~, ~ . ~}\right.\right.
$$

and

$$
\operatorname{Pr}_{h^{n} \in H^{n}}\left[\left\lvert\, A_{n}\left(h^{n}\right)-\underset{C_{n}}{\left.\operatorname{Pr}\left(G \circ h^{n}\right) \left\lvert\, \geq \frac{1}{\log ^{4 c} n}\right.\right] \leq \frac{1}{n^{2}} . . . . ~}\right.\right.
$$

By Lemma 5.5 we derandomize $A_{n}$ to get $B_{n}$ so that for all but $\frac{1}{n}$ of the $w \in \mathcal{F}^{n}$ :

$$
\begin{equation*}
\left|B_{n}(w)-\operatorname{Pr}_{n}\left(G^{n} \circ w\right)\right|<\frac{1}{\log ^{4 c} n} \tag{5.5}
\end{equation*}
$$

and for all but $\frac{1}{n}$ of the $h^{n} \in \mathrm{H}^{n}$ :

$$
\begin{equation*}
\left|B_{n}\left(h^{n}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(G^{n} \circ h^{n}\right)\right|<\frac{1}{\log ^{4 c} n} . \tag{5.6}
\end{equation*}
$$

We now break the proof into the two cases mentioned earlier. The first case covers the situation in which there is a clear separation between the value output by $B_{n}$ on functions from $\mathcal{F}^{n}$ versus functions from $H^{n}$. The second case handles the random coin flipping.

Let $K_{n}^{\prime}=\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)+\bar{\delta}_{n}\left(1-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right)$, and let $L_{n}^{\prime}=\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)-\bar{\delta}_{n} \operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)$. We note that in the following proof that the variable $K_{n}$ seems to be used many times, but there appears to be little justification for the variable $L_{n}^{\prime}$. We note that it replaces the use of $K_{n}^{\prime}$ in the symmetric version of this proof, and therefore it is included in this version of the proof as well.
Case 1: $\operatorname{Pr}_{k^{n} \in H^{n}}\left[B_{n}\left(h^{n}\right) \geq K_{n}^{\prime}\right] \geq \bar{\epsilon}_{n}=\epsilon(n)+\frac{1}{\log ^{c} n}$
We create the decision circuit $\bar{B}_{n}(w)$ which accepts $w$ iff $B_{n}(w) \geq K_{n}^{\prime}$.

$$
\begin{aligned}
\left|\operatorname{Pr}_{B_{n}}\left(H^{n}\right)-\operatorname{Pr}\left(\mathcal{F}^{n}\right)\right| & \geq \epsilon(n)+\frac{1}{\log ^{c} n}-\operatorname{Pr}_{B_{n}}\left(\mathcal{F}^{n}\right) \\
& \geq \epsilon(n)+\frac{1}{\log ^{c} n}-\frac{1}{n}-\frac{1}{n} \quad \text { (Equation } 5.2 \text { \& sampling error) } \\
& \geq \epsilon(n)+\frac{1}{\log ^{c+1} n}
\end{aligned}
$$

Notice that the second inequality follows from constraints on the probability that a random function $w \in \mathcal{F}^{n}$ is in the set $S^{n}$ defined in equation 5.2. This is because, for all sufficiently large $n, K_{n}^{\prime}$ is greater than the value of $K_{n}$. As previously seen, the probability that a random $w \in \mathcal{F}^{n}$ has the property that $\operatorname{Pr}_{C_{n}}\left(G^{n} \circ w\right)>K_{n}$, is the probability that $w$ is contained in the set $S^{n}$. But, we know that the probability of a random $w \in \mathcal{F}^{n}$ being in the set $S^{n}$ is less than $\frac{1}{n}$. Clearly it follows that the probability that a random $w \in \mathcal{F}^{n}$ has the property that $\operatorname{Pr}_{C_{n}}\left(G^{n} \circ w\right)>K_{n}^{\prime}$ is smaller than $\frac{1}{n}$, for the set of functions with this property is necessarily a subset of $S^{n}$. Finally, since $B_{n}(w)$ approximates $\operatorname{Pr}_{C_{n}}\left(G^{n} \circ w\right)$ to within a value of $\frac{1}{\log ^{1+c} n}$, for all but $\frac{1}{n}$ of the $w \in \mathcal{F}^{n}$, and since it still holds that for all sufficiently large $n$ that $K_{n}^{\prime}-\frac{1}{\log ^{4 c_{n}}}$ is greater than $K_{n}$, there can only be a fraction $\frac{2}{n}$ of the $w \in \mathcal{F}^{n}$ for which $B^{\prime}(w)$ accepts. The first $\frac{1}{n}$ of these functions are the ones in $S^{n}$. The last $\frac{1}{n}$ of these functions are the functions $w$ in which $B(w)$ 's approximation of $\operatorname{Pr}_{C_{n}}\left(G^{n} \circ w\right)$ is off by more than a factor of $\frac{1}{\log ^{\frac{1}{4 c} n}}$.

We let $\Upsilon_{n}=\bar{B}_{n}$, and we are finished Case 1.
Case 2: $\operatorname{Pr}_{h^{n} \in H^{n}}\left[B_{n}\left(h^{n}\right) \geq K_{n}^{\prime}\right]<\bar{\epsilon}_{n}=\epsilon(n)+\frac{1}{\log ^{\varepsilon} n}$
Let $q(w)=B_{n}(w)$ and let

$$
q^{\prime}(w)= \begin{cases}K_{n}^{\prime} & \text { if } K_{n}^{\prime}<B_{n}(w) \\ q(w) & \text { if } L_{n}^{\prime} \leq B_{n}(w) \leq K_{n}^{\prime} \\ L_{n}^{\prime} & \text { if } B_{n}(w)<L_{n}^{\prime}\end{cases}
$$

Let

$$
p(w)=\frac{q(w)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)}{\delta}+\underset{C_{n}}{\operatorname{Pr}\left(\mathcal{F}^{n}\right) \text { and let } p^{\prime}(w)=\frac{q^{\prime}(w)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)}{\bar{\delta}_{n}}+\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) . . . . . .}
$$

We use $B_{n}$ to create a probabilistic decision circuit $D_{n}$, which accepts $w$ with probability $p^{\prime}(w)$. We will show that $D_{n}$ distinguishes between $H^{n}$ and $\mathcal{F}^{n}$ with probability at least $\epsilon(n)+\frac{1}{\log ^{2 c} n}$, proving the theorem. However, we first make some observations about the circuit $B_{n}$ which we will use later in the proof. Notice that,

$$
\left|E_{w \in \mathcal{F n}}\left(B_{n}(w)\right)-\operatorname{Pr}_{n}\left(G^{n} \circ \mathcal{F}^{n}\right)\right| \leq \frac{1}{\log ^{4 c} n}+\frac{1}{n} \quad \text { and }
$$

These facts follow from equations (5.5) and (5.6) respectively. We use the first inequality to get an upper bound on $\operatorname{Pr}_{D_{n}}\left(\mathcal{F}^{n}\right)$, and the second to get a lower bound on $\operatorname{Pr}_{D_{n}}\left(H^{n}\right)$.

We observe that $\operatorname{Pr}_{D_{n}}\left(\mathcal{F}^{n}\right)=E_{w \in \mathcal{F}^{n}}\left(p^{\prime}(w)\right)$. We first produce an upper bound on the expectation of $p(w)$, and then we compensate for the possible discrepancy between it and $p^{\prime}(w)$.

$$
\begin{aligned}
E_{w \in \mathcal{F n}}(p(w)) & =\frac{E_{w \in \mathcal{F} n}(q(w))-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)}{\bar{\delta}_{n}}+\operatorname{Pr}\left(\mathcal{F}^{n}\right) \\
& =\frac{E_{w \in \mathcal{F}^{n}}\left(B_{n}(w)\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)}{\bar{\delta}_{n}}+\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right) \\
& \leq \frac{\operatorname{Pr}_{C_{n}}\left(G^{n} \circ \mathcal{F}^{n}\right)+\frac{1}{\log ^{1 / C_{n}}}+\frac{1}{n}-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)}{\bar{\delta}_{n}}+\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right) \\
& =\frac{\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)+\frac{1}{\log ^{4 c_{n}}}+\frac{1}{n}-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)}{\bar{\delta}_{n}}+\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right) \\
& =\frac{\frac{1}{\log ^{4 C_{n}}}+\frac{1}{n}}{\bar{\delta}_{n}}+\underset{C_{n}}{\operatorname{Pr}\left(\mathcal{F}^{n}\right)} \\
& \leq \operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)+\frac{1}{\log ^{3 c_{n}} n}
\end{aligned}
$$

We need to take into account the possibility that $p(w)<0$, for this can never happen with $p^{\prime}(w)$. This is because $q(w)$ may be smaller than $L_{n}^{\prime}$ whereas this cannot be the case with $q^{\prime}(w)$. Fortunately, by Equation 5.3 , we know that the probability that $q(w)<L_{n}^{\prime}$ is non-significant.

$$
\begin{aligned}
\operatorname{Pr}_{B_{n}}\left(\mathcal{F}^{n}\right) & =E_{w \in \mathcal{F} n}\left(p^{\prime}(w)\right) \\
& \leq E_{w \in \mathcal{F} n}(p(w))+\frac{1}{n}+\frac{1}{n} \quad \text { (Equation } 5.3 \text { and sampling error) } \\
& \leq \operatorname{Pr}\left(\mathcal{F}^{n}\right)+\frac{1}{\log ^{3 c} n}+\frac{1}{n}+\frac{1}{n} \\
& \leq \operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)+\frac{1}{\log ^{2 c} n}
\end{aligned}
$$

We now give a lower bound for $\operatorname{Pr}_{D_{n}}\left(H^{n}\right)=E_{h^{n} \in H^{n}}\left(p^{\prime}\left(h^{n}\right)\right)$. This is done by showing a lower bound for $E_{h^{n} \in \mathrm{H}^{n}}\left(p\left(h^{n}\right)\right)$, and then compensating for the possible difference between $p\left(h^{n}\right)$ and $p^{\prime}\left(h^{n}\right)$.

$$
\begin{aligned}
& E_{h^{n} \in \mathrm{H}^{n}}\left(p\left(h^{n}\right)\right)=\frac{E_{h^{n} \in \mathrm{H}^{n}}\left(q\left(h^{n}\right)\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)}{\bar{\delta}_{n}}+\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right) \\
& =\frac{E_{h^{n} \in \mathbf{H}^{n}}\left(B_{n}\left(h^{n}\right)\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)}{\bar{\delta}_{n}}+\operatorname{P}_{C_{n}}\left(\mathcal{F}^{n}\right) \\
& \geq \frac{\operatorname{Pr}_{C_{n}}\left(G^{n} \circ H^{n}\right)-\frac{1}{\log ^{1 c_{n}}}-\frac{1}{n}-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)}{\delta_{n}}+\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) \\
& \geq \frac{\bar{\epsilon}_{n} \bar{\delta}_{n}\left(2-\bar{\delta}_{n}\right)-\frac{1}{\log ^{\frac{1}{c} c_{n}}}-\frac{1}{n}}{\bar{\delta}_{n}}+\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) \\
& \geq \bar{\epsilon}_{n}\left(2-\bar{\delta}_{n}\right)-\frac{1}{\bar{\delta}_{n} \log ^{4 c} n}-\frac{1}{\bar{\delta}_{n} n}+\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) \\
& \geq \bar{\epsilon}_{n}\left(2-\bar{\delta}_{n}\right)-\frac{1}{\log ^{3 c} n}+\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)
\end{aligned}
$$

We compensate for the $h^{n} \in H^{n}$ such that $p\left(h^{n}\right)>1$. This occurs when $q\left(h^{n}\right)>K^{\prime}$, and thus by virtue of being in case 2, this can only occur for a fraction $\bar{\epsilon}_{n}=\epsilon(n)+\frac{1}{\log ^{c} n}$ of the $h^{n} \in \mathrm{H}^{n}$.

By Equation 5.4 we know that $\forall h_{k}^{n} \in \mathrm{H}^{n}\left|\operatorname{Pr}_{C_{n}}\left(\mathrm{G} \circ h_{k}^{n}\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right|<\delta+\frac{1}{n}$. Note that $\operatorname{Pr}_{C_{n}}\left(\mathrm{G}^{n} \circ h_{k}^{n}\right)$ is within $\frac{1}{\log ^{1+c_{n}}}$ of $B_{n}\left(h^{n}\right)=q\left(h^{n}\right)$, for all but $\frac{1}{n}$ of the $h^{n} \in H^{n}$. Further, the $\operatorname{Pr}_{C_{n}}\left(\mathrm{G}^{n} \circ h_{k}^{n}\right)$ is at most one. It follows that $\left|q\left(h^{n}\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right|<\delta+\frac{1}{\log ^{4} c_{n}}+\frac{1}{n}$, for most $h^{n} \in H^{n}$. We conclude that for all but $\frac{1}{n}$ of the $h^{n} \in H^{n}$ :

$$
q(h)<\min \left\{\left(1+\frac{1}{\log ^{4 c} n}+\frac{1}{n}\right),\left(\underset{C_{n}}{\left.\left.\operatorname{Pr}\left(\mathcal{F}^{n}\right)+\delta+\frac{1}{\log ^{4 c} n}+\frac{1}{n}\right)\right\} . . . . . .}\right.\right.
$$

We show that in both of the above cases the distance between $q\left(h^{n}\right)$ and $q^{\prime}\left(h^{n}\right)=K_{n}^{\prime}$ is less than $\bar{\delta}_{n}(1-\delta)$.

Sub Case 1) $\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right) \leq(1-\delta)$
It must be the case that $q\left(h^{n}\right)<\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)+\delta+\frac{1}{\log ^{\frac{1}{c} n}}+\frac{1}{n}$ and $q^{\prime}\left(h^{n}\right)=K_{n}^{\prime}=\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)+$ $\bar{\delta}_{\mathrm{n}}\left(1-\operatorname{Pr}_{C_{\mathbf{n}}}\left(\mathcal{F}^{n}\right)\right)$.The result follows:

$$
q\left(h^{n}\right)-q^{\prime}\left(h^{n}\right)<\left(\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)+\delta+\frac{1}{\log ^{4 c} n}+\frac{1}{n}\right)-\left(\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)+\bar{\delta}_{n}\left(1-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right)\right.
$$

$$
\begin{aligned}
& =\left(\delta+\frac{1}{\log ^{4 c} n}+\frac{1}{n}\right)-\left(\delta+\frac{1}{\log ^{c} n}\right)\left(1-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right) \\
& =\left(\delta+\frac{1}{\log ^{4 c} n}+\frac{1}{n}\right)-\delta+\delta \operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)-\frac{1}{\log ^{c} n}+\frac{1}{\log ^{c} n}{\underset{C}{n}}^{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) \\
& =\left(\frac{1}{\log ^{4 c} n}+\frac{1}{n}\right)+\delta \underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)-\frac{1}{\log ^{c} n}+\frac{1}{\log ^{c} n} \operatorname{Pr}_{n}\left(\mathcal{F}^{n}\right) \\
& <\delta \operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)+\frac{1}{\log ^{c} n} \operatorname{Pr}_{\boldsymbol{n}}\left(\mathcal{F}^{n}\right) \\
& =\bar{\delta}_{n}{\underset{C}{n}}^{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) \\
& <\bar{\delta}_{n}(1-\delta) \text {. }
\end{aligned}
$$

Sub Case 2) $\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)>(1-\delta)$
In this case there must exist a $u$ such that $0<u \leq \delta$ where $\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)=1-\delta+u$. The result follows:

$$
\begin{aligned}
q(h)-q^{\prime}(h) & <1+\frac{1}{\log ^{4 c} n}+\frac{1}{n}-K_{n}^{\prime} \\
& =\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)+\delta-u+\frac{1}{\log ^{4 c} n}+\frac{1}{n}-K_{n}^{\prime} \\
& =\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)+\delta-u+\frac{1}{\log ^{4 c} n}+\frac{1}{n}-\left(\underset{C_{n}}{\operatorname{Pr}\left(\mathcal{F}^{n}\right)}+\bar{\delta}_{n}\left(1-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right)\right. \\
& =-u+\frac{1}{\log ^{4 c} n}+\frac{1}{n}-\left(\bar{\delta}_{n}\left(1-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right)\right. \\
& =-u+\frac{1}{\log ^{4 c} n}+\frac{1}{n}-\left(\bar{\delta}_{n}-\bar{\delta}_{n}{\underset{C r}{n}}^{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right) \\
& =-u+\frac{1}{\log ^{4 c} n}+\frac{1}{n}-\left(\delta+\frac{1}{\log ^{c} n}-\bar{\delta}_{n}{\underset{C r}{n}}\left(\mathcal{F}^{n}\right)\right) \\
& =-u-\delta+\bar{\delta}_{n} \operatorname{Pr}\left(\mathcal{F}^{n}\right) \\
& =-u-\delta+\bar{\delta}_{n}(1-\delta+u) \\
& =-u-\delta+\bar{\delta}_{n}(1-\delta)+\bar{\delta}_{n} u \\
& =\bar{\delta}_{n}(1-\delta)+\bar{\delta}_{n} u-u-\delta \\
& \leq \bar{\delta}_{n}(1-\delta) .
\end{aligned}
$$

Using the bound on the distance between $p\left(h^{n}\right)$ and $p^{\prime}\left(h^{n}\right)$ derived above and applying it to the previous lower bound for $\operatorname{Pr}_{D_{n}}\left(H^{n}\right)$ we get that for all but $\frac{1}{n}$ of $h^{n} \in H^{n}$ :

$$
\begin{aligned}
{\underset{D}{n}}^{\operatorname{Pr}}\left(H^{n}\right) & =E_{h^{n} \in H^{n}}\left(p^{\prime}\left(h^{n}\right)\right) \\
& \geq E_{h^{n} \in H^{n}}\left(p\left(h^{n}\right)\right)-\left(\epsilon(n)+\frac{1}{\log ^{c} n}\right)(1-\delta) \\
& =E_{h^{n} \in H^{n}}\left(p\left(h^{n}\right)\right)-\left(\bar{\epsilon}_{n}\right)(1-\delta)
\end{aligned}
$$

For the $\frac{1}{n}$ of the $h^{n} \in H^{n}$ that are not within the distance given by the Chernoff bound, we know that they differ by at most 1 from the value being estimated, and therefore can conclude:

$$
\begin{aligned}
{\underset{D r}{n}}^{\operatorname{Pr}\left(H^{n}\right)} & =E_{h^{n} \in \mathcal{H}^{n}}\left(p^{\prime}\left(h^{n}\right)\right) \\
& \geq E_{h^{n} \in H^{n}}\left(p\left(h^{n}\right)\right)-\left[\bar{\epsilon}_{n}\right][(1-\delta)]-\frac{1}{n} \\
& =\left[\bar{\epsilon}_{n}\left(2-\bar{\delta}_{n}\right)-\frac{1}{\log ^{3 c} n}+\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right]-\left[\bar{\epsilon}_{n}\right][(1-\delta)]-\frac{1}{n} \\
& =\left[\bar{\epsilon}_{n}\left(2-\bar{\delta}_{n}-1+\delta\right)-\frac{1}{\log ^{2 c} n}+\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right] \\
& =\left[\bar{\epsilon}_{n}\left(1-\bar{\delta}_{n}+\delta\right)-\frac{1}{\log ^{2 c} n}+\underset{C_{n}}{\left.\operatorname{Pr}\left(\mathcal{F}^{n}\right)\right]}\right. \\
& =\left[\bar{\epsilon}_{n}\left(1-\delta-\frac{1}{\log ^{c} n}+\delta\right)-\frac{1}{\log ^{2 c} n}+\underset{C_{n}}{\left.\operatorname{Pr}\left(\mathcal{F}^{n}\right)\right]}\right. \\
& =\bar{\epsilon}_{n}\left(1-\frac{1}{\log ^{c} n}\right)-\frac{1}{\log ^{2 c} n}+\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right) .
\end{aligned}
$$

By taking the difference between the lower bound on the acceptance rate of $D_{n}$ on $\mathrm{H}^{n}$ and the upper bound on the acceptance rate of $D_{n}$ on $\mathcal{F}^{n}$ we achieve the proper distinguishing probability.

$$
\begin{aligned}
\operatorname{Pr}_{n}\left(H^{n}\right)-\underset{D_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) & \geq\left[\bar{\epsilon}_{n}\left(1-\frac{1}{\log ^{c} n}\right)-\frac{1}{\log ^{2 c} n}+\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right]-\left[\operatorname{Pr}\left(\mathcal{F}^{n}\right)+\frac{1}{\log ^{2 c} n}\right] \\
& \geq \bar{\epsilon}_{n}\left(1-\frac{1}{\log ^{c} n}\right) \\
& =\left(\epsilon(n)+\frac{1}{\log ^{c} n}\right)\left(1-\frac{1}{\log ^{c} n}\right) \\
& =\left(\epsilon(n)-\frac{\epsilon(n)}{\log ^{c} n}+\frac{1}{\log ^{c} n}-\frac{1}{\log ^{2 c} n}\right) \\
& \geq \epsilon(n)+\frac{1}{\log ^{2 c} n}
\end{aligned}
$$

All that remains is to derandomize $D_{n}$. In this case we let $\Upsilon_{n}$ be the derandomized version of $D_{n}$. We have now constructed the family of circuits $\left\{\Upsilon_{n}\right\}$ which was required to prove the lemma.

### 5.4 Towards Complete Security

Our goal is to show that by composing enough partially secure PRPGs together we can construct a PRPG which is completely secure. So far our efforts along these lines have failed. We will now discuss some possible future research directions which will allow us to attain our goal.

Observe that although the Isolation Lemma is tight, we have been unable to find a set of three permutation generators which when composed together are as insecure as is permitted by the direct application of the Isolation Lemma. In particular, in our previous example we showed that there are generators $G$ which are $1-\delta$ secure, and generators $H$ which are $1-\epsilon$ secure, but whose composition $G \circ H$ is $\epsilon \delta(2-\max \{\epsilon, \delta\})$ secure. However, a crucial requirement of this example is that a fraction $\delta$ of the keys in $G$ correspond to the identity permutation, and a fraction $\epsilon$ of the keys in $H$ correspond to the identity permutation. However, only a fraction $\epsilon \delta$ of the keys for $\mathrm{G} \circ \mathrm{H}$ correspond to the
identity permutation, whereas its security is $\epsilon \delta(2-\max \{\epsilon, \delta\})$. This fact prevents us from composing $G \circ H$ with a $1-\gamma$ secure generator $F$ to produce a generator $\mathrm{Fo}(\mathrm{GoH})$ which is $\epsilon \delta \gamma(2-\max \{\epsilon, \delta\})(2-\max \{\epsilon, \delta, \gamma\})$ secure. The observation we make is that the fraction of keys for $\mathrm{G} \circ \mathrm{H}$ which correspond to permutations which are easily distinguishable from random functions is much smaller than the resulting security of $G \circ H$. Further, it is this lack of "special", and therefore easily identifiable, permutations in $\mathrm{G} \circ \mathrm{H}$ which prevents us from composing the generator with a third, and achieving the security bound specified by the Isolation Lemma.

We conjecture that the security of the composition of three or more partially secure PRPGs is stronger than the bound which results from multiple applications of the Isolation Lemma. However, we suspect that the current notion of security may be too coarse to prove such a statement. Therefore, we propose a finer notion of security.

Definition 5.1 (Weak Security) We say that a PRPG (PRFG) G is $1-\delta(n)$ weaksecure if for all adversaries $A^{f}$ which have the property that for all $c>0$, and for all sufficiently large $n$ :

$$
\operatorname{Pr}_{f \in \mathcal{F} n}\left[A^{f}=1\right] \leq \frac{1}{n^{c}} ;
$$

then it is the case that for all sufficiently large $n$ and all $c>0$ :

$$
\left|\operatorname{Pr}_{f \in \mathcal{F}^{n}}\left[A^{f}=1\right]-\operatorname{Pr}_{g \in \mathrm{G}^{n}}\left[A^{g}=1\right]\right| \leq \delta(n)+\frac{1}{n^{c}} .
$$

We note that by some minor modifications to the arguments presented in this chapter, it is easy to show that the composition of a $1-\delta$ weakly secure PRPG and a $1-\epsilon$ weakly secure PRPG is $1-\delta \epsilon$ weakly-secure (This follows from the fact that sub case 2 of case 2 in section 5.3.1 disappears). This shows that there is no way of increasing the fraction of permutations which are easily identifiable from random in the composition of two partially secure PRPGs. We suspect that further work in this direction may point the way to developing a theorem which shows that a polynomial in $\boldsymbol{n}$ number of compositions of a $1-\delta$ partially secure PRPG results in a 1 secure PRPG.

## Chapter 6

## The Luby-Rackoff Operator

In Chapter 5 we showed how to take partially secure PRPGs and compose them to produce a PRPG which is more secure than any of the constituent PRPGs. In this chapter we will show that an analog is possible for PRFGs. As was demonstrated in Chapter 4 , this cannot be done directly by composing PRFGs, and so we introduce a new operator $\square$ (read as box). We will show that the operator acting on functions has some properties in common with the composition operator acting on permutations.

Given two functions $f_{1}$ and $f_{2}$ such that $f_{i}:\{0,1\}^{n} \rightarrow\{0,1\}^{n}$, for $i \in\{1,2\}$, we define the operator $\square$ as:

$$
\left(f_{1} \square f_{2}\right)(x)=f_{1}\left(f_{1}(x) \oplus f_{2}(x)\right) \oplus f_{2}\left(f_{1}(x) \oplus f_{2}(x)\right)
$$

The operator was originally defined in some working notes by Luby and Rackoff [13] in which an incomplete argument is given to show that two partially secure PRFGs can be combined by the $\square$ operator, so that the resulting PRFG is more secure than either of its constituents. We will now complete the argument, and show how these results relate to the arguments presented in Chapter 5.

### 6.1 Two Technical Lemmas

In Chapter 5 we presented two Isolation Lemmas which proved that the composition of two partially secure PRPGs resulted in a third which was more secure than either of the constituent generators. However, in both proofs the only arguments which made use of specific properties of the composition operator or the permutation property of the PRPGs were contained in Lemmas 5.6 and 5.7. Therefore, if we can prove corresponding lemmas for the $\square$ operator and PRFGs, then equivalent Isolation Lemmas will hold for the operator acting on PRFGs.

Before proving these correspondences we standardize the circuits that will be used in the remainder of this chapter. Observe that whenever we have a circuit $C_{n}$ which has oracle-gates which correspond to a function $f$, it can be easily modified, by adding at most a polynomial number of gates and connections between gates, to render a new circuit $\widehat{C_{n}}$, where $\widehat{C_{n}}$ never repeats queries to $f$ on the same input. Further, any polynomial sized family of circuits $\left\{C_{n}\right\}$, in which each circuit $C_{n}$ has at most $m(n)$ oracle gates, can be easily be modified to a equivalent family of circuits $\left\{\bar{C}_{n}\right\}$, where each circuit $\overline{C_{n}}$ has exactly $m(n)$ oracle gates. We shall assume that all circuits and circuit families in this chapter are of the modified forms described above, unless specifically mentioned otherwise.

We now state the first property of the operator, which we will use to develop one of the above mentioned correspondences.

Lemma 6.1 Given any family of decision-circuits $\left\{C_{n}\right\}$, for each $f \in \mathcal{F}^{n}$,

$$
\left|\operatorname{Pr}\left(f \square \mathcal{F}^{n}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \leq \frac{2 m^{2}(n)}{2^{n}}
$$

where $m(n)$ is the number oracle gates in the circuit $C_{n}$.

Intuitively the lemma shows that the operator maintains the security of its strongest operand, for there is a trivial circuit which distinguishes $f$, but by the above lemma
there can be no such circuit for $f \square \mathcal{F}^{n}$. For completeness we give the proof of Lemma 6.1 originally presented by Luby and Rackoff in [13].

Proof of Lemma 6.1: We will perform two experiments which will result in two random variables $R_{1}$ and $R_{2}$. We will show that $E\left[R_{1}\right]=\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)$, and then show that $E\left[R_{2}\right]$ differs from $\operatorname{Pr}_{C_{n}}\left(f \square \mathcal{F}^{n}\right)$ by at most a negligible amount. Finally, we will observe that $R_{1}=R_{2}$, proving the result.

In the first experiment we let $X_{1}, \ldots, X_{m(n)}$ be i.i.d.r.v. from the uniform distribution over $\{0,1\}^{n}$. Let $g_{1}, \ldots, g_{m(n)}$ be the oracle gates in $C_{n}$. We define $R_{1}$ to be the output of the circuit $C_{n}$ when we replace the output of gate $g_{i}$ with the value $X_{i}$. Let $\alpha_{i}$ be the input to oracle gate $g_{i}$. Note that for $i$, such that $1<i \leq m(n)$, that $\alpha_{i}$ is completely determined by $C_{n}$ and $X_{1}, \ldots, X_{i-1}$. This taken in conjunction with the fact that for $i \neq j$ that $\alpha_{i} \neq \alpha_{j}$ and by the construction of $C_{n}$, we observe that $\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)=E\left[R_{1}\right]$.

For the second experiment, observe that $\operatorname{Pr}_{C_{n}}\left(f \square \mathcal{F}^{n}\right)=\operatorname{Pr}_{f^{\prime} \in \mathcal{F} n}\left[C_{n}\left(f^{\prime} \circ f^{\prime}\right)=1\right]$. This follows from two observations: first $\left(f_{1} \square f_{2}\right)(x)=\left(f_{1} \oplus f_{2}\right) \circ\left(f_{1} \oplus f_{2}\right)(x)$; secondly, for any $\hat{f} \in \mathcal{F}^{n}$ the set $\hat{f} \oplus \mathcal{F}^{n}$ is isomorphic to $\mathcal{F}^{n}$. We construct the circuit $\widehat{C}_{n}$ by replacing each oracle-gate $g_{i}$ in the circuit $C_{n}$ with two oracle-gates $h_{i}$ and $h_{i}^{\prime}$. We feed the original input of $g_{i}$ into $h_{i}$; feed the output of $h_{i}$ into $h_{i}^{\prime}$; and let the output of $h_{i}^{\prime}$ replace the original output of $g_{i}$. We now let $X_{1}, \ldots, X_{m(n)}$ and $Y_{1}, \ldots, Y_{m(n)}$ be i.i.d.r.v. from the uniform distribution over $\{0,1\}^{n}$. We define $R_{2}$ to be the the output of $\widehat{C}_{n}$ when we replace the output of gate $h_{i}$ with $Y_{i}$, and the output of $h_{i}^{\prime}$ with $X_{i}$, for $1 \leq i \leq m(n)$. Let $\alpha_{i}$ be the input to gate $h_{i}$, and notice that $\alpha_{i}$ is completely determined by $\widehat{C}_{n}$ and $X_{1}, \ldots, X_{i-1}$. Clearly $E\left[R_{1}\right]=E\left[R_{2}\right]$. But we will show that $E\left[R_{2}\right]$ is negligibly close to $\operatorname{Pr}_{C_{n}}\left(f \square \mathcal{F}^{n}\right)$. We consider the set

$$
\text { good }=\left\{X_{1}, \ldots, X_{m(n)}, Y_{1}, \ldots, Y_{m(n)} \mid \text { the relation }\left\{\left(\alpha_{i}, Y_{i}\right)\right\} \cup\left\{\left(Y_{i}, X_{i}\right)\right\} \text { is one-to-one }\right\}
$$

The set contains all of the choices of $X_{1}, \ldots, X_{m(n)}, Y_{1}, \ldots, Y_{m(n)}$ which define a partial function in the above experiment. We then define the set bad to be the complement of
the set good. Notice that:

$$
\begin{aligned}
\operatorname{Pr}_{C_{n}} \operatorname{Pr}\left(f \square \mathcal{F}^{n}\right) & =\operatorname{Pr}_{f^{\prime} \in \mathcal{F}^{n}}\left[C_{n}\left(f^{\prime} \circ f^{\prime}\right)=1\right] \\
& =\operatorname{Pr}_{f^{\prime} \in \mathcal{F}^{n}}\left[\widehat{C}_{n}\left(f^{\prime}\right)=1\right] \\
& =E_{f^{\prime} \in \mathcal{F} n}\left[\widehat{C}_{n}\left(f^{\prime}\right)\right] \\
& =E_{t=\left(X_{\mathrm{L}}, \ldots, X_{m(n)}, Y_{1}, \ldots, Y_{m(n)}\right) \in\{0,1\}^{n \times 2 m(n)}}\left[R_{2} \mid t \in \operatorname{good}\right]
\end{aligned}
$$

Finally, observe that a tuple $\left(X_{1}, \ldots, X_{m(n)}, Y_{1}, \ldots, Y_{m(n)}\right)$ is definitely in good if $\alpha_{1}, . ., \alpha_{m(h)}, Y_{1}, \ldots, Y_{m(n)}$ are all distinct. But by the construction of $\widehat{C}_{n}, \alpha_{i} \neq \alpha_{j}$ for $i \neq j$, and thus the probability that these values are not all distinct is bounded above by $\frac{2 \mathrm{~m}^{2}(n)}{2^{n}}$. Observing that:

$$
\begin{aligned}
\underset{C_{n}}{\operatorname{Pr}\left(\mathcal{F}^{n}\right)} & =E\left[R_{2}\right] \\
& =E\left[R_{2} \mid \text { good }\right] \operatorname{Pr}[\text { good }]+E\left[R_{2} \mid \text { bad }\right] \operatorname{Pr}[\text { bad }] \\
& \leq \operatorname{Pr}\left(f \square \mathcal{F}^{n}\right) \operatorname{Pr}[\text { good }]+E\left[R_{2} \mid \text { bad }\right] \frac{2 m^{2}}{2^{n}} \\
& \leq \underset{C_{n}}{\operatorname{Pr}}\left(f \square \mathcal{F}^{n}\right)+\frac{2 m^{2}(n)}{2^{n}},
\end{aligned}
$$

the result follows.

Corollary 6.1 For each $n$ let $S^{n} \subset \mathcal{F}^{n}$, and let $\left\{C_{n}\right\}$ be a family of circuits, then

$$
\left|\operatorname{Pr}_{C_{n}}\left(S^{n} \square \mathcal{F}^{n}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right|<\frac{2 m^{2}(n)}{2^{n}}
$$

where $m(n)$ is the number oracle gates in the circuit $C_{n}$.
Notice the similarities between Corollary 6.1 and Lemma 5.6. In effect, except for a negligibly small chance of error, we have proven the equivalent of Lemma 5.6 except that the composition operator and a set of permutations has been replaced by the operator and a set of functions, respectively. However, because the probability of error is negligible it does not have an effect on arguments made in the proofs of the Isolation Lemmas. Therefore, we can intuitively treat $\operatorname{Pr}_{C_{n}}\left(S^{n} \square \mathcal{F}^{n}\right)=\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)$.

We now present the second property of the $\square$ operator. As we used Lemma 5.8 to prove Lemma 5.7 , we will similarly use a lemma, which corresponds to Lemma 5.8 , to help prove a Lemma which corresponds to Lemma 5.7. In order to prove the lemma which corresponds to Lemma 5.8, we first need to prove the lemma given below.

Lemma 6.2 For every $n$ let $\widehat{S}^{n} \subset \mathcal{F}^{n}$ be a set such that $\frac{\left|\widehat{S}^{n}\right|}{\left|\bar{n}^{n}\right|} \geq \frac{1}{n^{c}}$, for some constant $c>0$.Let $\left\{B_{n}\right\}$ be an infinite family of polynomial in $n$ sized oracle circuits. For each $n$, let $B_{n}$ have exactly $m(n)$ oracle gates. Then for all $k$ and all sufficiently large $n$

$$
\operatorname{Pr}_{\hat{3} \in \hat{\mathcal{S}}^{n}}\left[B_{n}(\hat{s} \circ \hat{s})=\left(\eta_{k}, \hat{s}\left(\eta_{k}\right)\right)\right] \leq \frac{n^{c} \cdot m(n)^{3}}{2^{n}}
$$

where $\eta_{k}$ is the $k^{\text {th }}$ query made to a function oracle gate by $B_{n}$ when given $\hat{s} \circ \hat{s}$ as an oracle.

Proof: We first give a bound on the probability when $\hat{s}$ is chosen uniformly at random from the set $\mathcal{F}^{n}$. Later, we use this bound to derive one in which $\hat{s}$ is considered to be chosen from the set $\widehat{S}^{n}$.

Let $B_{n}$ be a polynomial sized oracle circuit which has $m(n)$ oracle gates. Further, suppose that $B_{n}$ never performs the same query twice. Let $\Gamma_{1}, \ldots, \Gamma_{m(n)}$ be the oracle gates of $B_{n}$. Each gate $\Gamma_{k}$ is supposed to represent the function $\hat{s} \circ \hat{s}$. Therefore, we replace each gate $\Gamma_{k}$ with the gates $\Gamma_{k}^{1}$ and $\Gamma_{k}^{2}$, where these gates each represent the function $\hat{s}$. This is done by taking the input to $\Gamma_{k}$ and redirecting it to $\Gamma_{k}^{1}$; taking the output of $\Gamma_{k}^{1}$ and making it the input of $\Gamma_{k}^{2}$; and replacing the output of $\Gamma_{k}$ with the output of $\Gamma_{k}^{2}$.

We perform an experiment. We choose $\zeta_{1}, \ldots, \zeta_{m(n)}$ uniformly at random from the set $\{0,1\}^{n \times m(n)}$, and we choose $\lambda_{1}, \ldots, \lambda_{m(n)}$ uniformly at random from the set $\{0,1\}^{n \times m(n)}$. We will evaluate $B_{n}$ in two different manners using the random choices. Define the output of gate $\Gamma_{b}^{a}$ to be $\rho_{b}^{a}$. For each $k$, we call $\eta_{k}$ the input to $\Gamma_{k}^{1}$.

For the first evaluation of $B_{n}$ we do the following. For each $k$ we set $\rho_{k}^{2}$ to be $\zeta_{k}$. Notice that by fixing the $\rho_{1}^{2}, \ldots, \rho_{m(n)}^{2}$ we have completely determined $\eta_{1}, \ldots, \eta_{m(n)}$, the inputs to
each oracle gate. Further, we have determined the output of the circuit $B_{n}$, which we will denote the output by the pair ( $\eta_{k}, \mu$ ), where $\mu$ is $B_{n}$ 's guess of $s\left(\eta_{k}\right)$, for some $k$. Now for each $k$ we set $\rho_{k}^{1}$ to be $\lambda_{k}$.

Observe that these choices do not correspond to randomly choosing a function $t \in \mathcal{F}^{\boldsymbol{n}}$, and then running $B_{n}$ on the oracle $t$ o $t$. In fact, it may be the case that these random choices do not even correspond to functions. For example, it may be the case that $\lambda_{j}=\lambda_{k}$, but that $\zeta_{j} \neq \zeta_{k}$, and therefore the choices would not even correspond to a function, as functions are single-valued. However, had these choices corresponded to a random function, then the probability that the output of the circuit, $\left(\eta_{k}, \mu\right)$, was correct, is the probability that $\mu=\lambda_{k}$. We note that the probability that $\mu=\lambda_{k}$ is exactly $\frac{1}{2^{n}}$. Let $\mathcal{E}_{1}$ be the event that $\mu=\lambda_{k}$, during the first evaluation of $B_{n}$.

As stated earlier, it is clearly the case that the above experiment does not correspond to choosing a function $\hat{s}$ uniformly at random from $\mathcal{F}^{n}$, and then using $\hat{s} \circ \hat{s}$ as an oracle for $B_{n}$. We will correct this in the second evaluation of $B_{n}$. Further, we will now bound the number of cases in which the first evaluation fails to correspond to choosing a function $\hat{s} \in \mathcal{F}^{n}$, and then using $\hat{s} \circ \hat{s}$ as an oracle.

We now describe the second evaluation of $B_{n}$, and we note that it simulates $B_{n}$ running on the input $\hat{s} \circ \hat{s}$. We consider the gates of $B_{n}$ in the order $\Gamma_{1}^{1}, \Gamma_{1}^{2}, \ldots, \Gamma_{m(n)}^{1}, \Gamma_{m(n)}^{2}$. We need to set the outputs of the gates $\Gamma_{b}^{a}$ in fashion which is equivalent to how they would be set if we were to truly choose an $\hat{s}$ uniformly at random from $\mathcal{F}^{n}$, and then use $\hat{\boldsymbol{s}} \circ \hat{\boldsymbol{s}}$ as an oracle. To accomplish this we set the outputs of the oracle gates to be the random choices, as before, except when the random choice would not correspond to a function of the form $\hat{s} \circ \hat{s}$. This occurs when the random choice for the output is not consistent with previously assigned random outputs of previous gates. If such an exception arises, then we fix the output of the gate so that it is consistent with the outputs of previous gates. Formally, we set $\rho_{k}^{1}$ to be $\lambda_{k}$, unless there exists a $\rho_{\ell}^{1}=\rho_{k}^{1}$, for $\ell<k$, and in this case we set $\rho_{k}^{1}$ to be $\rho_{\ell}^{2}$. Similarly, we set $\rho_{k}^{2}$ to be $\zeta_{k}$, unless there exists an $\ell<k$ where $\rho_{k}^{1}=\rho_{\ell}^{1}$
or $\rho_{k}^{1}=\eta_{\ell}$, in which case we set $\rho_{k}^{2}$ to be $\rho_{l}^{2}$ or $\rho_{l}^{1}$ respectively.
When there exists a $k$ such that we do not set $\rho_{k}^{1}$ to be $\lambda_{k}$, or we do not set $\rho_{k}^{2}$ to be $\zeta_{k}$ then we say a collision has occurred. Intuitively, this is when a random choice was not consistent with a function of the form $\hat{s} \circ \hat{s}$, and so the output of a gate was set so that it was consistent with previous gates outputs.

Notice that it is only when collisions occur that the first evaluation of $B_{n}$ differs from the second evaluation of $B_{n}$.

Let $\mathcal{E}_{2}$ be the event of a collision occurring in second evaluation of $B_{n}$. We bound the probability of event $\mathcal{E}_{2}$. Clearly, if all of the $\eta$ 's and $\lambda$ 's are distinct, then no collision can occur. We bound this probability.

$$
\begin{aligned}
\operatorname{Pr}_{\substack{\mathcal{C l}_{1}, \ldots, \zeta_{m(n)} \in\left\{0,11^{n \cdot m(n)} \\
\lambda_{1}, \ldots, \lambda_{m(n)} \in\{0,1)^{n \cdot m(n)}\right.}}\left[\left|\left\{\lambda_{j}, \eta_{k} \mid 1 \leq j, k \leq m(n)\right\}\right|<2 m(n)\right] & \leq \frac{m(n)}{2^{n}}+\frac{m(n)+1}{2^{n}}+\cdots+\frac{2 m(n)}{2^{n}} \\
& =\frac{(m(n))^{2}+(1+\cdots+m(n))}{2^{n}} \\
& =\frac{(m(n))^{2}+\left(m(n)^{2}+m(n)\right)}{2^{n+1}} \\
& =\frac{2 m(n)^{2}+m(n)}{2^{n+1}}
\end{aligned}
$$

Define $\mathcal{E}_{3}$ to be the event that $B_{n}$ outputs $\left(\eta_{k}, \mu\right)$, and $\mu=\rho_{k}^{1}$ during the second evaluation of $B_{n}$. Clearly, this is only the case when the output of gate $\Gamma_{k}^{1}$ is equal to $\mu$. Therefore, $\mathcal{E}_{3} \subset \mathcal{E}_{1} \cup \mathcal{E}_{2}$, and this implies that:

$$
\begin{aligned}
\operatorname{Pr} \mathcal{E}_{3} & \leq \operatorname{Pr} \mathcal{E}_{1}+\operatorname{Pr} \mathcal{E}_{2} \\
& =\frac{1}{2^{n}}+\frac{2 m(n)^{2}+m(n)}{2^{n+1}} \\
& =\frac{2+2 m(n)^{2}+m(n)}{2^{n+1}}
\end{aligned}
$$

Notice that by the definition of $\mathcal{E}_{3}$, and the design of the experiment that

$$
\underset{\hat{s} \in \mathcal{F}{ }^{\operatorname{Pr}}}{ }\left[B_{n}(\hat{s} \circ \hat{s})=\left(\eta_{k}, \hat{s}\left(\eta_{k}\right)\right)\right]=\operatorname{Pr}\left[\mathcal{E}_{3}\right] \leq \frac{2+2 m(n)^{2}+m(n)}{2^{n+1}}
$$

We now remember that we are interested in the case in which $\hat{s}$ is chosen uniformly at random for the set $\widehat{S}^{n}$, and not the set $\mathcal{F}^{n}$. We use our bound for the latter case to derive a bound for the former.

$$
\begin{aligned}
\operatorname{Pr}_{\hat{s} \in \hat{S}^{n}}\left[\mathcal{E}_{3}\right] & =\operatorname{Pr}_{\hat{s} \in \mathcal{F}^{n}}\left[\mathcal{E}_{3} \mid \hat{s} \in \widehat{S}^{n}\right] \\
& \leq \frac{\operatorname{Pr}_{\hat{s} \in \mathcal{F}^{n}}\left[\mathcal{E}_{3}\right]}{\operatorname{Pr}_{\hat{s} \in \widehat{S}^{n}}} \\
& =\frac{\frac{2+2 m(n)^{2}+m(n)}{2^{n+1}}}{\frac{1}{n^{c}}} \\
& =\frac{n^{c}\left(2+2 m(n)^{2}+m(n)\right)}{2^{n+1}} \\
& \leq \frac{n^{c} \cdot m(n)^{3}}{2^{n}}
\end{aligned}
$$

We now prove the lemma which corresponds to Lemma 5.8. Remember that this lemma is only used to later prove a lemma which corresponds to Lemma 5.7.

Lemma 6.3 For every $n$ let $S^{n} \subset \mathcal{F}^{n}$ be a set such that $\frac{\left|S^{n}\right|}{\left|\mathcal{F}^{n}\right|} \geq \frac{1}{n^{c}}$, for some constant $c>0$. Let $\left\{C_{n}\right\}$ be an infinite family of polynomial in $n$ sized decision-circuits. We show that for any $d>0$, there exists an $r>0$ such that for all sufficiently large $n$ :

$$
\underset{\left(f, s_{1}, \ldots, s_{n^{r}}\right) \in \mathcal{F}^{n} \times\left(S^{n}\right)^{n^{r}}}{\operatorname{Pr}}\left[\left|\frac{1}{n^{r}} \sum_{i=1}^{n^{r}} C_{n}\left(f \square s_{i}\right)-\underset{C_{n}}{\operatorname{Pr}\left(\mathcal{F}^{n}\right)}\right|>\frac{1}{n^{d}}\right]<\frac{1}{2^{n / 3}}
$$

Proof: We will assume that $r$ is fixed, and later we will show how to determine $r$ 's value.

First we define two experiments. In the first experiment pick random $\left(f_{1}, f_{2}, \ldots, f_{n}^{r}\right) \in$ $\left(\mathcal{F}^{n}\right)^{n^{r}}$, and evaluate $C_{n}\left(f_{i}\right)$ for each $i \in\left\{1, . ., n^{r}\right\}$. Define the event $E_{1}$ to be:

$$
\left|\frac{1}{n^{r}} \sum_{i=1}^{n^{r}} C_{n}\left(f_{i}\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right|>\frac{1}{n^{d}} .
$$

In the second experiment pick random $\left(f, s_{1}, \ldots, s_{n}^{r}\right) \in S^{n} \times\left(\mathcal{F}^{n}\right)^{n^{r}}$, and evaluate $C_{n}\left(f \square_{s_{i}}\right)$ for each $i$, where $1 \leq i \leq n^{r}$. Define the event $E_{2}$ to be:

$$
\left|\frac{1}{n^{r}} \sum_{i=1}^{n^{r}} C_{n}\left(f \square s_{i}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right|>\frac{1}{n^{d}} .
$$

Using the Chernoff Bound (Lemma 5.4) we choose an $r$ where the probability of event $E_{1}$ occurring in the first experiment is less than $\frac{1}{2 \frac{1}{2}}$. We will show that the probability of event $E_{2}$, in the second experiment, is negligibly close to the probability of $E_{1}$, in the first experiment. The lemma will follow.

We will perform a third experiment, in which we model both of the first two experiments by considering two different methods of evaluating the circuit $C_{n}, n^{r}$ times. However, in the third experiment we use a modified form of $C_{n}$. To modify $C_{n}$, we replace every function gate, $g_{j}$, which corresponds to the function $f_{i}$ or $f \square s_{i}$ in experiments one and two respectively, with two function gates, $I_{\boldsymbol{j}}$ and $O_{j}$. In the first experiment the gate $I_{j}$ would correspond to the identity function and $O_{\boldsymbol{j}}$ would correspond to the function $f_{i}$. In the second experiment both $I_{j}$ and $O_{J}$ would correspond to the functions $f \oplus s_{i}$. We redirect $g_{j}$ 's input to $I_{j}$; we take the output of $I_{j}$ and use it as the input for $O_{j}$; and we replace the outputs of $g_{j}$ with the output of $O_{j}$. Notice that the behavior of $C_{n}$ has not been modified. We have simply modified the representation of the circuit. We let $I_{j}^{i}$ and $O_{j}^{i}$ correspond to the gates which replaced $g_{j}$, in the manner previously mentioned, during the evaluation of $C_{n}\left(f_{i}\right)$ and $C_{n}\left(f \square s_{i}\right)$, respectively in our models of experiments one and two.

First we choose $s_{1}, \ldots, s_{n^{r}} \in S^{n} ; \beta_{1}^{1}, \ldots, \beta_{m(n)}^{1}, \ldots, \beta_{1}^{n^{r}}, \ldots, \beta_{m(n)}^{n^{r}} \in\{0,1\}^{n \times m(n) \times n^{r}} ;$ and then we choose $\gamma_{1}^{1}, \ldots, \gamma_{m(n)}^{1}, \ldots, \gamma_{1}^{n^{r}}, \ldots, \gamma_{m(n)}^{n^{r}} \in\{0,1\}^{n \times m(n) \times n^{r}}$. We represent the input to the gate $I_{b}^{a}$ as $\alpha_{b}^{a}$; we represent the input to the gate $O_{b}^{a}$ as $\widehat{\beta}_{b}^{a}=I_{b}^{a}\left(\alpha_{b}^{a}\right)$; and we represent the output of the gate $O_{b}^{a}$ as $\widehat{\gamma}_{b}^{a}$.

We can now model the first experiment in experiment three. Rather than performing $n^{r}$ evaluations of $C_{n}$ using a different random function as the oracle for each evaluation, as in experiment one; we equivalently consider $n^{r}$ evaluations of $C_{n}$, where in each evaluation
$i$ and for each $j$ we assign the random element $\gamma_{i}^{j} \in\{0,1\}^{n}$ to $\widehat{\gamma}_{b}^{a}$, the output of $O_{j}^{i}$. Notice that by fixing the outputs of each gate we have completely determined the behaviour of the circuit. Further, notice that because of the way we have set the outputs to the gates $O_{j}^{i}$, we have perfectly modeled experiment one. Let $\mathcal{E}_{1}$ be the event in experiment three which corresponds to event $E_{1}$ in experiment one.

Observe that we cannot model the second experiment in the same manner, for we are dealing with functions of the form $f \square s_{i}=\left(f \oplus s_{i}\right) \circ\left(f \oplus s_{i}\right)(x)$. An example of the problems which can occur is as follows: it is possible for the gates $g_{i}$ and $g_{j}$ to be queried on two different inputs $\alpha_{i} \neq \alpha_{j}$, but it may be the case that $\left(f \oplus s_{i}\right)\left(\alpha_{i}\right)=\left(f \oplus s_{i}\right)\left(\alpha_{j}\right)$, and in this case the outputs of the two gates must be identical. However, the chances of randomly assigning the same output to both gates is negligibly small. We will remedy this problem, and present a model of experiment two which is similar to our model of experiment one.

We consider the gates in the following
$I_{1}^{1}, O_{1}^{1}, \cdots, I_{m(n)}^{1}, O_{m(n)}^{\mathrm{I}}, \cdots, I_{1}^{n^{r}}, O_{1}^{n^{r}}, \cdots, I_{m(n)}^{n^{r}}, O_{m(n)}^{n^{r}}$. If the input $\alpha_{b}^{a}$ of gate $I_{b}^{a}$ is not equal to any $\alpha_{d}^{c}$ or $\widehat{\beta}_{d}^{c}$ for any $(c, d)<(a, b)$, then we assign its output to be $\widehat{\beta}_{b}^{a} \leftarrow \beta_{b}^{a}$. Similarly, if the input $\widehat{\beta}_{b}^{a}$ of gate $O_{b}^{a}$ is not equal to any $\alpha_{d}^{c}$ or $\widehat{\beta}_{d}^{c}$ for any $(c, d)<(a, b)$, then we assign its output to be $\widehat{\gamma}_{b}^{a} \leftarrow \gamma_{b}^{a}$.

Alternatively, if the input $\alpha_{b}^{a}$ of gate $I_{b}^{a}$ is equal to $\alpha_{d}^{c}$ or $\widehat{\beta}_{d}^{c}$, for some $(c, d) \leq(a, b)$, then we assign its output to be $\widehat{\beta}_{b}^{a} \leftarrow f\left(\alpha_{b}^{a}\right) \oplus s^{a}\left(\alpha_{b}^{a}\right)$. Similarly, if the input $\widehat{\beta}_{b}^{a}$ of gate $O_{b}^{a}$ is equal to $\alpha_{d}^{c}$ or $\widehat{\beta}_{d}^{c}$, for some $(c, d) \leq(a, b)$, then we assign its output to be $\widehat{\gamma}_{b}^{a} \leftarrow f\left(\widehat{\beta}_{b}^{a}\right) \oplus s^{a}\left(\widehat{\beta}_{b}^{a}\right)$. These $n^{r}$ evaluations of $C_{n}$ now perfectly model experiment two. Let $\mathcal{E}_{2}$ be the event in experiment three which corresponds to $E_{2}$ in experiment two.

It is now easy to see that our view of assigning outputs to the function gates, as in the model of experiment one, goes wrong only if one of the gates $O_{b}^{a}$ is queried on an input which was previously queried for a gate $O_{d}^{c}$ or $I_{d}^{c}$, for some $(c, d) \neq(a, b)$. We formalize this by defining a collision. Given our random choices in experiment three, we say that
a collision has occurred if there exists $(a, b) \neq(c, d)$ s.t. $\widehat{\beta}_{b}^{a}=\widehat{\beta}_{d}^{c}$ or $\widehat{\beta}_{b}^{a}=\alpha_{d}^{c}$. Let $\mathcal{E}_{3}$ be the event that a collision has occurred in our model of experiment two.

Given the new perspective of the experiments, observe that $\mathcal{E}_{2} \subset \mathcal{E}_{1} \cup \mathcal{E}_{3}$, and this implies that $\operatorname{Pr}\left[\mathcal{E}_{2}\right] \leq \operatorname{Pr}\left[\mathcal{E}_{1}\right]+\operatorname{Pr}\left[\mathcal{E}_{3}\right]$. We will now show that the $\operatorname{Pr}\left[\mathcal{E}_{3}\right] \leq \frac{1}{2^{n / 2}}$, and this combined with the fact that $\operatorname{Pr}\left[E_{1}\right] \leq \frac{1}{2^{n / 2}}$, and thus that $\operatorname{Pr}\left[\mathcal{E}_{1}\right] \leq \frac{1}{2^{n / 2}}$, implies that $\operatorname{Pr}\left[\mathcal{E}_{2}\right] \leq \frac{1}{2^{n / 3}}$ and thus that $\operatorname{Pr}\left[E_{2}\right] \leq \frac{1}{2^{n / 3}}$. This proves the lemma.

It remains to show that $\operatorname{Pr}\left[\mathcal{E}_{3}\right] \leq \frac{1}{2^{\pi / 2}}$. In order to do this we consider a forth experiment in which we fix $f \in \mathcal{F}^{n}$. We then draw a function $s$ uniformly at random from the set $S^{n}$. We then run the circuit $C_{n}$ on the oracle $f \square s$, by having the gates $I_{j}$ and $O_{j}$ compute the function $f \oplus s$, for each $j$.

Remember that $\alpha_{1}^{i}$ is completely determined by $C_{n}$, but for $j>1$ the choice of $\alpha_{j}^{i}$ can be dependent on $\widehat{\gamma}_{1}^{i}, \widehat{\gamma}_{2}^{i}, \ldots, \widehat{\gamma}_{j-1}^{i}$. Further, $\alpha_{j}^{i} \neq \alpha_{k}^{i}$, for $j \neq k$, by the construction of $C_{n}$. In experiment four, we will consider the probabilities involved in four cases. We show that in first three cases the probability is trivially bounded, and show that in the last case reduces to the problem taken care of in Lemma 6.2. Before describing the cases we define a set used in the four cases. The set is $\widehat{S}^{n}=\left\{f \oplus s_{i} \mid s_{i} \in S^{n}\right\}$.
Case 1: There exists an $i, j, k, l$, where $k=i$, such that $\widehat{\beta}_{j}^{i}=\widehat{\beta}_{l}^{k}$.
Fix $f \in \mathcal{F}^{n}$, and consider:

$$
\operatorname{Pr}_{s_{i} \in \mathcal{S}^{n}}\left[\widehat{\beta}_{j}^{i}=\widehat{\beta}_{l}^{k}\right]=\operatorname{Pr}_{s_{i} \in \mathcal{S}^{n}}\left[f\left(\alpha_{j}^{i}\right) \oplus s_{i}\left(\alpha_{j}^{i}\right)=f\left(\alpha_{l}^{k}\right) \oplus s_{k}\left(\alpha_{l}^{k}\right)\right]
$$

Notice that when $i=k$ that

$$
\operatorname{Pr}_{s_{i} \in \mathcal{S}^{n}}\left[f\left(\alpha_{j}^{i}\right) \oplus s_{i}\left(\alpha_{j}^{i}\right)=f\left(\alpha_{l}^{k}\right) \oplus s_{k}\left(\alpha_{l}^{k}\right)\right]=\operatorname{Pr}_{\hat{s}_{i} \in \hat{S}^{n}}\left[\hat{s}_{i}\left(\alpha_{j}^{i}\right)=\hat{s}_{i}\left(\alpha_{l}^{k}\right)\right]
$$

A simple argument shows that this probability is bounded by $\frac{m(n)^{2} \cdot n^{c}}{2^{n}}$. To notice this, observe that given a function $\hat{s}$, drawn uniformly at random from $\mathcal{F}^{\boldsymbol{n}}$, and given $\boldsymbol{m}(n)$ queries, that for any two distinct queries $a$ and $b$, the probability that $\hat{s}(a)=\hat{s}(b)$ is less than $\frac{m(n)^{2}}{2^{n}}$. Then notice when $\hat{s}$ is drawn uniformly at random from $\widehat{S}^{n}$ instead of $\mathcal{F}^{n}$, we can easily bound the probability by $\frac{m(n)^{2} \cdot n^{c}}{2^{n}}$.

Case 2: There exists an $i, j, k, l$, where $k \neq i$, such that $\widehat{\beta}_{j}^{i}=\widehat{\beta}_{l}^{k}$.
Fix $f \in \mathcal{F}^{n}$, and consider.

$$
\operatorname{Pr}_{s_{i} \in \mathcal{S}^{n}}\left[\widehat{\beta}_{j}^{i}=\widehat{\beta}_{l}^{k}\right]=\operatorname{Pr}_{s_{i} \in \mathcal{S}^{n}}\left[f\left(\alpha_{j}^{i}\right) \oplus s_{i}\left(\alpha_{j}^{i}\right)=f\left(\alpha_{l}^{k}\right) \oplus s_{k}\left(\alpha_{l}^{k}\right)\right]
$$

Notice that $f\left(\alpha_{l}^{k}\right) \oplus s_{k}\left(\alpha_{l}^{k}\right)$ and $\alpha_{l}^{k}$ are completely independent of the choice of $s_{i}$. Therefore, we can bound this probability by bounding the probability

$$
\operatorname{Pr}_{s_{i} \in \mathcal{S}^{n}}\left[f\left(\alpha_{j}^{i}\right) \oplus s_{i}\left(\alpha_{j}^{i}\right)=\vartheta\right]
$$

where $\vartheta$ is a fixed value which has been chosen, non-uniformly, to maximize the above probability. This gives the following probability,

$$
\operatorname{Pr}_{s_{i} \in \mathcal{S}^{n}}\left[f\left(\alpha_{j}^{i}\right) \oplus s_{i}\left(\alpha_{j}^{i}\right)=\vartheta\right]=\operatorname{Pr}_{\hat{s}_{i} \in \hat{S}^{n}}\left[\hat{s}_{i}\left(\alpha_{j}^{i}\right)=\vartheta\right]
$$

This probability is bound to be less than $\frac{m(n) \cdot n^{c}}{2^{n}}$. To notice this, observe that given a function $\hat{s}$, drawn uniformly at random from $\mathcal{F}^{n}$, and given $m(n)$ queries, that for any query $a$, the probability that $\hat{s}(a)=\vartheta$ is less than $\frac{m(n)}{2^{n}}$. Then notice when $\hat{s}$ is drawn uniformly at random from $\widehat{S}^{n}$ instead of $\mathcal{F}^{n}$, we can easily bound the probability by $\frac{m(n) \cdot n^{c}}{2^{n}}$.
Case 3: There exists an $i, j, k, l$, where $k \neq i$, such that $\widehat{\beta}_{j}^{i}=\alpha_{i}^{k}$.
Again we fix $f \in \mathcal{F}^{\boldsymbol{n}}$, and consider:

$$
\begin{aligned}
\operatorname{Pr}_{s_{i} \in \mathcal{S}^{n}}\left[\alpha_{l}^{k}=\widehat{\beta}_{j}^{i}\right] & =\operatorname{Pr}_{s_{i} \in \mathcal{S}^{n}}\left[\alpha_{l}^{k}=f\left(\alpha_{j}^{i}\right) \oplus s_{i}\left(\alpha_{l}^{i}\right)\right] \\
& =\operatorname{Pr}_{\hat{s}_{i} \in \mathcal{S}^{n}}\left[\alpha_{l}^{k}=\hat{s}_{i}\left(\alpha_{j}^{i}\right)\right]
\end{aligned}
$$

Similarly to the previous case, notice that $\alpha_{l}^{k}$ is completely independent of the choice of $\hat{s}_{i}$. Therefore, we can bound this probability by bounding the following probability,

$$
\operatorname{Pr}_{\hat{s}_{i} \in \hat{S}^{n}}\left[\vartheta^{\prime}=\hat{s}_{i}\left(\alpha_{j}^{i}\right)\right]
$$

where $\vartheta^{\prime}$ is a fixed value which has been chosen, non-uniformly, to maximize the above probability. This probability is bound to be less than $\frac{m(n) \cdot n^{c}}{2^{n}}$, for the same reasons as those mentioned in Case 2.

Case 4: There exists an $i, j, k, l$, where $k=i$, such that $\widehat{\beta}_{j}^{i}=\alpha_{l}^{k}$. Again we fix $f \in \mathcal{F}^{n}$, and consider:

$$
\begin{aligned}
\operatorname{Pr}_{s_{i} \in \mathcal{S}^{n}}\left[\alpha_{l}^{k}=\widehat{\beta}_{j}^{i}\right] & =\operatorname{Pr}_{s_{i} \in \mathcal{S}^{n}}\left[\alpha_{l}^{k}=f\left(\alpha_{j}^{i}\right) \oplus s_{i}\left(\alpha_{l}^{i}\right)\right] \\
& =\operatorname{Pr}_{\boldsymbol{s}_{i} \in \widehat{S}^{n}}\left[\alpha_{l}^{k}=\hat{s}_{i}\left(\alpha_{j}^{i}\right)\right]
\end{aligned}
$$

Clearly, the probability in last cases is bound by a bound on the following probability. First, let $\left\{B_{n}\right\}$ be any polynomial in $n$ sized circuit family, in which every circuit $B_{n}$ has $m(n)$ oracle gates. We consider the probability when $\hat{s}$ is chosen uniformly at random from the set $\widehat{S}^{n}$, and that when given as an input the oracle $\hat{s} \circ \hat{s}$ the circuit $B_{n}$ outputs the pair ( $\alpha, \mu$ ), where $\alpha$ corresponds to one of the oracle queries made by $B_{n}$, and $\mu=\hat{s}(\alpha)$. Fortunately, the above bound is given in Lemma 6.2.

Therefore, for each choice of $i, j, k$ and $l$ there are two possible ways in which a collision can occur, and for each way the probability of it occurring is bounded above by $\frac{n^{c} \cdot m(n)^{3}}{2^{n}}$ from the results of Lemma 6.2. There are a total of $\frac{n^{2 r}+m(n)^{2}}{2}$ choices of $i, j, k, l$, so:

$$
\operatorname{Pr}\left(\mathcal{E}_{3}\right) \leq \frac{n^{c} \cdot m(n)^{3}\left(\frac{n^{2 r}+m(n)^{2}}{2}\right)}{2^{n}}<\frac{1}{2^{\frac{n}{2}}}
$$

and the lemma is proved.

Corollary 6.2 For each $n$, let $S^{n} \subset \mathcal{F}^{n}$ by any set such that $\frac{\left|S^{n}\right|}{\left|\mathcal{F}^{n}\right|} \geq \frac{1}{n^{d}}$, for some constant d, and let $\left\{C_{n}\right\}$ be a polynomial sized family of decision-circuits. Then for every constant c, and for all but $\frac{1}{2^{n / 3}}$ of the $w \in \mathcal{F}^{n}$ :

$$
\left\lvert\, \underset{C_{n}}{\operatorname{Pr}\left(w \square S^{n}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) \left\lvert\,<\frac{1}{n^{c}}\right., ~}\right.
$$

for sufficiently large $n$.

This corollary corresponds directly to Lemma 5.7 , but with the composition operator and the set of permutations $S^{n}$ replaced by the $\square$ operator and the set of functions $S^{n}$.

### 6.2 Isolation Lemmas

Now by Corollaries 6.1 and 6.2 we can prove both versions of the Isolation Lemmas in Chapter 5 by replacing the composition operator with the operator, and by replacing the PRPGs with PRFGs ${ }^{1}$. For completeness, we state both versions of Isolation Lemmas and their corresponding composition theorems below.

Lemma 6.4 (Box Isolation Lemma - Weak Version) -
There exist fixed polynomials $p_{1}$ and $p_{2}$ such that for all $0 \leq \epsilon, \delta \leq 1$; polynomials $c_{G}, c_{\mathrm{H}}$, and $s_{F}$; and function generators H and G , where $c_{\mathrm{G}}(n)$ and $c_{\mathrm{H}}(n)$ bound from above the size of the circuits which compute $\mathrm{G}^{n}$ and $\mathrm{H}^{n}$ respectively. Define $\mathrm{F}=\mathrm{G} \square \mathrm{H}$.

Hypothesis: If there exists a family of decision-circuits $\left\{C_{n}\right\}$, where for each $n$ the circuit $C_{n}$ is of size $s_{F}(n)$, and for some $c>0$ and infinitely many $n$ :

$$
\left|\operatorname{Pr}_{C_{n}}\left(F^{n}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \epsilon \delta(2-\max \{\epsilon, \delta\})+\frac{1}{n^{c}}
$$

Conclusion: Ther for each sufficiently large $n$ there exists either a decision-circuit $\Lambda_{n}$ of size $p_{1}\left(n^{c} \cdot c_{G}(n)\right) s_{F}(n)$ for which:

$$
\left\lvert\, \underset{\Lambda_{n}}{\operatorname{Pr}\left(G^{n}\right)-\underset{\Lambda_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) \left\lvert\, \geq \delta+\frac{1}{n^{24 c}}\right. ;, ~ ; ~}\right.
$$

or a decision-circuit $\Upsilon_{n}$ of size $p_{2}\left(n^{c}-c_{G}(n)\right) s_{F}(n)$ for which:

$$
\left|\operatorname{Pr}_{\mathbf{r}}\left(H^{n}\right)-\underset{\mathrm{Y}_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \epsilon+\frac{1}{n^{3 c}} ;
$$

or a decision-circuit $\Xi_{n}$ of size $c_{H}(n)+s_{F}(n)$ for which:

$$
\left|\operatorname{Pr}_{\Xi_{n}}\left(G^{n}\right)-\underset{\Xi_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \delta+\frac{1}{n^{6 c}} .
$$

Corollary 6.3 (Box Theorem) Let G be a $1-\epsilon$ secure PRPG. Then for each positive integer $c$, the generator $F=\underbrace{(\cdots(G \square G) \square \cdots \square G)}_{c \text { limet }}$ is $1-\theta$ secure, where $\theta=\epsilon^{c}(2-\epsilon)^{c-1}$.

[^8]For the theorem below we need a definition for a semi-secure PRFG which is analogous to the definition of a semi-secure PRPG. We will call a PRFG, H, $1-\epsilon$ semi-secure if there exists no family of polynomial-sized decision- circuits $\left\{C_{n}\right\}$ which has an advantage of $\epsilon+\frac{1}{\log ^{c} n}$ in distinguishing $H^{n}$ from $\mathcal{F}^{n}$, for some constant $c$ and infinitely many $n$.

## Lemma 6.5 (Box Isolation Lemma - Stronger Version) -

There exist a fixed polynomials $p_{1}$ and a fixed poly-logarithmic function $p_{2}$ such that for all $0<\delta<1 ; \epsilon: \mathbb{N} \rightarrow[0,1]$, where for all sufficiently large $n, \epsilon(n)<\delta$; polynomials $c_{G}, c_{\mathrm{H}}$, and $s_{\mathrm{F}}$; and function generators H and G , where $c_{G}(n)$ and $c_{\mathrm{H}}(n)$ bound from above the size of the circuits which compute $\mathrm{G}^{n}$ and $\mathrm{H}^{n}$ respectively. Define $\mathrm{F}=\mathrm{GaH}$.
Hypothesis: If there exists a family of decision-circuits $\left\{C_{n}\right\}$, where for each $n$ the circuit $C_{n}$ is of size $s_{F}(n)$, and for some $c>0$ and all sufficiently large $n$ :

$$
\left|\operatorname{Pr}_{C_{n}}\left(F^{n}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \epsilon(n) \delta(2-\max \{\epsilon, \delta\})+\frac{1}{\log ^{c} n}
$$

Conclusion: Then for infinitely many $n$ there exists either a decision-circuit $\Lambda_{n}$ of size $p_{1}\left(\log ^{c} n \cdot c_{G}(n)\right) s_{F}(n)$ for which:

$$
\left\lvert\, \underset{\Lambda_{n}}{\operatorname{Pr}\left(G^{n}\right)-\underset{\Lambda_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) \left\lvert\, \geq \delta+\frac{1}{n}\right. ;, ~ ; ~}\right.
$$

or a decision-circuit $\Upsilon_{n}$ of size $p_{2}\left(\log ^{c} n \cdot c_{G}(n)\right) s_{F}(n)$ for which:

$$
\left|\operatorname{Pr}_{\mathrm{Y}_{n}}\left(\mathrm{H}^{n}\right)-\underset{\mathrm{Y}_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \epsilon(n)+\frac{1}{\log ^{2 c} n}
$$

or a decision-circuit $\Xi_{n}$ of size $c_{H}(n)+s_{F}(n)$ for which:

$$
\left|\operatorname{Pr}_{\Xi_{n}}\left(H^{n}\right)-\underset{\Xi_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \delta+\frac{1}{n} .
$$

Theorem 6.1 (Box Theorem - Stronger Version) -
Let $G$ be a $1-\epsilon$ secure $P R P G$. Then for each $f \in \mathcal{O}(\log \log n)$ there exits a generator $\mathrm{F}=\underbrace{(\cdots(\mathrm{G} \square \mathrm{G}) \square \cdots \square \mathrm{G})}_{f(n) \text { eimet }}$ which is $1-\theta(n)$ semi-secure, where $\theta(n)=\epsilon^{f(n)}(2-\epsilon)^{f(n)-1}$.

As in the case of the composition operator acting on permutation generators, it is a natural question to ask if the Isolation Lemma's security parameter is still tight for the function generator that results from the $\square$ operator acting on partially secure PRFGs. We show in the next section that the Isolation Lemma is still tight.

### 6.3 The Isolation Lemma is Still Tight

We now show that there exist PRFGs $H$ and $G$ which are respectively $(1-\epsilon)$ and $(1-\delta)$ secure, but when boxed together are exactly $(1-\theta)$ secure, where $\theta=\epsilon \delta(2-\max \{\epsilon, \delta\})$. We point out to the reader that there is much in common between this construction and the construction presented in Chapter 5.

## The Construction of G and H

We present the construction of $G$, and note that the construction of $H$ is similar. To simplify the presentation we assume that $\delta$ and $\epsilon$ are of the form $\frac{1}{2^{c}}$ or $1-\frac{1}{2^{c}}$. Let $\widehat{\mathrm{G}}=\left\{\widehat{\mathrm{G}}^{n}:\{0,1\}^{\ell(n)} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ be a PRFG. We construct $\mathrm{G}^{n}=\left\{\mathrm{G}^{n}:\right.$ $\left.\{0,1\}^{\ell(n)+c} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ to be $1-\delta$ secure in two steps. We will describe the construction of $\mathrm{G}^{n}$, and note that the construction of $\mathrm{H}^{n}$ is similar.

First, for each $n$ we set a fraction $\delta$ of the keys of $\mathrm{G}^{\boldsymbol{n}}$ to correspond to the zero function, $f(x)=\overline{0} \forall x$, and the remainder to correspond to functions chosen from $\widehat{G}^{n}$.This is done in two different fashions dependent on the form of $\delta$ as described below:

Case $1\left(\delta=1-\frac{1}{2^{c}}\right)$ For each $k \in\{0,1\}^{\ell(n)+c}$ we set the function $g_{k}^{n} \in G^{n}$ to be the zero function, if the first $c$ bits of $k$ are not all 0 ; otherwise let $\bar{k}$ be the last $\ell(n)$ bits of $k$, and set $g_{k}^{n}=\hat{g}_{k}^{n}$, for $\hat{g}_{\bar{k}}^{n} \in \widehat{\mathbf{G}}^{n}$.

Case $2\left(\delta=\frac{1}{2^{c}}\right)$ For each $k \in\{0,1\}^{\ell(n)+c}$ we set the function $g_{k}^{n} \in G^{n}$ to be the zero function, if the first $c$ bits of $k$ are all 0 ; otherwise let $\bar{k}$ be the last $\ell(n)$ bits of $k$, and set $g_{k}^{n}=\hat{g}_{\hat{k}}^{n}$, for $\hat{g}_{\vec{k}}^{n} \in \widehat{G}^{n}$.

The second step in the construction of $\mathrm{G}^{n}$ is to ensure that for each $k \in\{0,1\}^{\ell(n)+c}$ we set the value of $g_{k}^{n}\left(g_{k}^{n}(\overline{0})\right)$ in one of two manners, depending on the form of $\delta$, as described below:

Case $1\left(\delta=1-\frac{1}{2^{c}}\right)$ For all keys $k$ set the first $c$ bits of $g_{k}^{n}\left(g_{k}^{n}(\overline{0})\right)$ to 0 ; the last $n-c$ bits remain as they were in $g_{k}^{n}(\overline{0})$.

Case $2\left(\delta=\frac{1}{2^{c}}\right)$ For all keys $k$ ensure that not all of the first $c$ bits of $g_{k}^{n}\left(g_{k}^{n}(\overline{0})\right)$ are 1 . If they are, then set them to be a member of the set $\{0,1\}^{c} \backslash\left\{1^{c}\right\}$ chosen uniformly at random ${ }^{2}$; the last $n-c$ bits remain as they were in $g_{k}^{n}\left(g_{k}^{n}(\overline{0})\right)$.

It remains to show that $G$ and $H$ are in fact $1-\delta$ and $1-\epsilon$ secure. We note that the proof is similar to the proof presented in Appendix A, and so we will not present it here.

## The Adversary

We now show that there exists an adversary which can $\epsilon \delta(2-\max \{\epsilon, \delta\})$ distinguish $F=G \square H$, and thereby show that the isolation lemma result is tight. We assume WLOG that $\delta \geq \epsilon$.

Given a function $w \in \mathcal{F}^{n}$ the adversary, $A$, accepts in one of the following two conditions, dependent on the form of $\delta$ :

Case $1\left(\delta=1-\frac{1}{2^{c}}\right)$ The adversary $A$ accepts iff the first $c$ bits of $w(w(\overline{0}))$ are 0 ;
Case $2\left(\delta=\frac{1}{2^{c}}\right)$ The adversary $A$ accepts iff the first $c$ bits of $w(w(\overline{0}))$ are not 1 .
We first consider the accepting probability of $A$ if $w$ was chosen uniformly from $F^{n}$. In this case $w$ is the result of applying the box operator to two functions chosen randomly from each of the respective generators $G^{n}$ and $H^{n}$. We partition the functions of $G^{n}$ into

[^9]two sets. We let $I_{\mathbf{G}^{n}}$ represent the set of all zero functions in $\mathbf{G}^{\mathbf{n}}$, and we let $P_{\mathbf{G}^{n}}$ represent the remaining functions in $\mathrm{G}^{n}$. Similarly, we partition $\mathrm{H}^{\boldsymbol{n}}$ into two sets $I_{\mathrm{H}^{n}}$ and $P_{\mathrm{H}^{n}}$. We now consider the probability that $w$ is accepted based on it being the result of applying the box operator to two functions from the above mentioned partitions.

Case $1\left(w \in I_{G^{n}} \square I_{\mathrm{H}^{n}}\right)$ The adversary will accept $w$ with probability 1. This case occurs with probability $\delta \epsilon$.

Case $2\left(w \in I_{G^{n}} \square P_{\mathrm{H}^{n}}\right)$ The adversary will accept $w$ with probability $\frac{1-\delta}{1-\epsilon}$, which is the probability that $A$ would accept a random function conditioned on the fact that the function $w$ was chosen from $P_{H^{n}}$. This case occurs with probability $\delta(1-\epsilon)$.

Case 3 ( $w \in P_{\mathrm{G}^{n}} \square I_{\mathrm{H}^{n}}$ ) The adversary will accept $w$ with probability 1. This case occurs with probability $(1-\delta) \epsilon$.

Case $4\left(w \in P_{\mathrm{G}^{n}} \square P_{\mathrm{H}^{n}}\right)$ The adversary will accept $w$ with probability $1-\delta$, which is the probability that the result of a random function evaluated at a random point in its range would meet the acceptance criteria of $A$. This case occurs with probability $(1-\delta)(1-\epsilon)$.

Clearly a random function $w$ will be accepted by $A$ with probability $1-\delta$. Now following the argument made in section 5.1.1 it is clear that the adversary has distinguishing probability of $\epsilon \delta(2-\max \{\epsilon, \delta\})$.

### 6.4 Scalability Issues

Again, our goal is to show that by scaling this construction large enough we can achieve a 1 -secure PRPG. We believe that this will be possible by considering the finer notion of security which is proposed in section 5.4. However, unless we can eventually achieve complete security by a construction $\underbrace{(\cdots(G \square G) \square \cdots \square G)}_{O(\log n)}$ then this construction will not
be useful. This is because the size of the construction becomes larger than any fixed polynomial for sufficiently large $n$, and thus the generator is not computable in polynomial time. This is due to the fact that the size of the implementation of the construction is exponential in the number of $\square$ operators which are used in the construction. This becomes obvious by rewriting the construction and substituting the operator with its definition. This may limit the usefulness of this construction in practice, for it may be too slow to practically consider. Further, it may prevent polynomial size constructions of this form from becoming PRFGs.

## Chapter 7

## The $\diamond$ Operator Generator

In the previous two chapters we have shown that the composition and $\square$ operators are security increasing for PRPGs and PRFGs respectively. Unfortunately, in both cases we have not been able to show that constructions based on them can give 1 -secure generators. In this chapter we will show that the $\diamond$ operator generator is security increasing when applied to PRFGs. Further, we will show that it can be used to construct completely secure PRFGs from $1-\delta$ secure PRFGs. We now remind the reader of the definition of the $\diamond$ operator generator.

We define the $\diamond$ operator generator (read as Diamond) as $\diamond=\left\{\diamond_{r_{1}, r_{2}}^{n} \mid n \in \mathbb{N} \wedge r_{1}, r_{2} \in\right.$ $\left.\{0,1\}^{n}\right\}$. Let $f_{1}$ and $f_{2}$ be two functions such that $f_{i}:\{0,1\}^{n} \rightarrow\{0,1\}^{n}$, for $i \in\{1,2\}$. For each $r_{1}, r_{2} \in\{0,1\}^{n}$ we define the operator $\diamond_{r_{1} \bullet r_{2}}^{n}$, which acts on the functions $f_{1}$ and $f_{2}$ as:

$$
\left(f_{1} \diamond_{r_{1} \odot r_{2}}^{n} f_{2}\right)(x)=f_{1}\left(x \oplus r_{1}\right) \oplus f_{2}\left(x \oplus r_{2}\right)
$$

We will now show that this operator-generator is security increasing, by modifying the arguments which were presented in Chapters 5 and 6 . We will see that the fact that $\diamond$ is an operator-generator, as opposed to an operator, does not have anything more than a cosmetic effect on the proof.

## Lemma 7.1 (Diamond Isolation Lemma) -

There exists a fixed polynomial $p_{2}$ such that for all $\epsilon, \delta: \mathbb{Z} \rightarrow[0,1]$; polynomials $c_{G}, c_{H}$, and $s_{\mathrm{F}}$; and function generators H and G , where $c_{\mathrm{G}}(n)$ and $c_{\mathrm{H}}(n)$ bound from above the size of the circuits which compute $\mathrm{G}^{\boldsymbol{n}}$ and $\mathrm{H}^{\boldsymbol{n}}$ respectively. Define $\mathrm{F}=\mathrm{G} \diamond \mathrm{H}$.

Hypothesis: If there exists a family of decision-circuits $\left\{C_{n}\right\}$, where for each $n$ the circuit $C_{n}$ is of size $s_{F}(n)$, and for some $c>0$ and infinitely many $n$ :

$$
\left\lvert\, \underset{C_{n}}{\operatorname{Pr}\left(F^{n}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) \left\lvert\, \geq \epsilon(n) \delta(n)+\frac{1}{n^{c}} .\right.}\right.
$$

Conclusion: Then for infinitely many $n$ there exists either a decision-circuit $\Upsilon_{n}$ of size $p_{2}\left(n^{c} \cdot c_{G}(n)\right) s_{F}(n)$ for which:

$$
\left|\operatorname{Pr}_{\mathrm{P}_{n}}\left(\mathrm{H}^{n}\right)-\underset{\mathrm{T}_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \epsilon(n)+\frac{1}{n^{3 c}} ;
$$

or a decision-circuit $\Xi_{n}$ of size $c_{H}(n)+s_{F}(n)$ for which:

$$
\left|\operatorname{Pr}_{\Xi_{n}}\left(G^{n}\right)-\underset{\Xi_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \delta(n)+\frac{1}{n^{6 c}} .
$$

Notice that in this version of the Isolation Lemma that there is only one circuit which distinguishes $G$ from random, and the resulting circuit is much smaller than the circuit needed to break $H$. This allows us to iteratively apply the Isolation Lemma to the construction $\underbrace{G \diamond \ldots \diamond G}_{p(n)}$ for a $p \in \Omega\left(\log ^{2} n\right) \cap\left(\cup_{i=1}^{\infty} \mathcal{O}\left(n^{i}\right)\right)$ and achieve complete security. Theorem 7.1 (Diamond Composition Theorem) -

Let $G$ be a $1-\epsilon$ secure PRFG. Then for each $p \in \Omega\left(\log ^{2} n\right) \cap\left(\cup_{i=1}^{\infty} \mathcal{O}\left(n^{i}\right)\right)$ the generator $\mathrm{F}=\underbrace{\mathrm{G} \diamond \ldots \diamond G}_{p(n)}$ is a secure PRFG.

Proof: Similar to the proof of corollary 5.1.

We will now present a proof of Lemma 7.1. As was the case in Chapter 6, the important changes to the proof rely on two technical lemmas. We present these lemmas first, and then give the complete proof. It will become evident that the fact that we are using an operator generator instead of an operator has no significant effect on the proof.

### 7.1 Two Technical Lemmas

As mentioned in Chapter 6, the proof of security presented in Chapter 5 only makes specific use of the properties of composition and permutations in Lemmas 5.6 and 5.7. Therefore, we will need to prove corresponding lemmas which correspond to the $\diamond$ operator. In this chapter we shall assume the same standardized circuits and circuit families which were presented in Section 6.1. Specifically, all circuits will never repeat oracle queries, and all circuits $C_{n}$ in a circuit family $\left\{C_{n}\right\}$ will perform exactly $m(n)$ queries, for some polynomial $m$.

Below we present the first technical lemma. It corresponds to Lemma 5.6.

Lemma 7.2 Given any decision-circuit $C$, for each $f \in \mathcal{F}^{n}$ and for each $r_{1}, r_{2} \in\{0,1\}^{n}$ :

$$
\operatorname{Pr}_{C}\left(f \diamond_{r_{1} \bullet r_{2}}^{n} \mathcal{F}^{n}\right)=\operatorname{Pr}_{C}\left(\mathcal{F}^{n}\right)
$$

Proof: First observe that for each $r_{2} \in\{0,1\}^{n}$, that the set $\left\{f^{\prime}\left(x \oplus r_{2}\right) \mid f^{\prime} \in \mathcal{F}^{n}\right\}=\mathcal{F}^{n}$. Then let $g(x)=f\left(x \oplus r_{1}\right)$, and observe that the $g \oplus \mathcal{F}^{n}=\mathcal{F}^{n}$, proving the result.

Corollary 7.1 Given any decision-circuit $C$, for each $f \in \mathcal{F}^{n}$ :

$$
\operatorname{Pr}_{C}\left(f \diamond \mathcal{F}^{n}\right)=\operatorname{Pr}_{C}\left(\mathcal{F}^{n}\right)
$$

We now present the second technical lemma, and from it we will derive a simple corollary which corresponds to Lemma 5.7. Notice that the proof of this lemma is very similar to the proof of Lemma 6.3.

Lemma 7.3 For every $n$ let $S^{n} \subset \mathcal{F}^{n}$ be a non-empty set. Let $\left\{C_{n}\right\}$ be a family of polynomial in $n$ sized decision-circuits. We show that for any $d>0$, there exists an $r>0$ such that for all sufficiently large $n$ :
$\underset{\left(f, k_{1}^{1}, k_{1}^{2}, \ldots, k_{n}^{1} r, k_{n}^{2} r, s_{1}, \ldots, s_{n} r\right) \in \mathcal{F} n \times\{0,1\}^{(2 n)\left(n^{r}\right)} \times\left(S^{n}\right)^{n r}}{\operatorname{Pr}}\left[\left|\frac{1}{n^{r}} \sum_{i=1}^{n^{1}} C_{n}\left(f \diamond_{k_{i}^{1} \otimes k_{i}^{2}}^{n} s_{i}\right)-\underset{C_{n}}{\operatorname{Pr}\left(\mathcal{F}^{n}\right)}\right|>\frac{1}{n^{d}}\right]<\frac{1}{2^{n / 3}}$.

Notice that in this lemma there is no restriction on the size of $S^{n}$, other than it being non-empty. It is this lack of restriction on the size of $S^{n}$ which ends up allowing us to iteratively apply the $\diamond$ operator a polynomial number of times, as is stated in the Diamond Composition Theorem (Theorem 7.1). Notice than in similar Lemmas presented in Chapters 5 and 6 (Lemmas 5.8 and 6.3), there is a requirement that $\frac{\left|S^{n}\right|}{\left|\mathcal{F}^{n}\right|} \geq \frac{1}{n^{c}}$, for some constant $c$. Therefore, this lemma is a much stronger result.

Proof: We will assume that $r$ is fixed, and later we will show how to determine $r$ 's value. First we define two experiments. In the first experiment pick random $\left(f_{1}, f_{2}, \ldots, f_{n}^{r}\right) \in$ $\left(\mathcal{F}^{n}\right)^{n^{r}}$, and evaluate $C_{n}\left(f_{i}\right)$ for each $i \in\left\{1, \ldots, n^{r}\right\}$. Define the event $E_{1}$ to be:

$$
\left|\frac{1}{n^{r}} \sum_{i=1}^{n^{r}} C_{n}\left(f_{i}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right|>\frac{1}{n^{d}} .
$$

In the second experiment pick random $\left(f, k_{1}^{1}, k_{1}^{2}, \ldots, k_{n^{r}}^{1}, k_{n^{r}}^{2}, s_{1}, \ldots, s_{n^{r}}\right) \in \mathcal{F}^{n} \times\{0,1\}^{(2 n)\left(n^{r}\right)} \times$ $\left(S^{n}\right)^{n^{r}}$, and evaluate $C_{n}\left(f \diamond_{k_{i}^{1} \bullet k_{i}^{2}} s_{i}\right)$ for each $i$, where $1 \leq i \leq n^{r}$. Define the event $E_{2}$ to be:

$$
\left|\frac{1}{n^{r}} \sum_{i=1}^{n^{r}} C_{n}\left(f \diamond_{k_{i}^{\prime} \bullet k_{i}^{n}}^{n} s_{i}\right)-\underset{C_{n}}{\operatorname{Pr}\left(\mathcal{F}^{n}\right)}\right|>\frac{1}{n^{d}}
$$

Using the Chernoff Bound (Lemma 5.4) we choose an $r$ where the probability of event $E_{1}$ occurring in the first experiment is less than $\frac{1}{2^{\frac{1}{2}}}$. We will show that the probability of event $E_{2}$ in the second experiment, is negligibly close to the probability of $E_{1}$ in the first experiment. The lemma will follow.

We will perform a third experiment, in which we model both of the first two experiments by considering two different methods of evaluating the circuit $C_{n}, n^{r}$ times.

We choose $s_{1}, \ldots, s_{r} \in S^{n} ; \gamma_{1}^{1}, \ldots, \gamma_{m(n)}^{1}, \ldots ., \gamma_{1}^{n^{\gamma}}, \ldots, \gamma_{m(n)}^{n^{r}} \in\{0,1\}^{n \cdot m(n) \cdot n^{r}} ;$ and $k_{1}^{1}, k_{1}^{2}, \ldots, k_{n^{r}}^{1}, k_{n^{r}}^{2} \in\{0,1\}^{2 n \cdot n^{r}}$. Let $g_{j}^{i}$ represent the $j^{\text {th }}$ oracle-gate of $C_{n}$ when evaluating $f_{i}$ in experiment one, or $f \diamond_{k_{i}^{i} \bullet k_{i}^{2}}^{n} s_{i}$ in experiment two. Let $I_{j}^{i}$ be the input to $g_{j}^{i}$ in the experiment and let $O_{j}^{i}$ be its output. We consider the gates in the following order $g_{1}^{1}, g_{2}^{1}, \ldots, g_{m(n)}^{1}, \ldots, g_{1}^{n^{r}}, \ldots g_{m(n)}^{n^{r}}$.

We model the first experiment of performing $n^{r}$ evaluations of $C_{n}$, where in each
evaluation we set $C_{n}$ 's oracle to be a random function by equivalently considering $\boldsymbol{n}^{\boldsymbol{r}}$ evaluations of $C_{n}$, where in the evaluation of $f_{j}$ we independently assign the element $\boldsymbol{\gamma}_{i}^{j}$ to $O_{j}^{i}$, the output of each oracle gate $g_{j}$ in $C_{n}$. Notice that by fixing, for each $\boldsymbol{i}$ and each $j$, the value of $O_{j}^{i}$, we have completely determined the behaviour of the circuit $C_{n}$ in each of the evaluations. Further, notice that because of the way we have set the outputs of the gates, we have perfectly modeled experiment one. Let $\mathcal{E}_{1}$ be the event in experiment three which corresponds to event $E_{1}$ in experiment one.

Observe that we could model the second experiment in the same manner if we could guarantee that $f$ would never be queried on the same input. Since our circuits never repeat queries, there is no worry that $C_{n}$ will every query $f \diamond_{k_{i}^{d} k_{i}^{2}}^{n} s_{i}$ on the same input. Unfortunately, in our experiment there is the possibility that for $i \neq j$ that $C_{n}$ will perform a query $\alpha$ on input $f \diamond_{k_{i}^{1} \diamond k_{i}^{2}}^{n} s_{i}$ and a query $\beta$ on input $f \diamond_{k_{j}^{1} \diamond k_{j}^{2}}^{n} s_{j}$ where $\alpha \oplus k_{i}^{1}=$ $\beta \oplus k_{j}^{1}$, and the result will be that $f$ is queried on the same input twice. However, we can model experiment two in a method similar to which we modeled experiment one.

For each evaluation of $C_{n}\left(f \diamond_{k_{i}^{\prime} \bullet k_{i}^{2}}^{n} s_{i}\right)$ we set the output $O_{\ell}^{i}$ to be $\gamma_{l}^{i}$, for each $\ell$ such that $1 \leq \ell \leq m(n)$. Observe that once the outputs of the oracle gates have been fixed, then all of the inputs to the oracle-gates are fixed, as is the output of the circuit. Now that the inputs and outputs have been fixed, the only way the $f$ can be queried on the same input twice is if we have bad choices for our $k_{i}^{1}$ 's. Now for each $i$, we check if there exists a pair $(a, b)$ and a $j$, where $(a, b)<(i, j)$, such that $I_{j}^{i} \oplus k_{i}^{1}=I_{b}^{a} \oplus k_{a}^{1}$; if such an ( $a, b$ ) and $j$ exist then we say a collision has occurred. A collision corresponds to a bad choice of $k_{i}^{1}$, and therefore $f$ is inadvertently being queried twice on the same input, and we therefore have to make sure that the two query responses are consistent. If a collision has occurred, then we reevaluate $C_{n}\left(f \diamond_{k_{i}^{j}-k_{i}^{2}}^{n} s_{i}\right)$ in the following manner. We reevaluate the gates $g_{1}^{i}, \ldots, g_{m(n)}^{i}$ in that order. For each gate $g_{j}^{i}$ we consider its input $I_{j}^{i}$, and the input to $f$ which is $I_{j}^{i} \oplus k_{i}^{1}$. If there exists a pair $(a, b)<(i, j)$ such that $I_{b}^{a} \oplus k_{a}^{1}=I_{j}^{i} \oplus k_{i}^{1}$, then $f$ is being queried on an input which has previously been performed
in the experiment. Therefore, we set the output $O_{j}^{i}$ to be $f\left(I_{j}^{i} \oplus k_{i}^{1}\right) \oplus s_{i}\left(I_{j}^{i} \oplus k_{i}^{2}\right)$, and force the oracle-gates to respond consistently to queries on $f$. If no such pair ( $a, b$ ) existed then the random response was a consistent response, and therefore we set $O_{j}^{i} \leftarrow \gamma_{j}^{i}$. Notice that this method of evaluating $C_{n}$, perfectly models experiment two. Let $\mathcal{E}_{2}$ be the event in experiment three which corresponds to $E_{2}$ in experiment two.

Note that the model of the second experiment is identical to the model of the first, except in those cases in which a collision occurs. We define $\mathcal{E}_{3}$ to be the event that a collision occurred during the third experiment. Clearly $\mathcal{E}_{2} \subset \mathcal{E}_{1} \cup \mathcal{E}_{3}$, which implies that $\operatorname{Pr}\left(\mathcal{E}_{2}\right) \leq \operatorname{Pr}\left(\mathcal{E}_{1}\right)+\operatorname{Pr}\left(\mathcal{E}_{2}\right)$. Since the probability of $E_{1}$ is less than $\frac{1}{2^{n / 2}}$, in experiment one, and thus the probability of $\mathcal{E}_{1}$ is less than $\frac{1}{2^{n / 2}}$, in experiment three; it suffices to show that the probability of event $\mathcal{E}_{3}$ in experiment three is less than $\frac{1}{2^{n / 2}}$ in order to prove that the probability of $\mathcal{E}_{2}$, in experiment three, is less than $\frac{1}{2^{n / 3}}$, which implies that the probability of $E_{2}$, in experiment two, is less than $\frac{1}{2^{n / 3}}$. This proves the lemma.

We now bound from above the probability of event $\mathcal{E}_{3}$. Let

$$
L_{i}=\left\{\left(\left(I_{j}^{i} \oplus k_{i}^{1}\right),\left(\gamma_{j}^{i} \oplus s_{i}\left(I_{j}^{i} \oplus k_{i}^{2}\right)\right)\right) \mid 1 \leq j \leq m(n)\right\}
$$

The set $L_{i}$ represents the pairs $(x, f(x))$ which would be needed in order to calculate the function $f \diamond_{k_{i}^{1} \bullet k_{i}^{2}}^{n} s_{i}$ for the queries which are made by $C_{n}$ during the evaluation of $C_{n}\left(f \diamond_{k_{i}^{j} \diamond k_{i}^{2}}^{n} s_{i}\right)$. Notice that $\left(I_{j}^{i} \oplus k_{i}^{1}\right)$ corresponds to the value at which $f$ is evaluated, given that the query $I_{j}^{i}$ is made to the function $f \diamond_{k_{i}^{j} \bullet k_{i}^{2}}^{n} s_{i}$. Similarly, $\left(\gamma_{j}^{i} \oplus s_{i}\left(I_{j}^{i} \oplus k_{i}^{2}\right)\right)$ corresponds to the value of $f$ evaluated on $\left(I_{j}^{i} \oplus k_{i}^{1}\right)$ given that $\left(f \diamond_{k_{i}^{1} \cdot k_{i}^{2}}^{n} s_{i}\right)\left(I_{j}^{i} \oplus k_{i}^{1}\right)=\gamma_{j}^{i}$. Since all of $C_{n}$ 's queries are unique it is the case that a repeated query to $f$ cannot be strictly contained in the evaluation of $C_{n}\left(f \diamond^{n} s_{i}\right)$, and therefore $\left|L_{i}\right|=m(n)$ for each $i$. We now consider the probability that during the evaluation of $C_{n}\left(f \diamond_{k_{i}^{1} \bullet k_{i}^{2}}^{n} s_{i}\right)$ that a collision occurs. We define the set

$$
T_{i}=\left\{I_{j}^{i} \oplus x \mid 1 \leq j \leq m(n) \wedge(x, f(x)) \in \cup_{\ell=1}^{i-1} L_{\ell}\right\}
$$

which contains all of the possible values of $k_{i}^{1}$ which will cause a collision. We observe
that $\left|T_{i}\right| \leq m(n)^{\mathbf{2}} \cdot i$. Therefore the probability of having a collision during the evaluation of $C_{n}\left(f \diamond_{k_{i}^{d} * k_{i}^{2}}^{n} s_{i}\right)$ is at most $\frac{m(n)^{2}-i}{2^{n}}$. We note that the probability of a collision occurring in experiment three is less than the sum of the probabilities of a collision occurring during the evaluations of $C_{n}\left(f \diamond_{k_{i}^{1} \bullet k_{i}^{2}}^{2_{i}}\right)$ for each $i$, where $1 \leq i \leq n^{r}$. Therefore:

$$
\begin{aligned}
\operatorname{Pr}\left[E_{3}\right] & \leq \frac{m(n)^{2}}{2^{n}}\left(1+2+\cdots+n^{r}\right) \\
& =\left(\frac{m(n)^{2}}{2^{n}}\right)\left(\frac{n^{r}\left(n^{r}+1\right)}{2}\right) \\
& =\frac{m(n)^{2} n^{2 r}+n^{r} m(n)^{2}}{2^{n+1}} \\
& \leq \frac{1}{2^{n}}
\end{aligned}
$$

This proves the lemma.

Corollary 7.2 For each $n$, let $S^{n} \subset \mathcal{F}^{n}$ be a non empty set, and let $\left\{C_{n}\right\}$ be a polynomial sized family of decision-circuits. Then for every constant $c$, and for all but $\frac{1}{2^{n / 3}}$ of the $w \in \mathcal{F}^{n}$.

$$
\left|\underset{C_{n}}{\operatorname{Pr}}\left(w \diamond S^{n}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right|<\frac{1}{n^{c}},
$$

for sufficiently large $n$.

This corollary corresponds to Lemma 5.7 , but is actually much stronger. Since there are no restrictions on the size of $S_{n}$ we are able to prove a lemma corresponding to Lemma 5.9 , as a direct consequence of Corollary 7.2

### 7.2 Proof Of Lemma 7.1

Assume that there exists a polynomial-sized decision-circuit family $\left\{C_{n}\right\}$ which for some constant $c>0$ and infinitely many $n$ distinguishes $F^{n}$ from $\mathcal{F}^{n}$ with probability at least $\epsilon(n) \delta(n)+\frac{1}{n^{c}}$.

Lemma 7.4 For $i>0$ and for each $n$ let

$$
K_{n}(i)=\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)+\frac{1}{n^{i}} \quad \text { and } \quad L_{n}(i)=\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)-\frac{1}{n^{i}}
$$

Let

Then given $i, j$ :

$$
\operatorname{Pr}_{w \in \mathcal{P}^{n}}\left(w \in S^{n}(i)\right) \leq \frac{1}{n^{j}} \quad \text { and } \quad \operatorname{Pr}_{w \in \mathcal{F} n}\left(w \in T^{n}(i)\right) \leq \frac{1}{n^{j}}
$$

for sufficiently large $n$.

Proof: This lemma follows directly from Corollary 7.2 as $\mathrm{G}^{\boldsymbol{n}} \subset \mathcal{F}^{\boldsymbol{n}}$.

Notice the similarities between this lemma, and Lemma 5.9. They are similar, but in Lemma 5.9 there exists the possibility of a family of circuits $\left\{\Lambda_{n}\right\}$, which does not exist in this lemma. It is the lack of this circuit family which permits the large number of applications of the Diamond Operator Generator in the Diamond Composition Theorem, and thereby allows us to attain a PRFG from a partially secure PRFG. Notice that the lack of the family of circuits $\{\Lambda\}$ in this lemma, as opposed to Lemma 5.9 , is made possible due to the lack of restrictions on the size of $S^{n}$ in Corollary 7.2, which is a direct consequence of the lack of restrictions on the size of $S^{n}$ in Lemma 7.3.

Lemma 7.5 Either there exists a family of decision-circuits $\left\{\Xi_{n}\right\}$, where for each $n$ the circuit $\Xi_{n}$ is of size $c_{H}(n)+s_{F}(n)$; and for infinitely many $n$ :

$$
\left|\operatorname{Pr}_{\underline{n}}\left(G^{n}\right)-\underset{\Xi_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \delta(n)+\frac{1}{n^{6 c}}
$$

or for all sufficiently large $n$ and all $h^{n} \in \mathrm{H}^{n}$ :

$$
\left|\operatorname{Pr}\left(G \diamond h^{n}\right)-\operatorname{Pr} C_{n}\left(\mathcal{F}^{n}\right)\right|<\delta(n)+\frac{1}{n^{6 c}} .
$$

Proof: If a family of circuits $\left\{\Xi_{n}\right\}$ with the above properties exists, then we are done. Otherwise, there exists no such family of circuits. Assume that for the circuit family $\left\{C_{n}\right\}$, there exists a set $\left\{h^{n} \in \mathbf{H}^{n} \mid n \in \mathbb{N}\right\}$ such that, for all sufficiently large $n$ it is the case that $\left|\operatorname{Pr}_{C_{n}}\left(G^{n} \diamond h^{n}\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right| \geq \delta(n)+\frac{1}{n^{\delta c}}$. We create the decision-circuit family $\left\{D_{n}\right\}$, where $D_{n}(w)=C_{n}\left(w \diamond h^{n}\right)$.

$$
\begin{align*}
\operatorname{Pr}_{D_{n}}\left(\mathrm{G}^{n}\right)-\underset{D_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) & =\underset{C_{n}}{\operatorname{Pr}}\left(G^{n} \diamond h^{n}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n} \diamond h^{n}\right) \\
& =\underset{C_{n}}{\operatorname{Pr}\left(G^{n} \diamond h^{n}\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)}  \tag{Lemma7.1}\\
& \geq \delta(n)+\frac{1}{n^{6 c}}
\end{align*}
$$

Letting $D_{n}=\Xi_{n}$ we have a constructed a family of decision-circuits $\left\{\Xi_{n}\right\}$. This contradicts the fact that there is no decision-circuit family of size $c_{H}(n)+s_{F}(n)$, such that for infinitely many $n$ :

$$
\left|\operatorname{Pr}_{\Xi_{n}}\left(\mathrm{G}^{n}\right)-\underset{\Xi_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \delta(n)+\frac{1}{n^{6 c}},
$$

proving the lemma.

### 7.2.1 Main Argument

We note that the Main Argument for the Diamond Isolation Lemma is similar, but easier than the Main Argument for the Composition Isolation Lemma presented earlier in section 5.3.1. One major difference is that it performs more sampling, and therefore achieves bounds under the normal security definition of a partially secure PRFG, and not under the semi-secure definition of PRFGS. Another difference is that there is no need to assume the existence of a family of circuits $\left\{\Lambda_{n}\right\}$, due to Lemma 7.4. The final difference is that we present the construction of only one type of circuit in this Main Argument. Unlike the Main Argument in Chapter 5, there is no need to present a circuit which flips biased coins. This is because for a random $\boldsymbol{w} \in \mathcal{F}^{n}$, Lemma 7.4 constrains the value of $\operatorname{Pr}_{C_{n}}\left[G^{n} \diamond w\right]$ to be much closer to $\operatorname{Pr}_{C_{n}}\left[\mathcal{F}^{n}\right]$, then is the case for the constraint imposed
on the distance between $\operatorname{Pr}_{C_{n}}\left[G^{n} \circ w\right]$ and $\operatorname{Pr}_{C_{n}}\left[\mathcal{F}^{n}\right]$, by Lemma 5.9.
Let $\bar{\epsilon}_{n}=\epsilon(n)+\frac{1}{n^{c}}$ and similarly let $\bar{\delta}_{n}=\delta(n)+\frac{1}{n^{c}}$. Then the family of circuits, $\left\{C_{n}\right\}$, is capable of distinguishing $G^{n} \diamond H^{n}$ from $\mathcal{F}^{n}$ with probability at least $\bar{\epsilon}_{n} \bar{\delta}_{n}$, for infinitely many $n$, as this value is strictly smaller than $\epsilon(n) \delta(n)+\frac{1}{n^{c}}$, for sufficiently large $n$. We assume that there exists no family of circuits $\left\{\Xi_{n}\right\}$, where each circuit $\Xi_{n}$ is of size $c_{\mathrm{H}}(n)+s_{\mathrm{F}}(n)$, such that for infinitely many $n$ :

$$
\left\lvert\, \underset{\underline{E r}_{n}}{\operatorname{Pr}\left(H^{n}\right)-\underset{\Xi_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) \left\lvert\, \geq \delta(n)+\frac{1}{n^{6 c}} . . . .\right.}\right.
$$

From the above assumption and Lemma 7.5, we know that for all sufficiently large $n$ and all $h^{n} \in \mathbf{H}^{n}$ :

$$
\begin{equation*}
\left|\operatorname{Pr}_{C_{n}}\left(G \diamond h^{n}\right)-\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right|<\delta(n)+\frac{1}{n^{6 c}} . \tag{7.1}
\end{equation*}
$$

From Lemma 7.4 we know that for

$$
\begin{equation*}
K_{n}=\operatorname{Pr}\left(\mathcal{F}_{n}^{n}\right)+\frac{1}{n^{5 c}} \quad \text { and } \quad S^{n}=\left\{w \in \mathcal{F}^{n}| |_{C_{n}} \operatorname{Pr}^{n}\left(G^{n} \diamond w\right) \geq K_{n}\right\} \tag{7.2}
\end{equation*}
$$

that $\operatorname{Pr}_{w \in \mathcal{F n}}\left(w \in S^{n}\right) \leq \frac{1}{n^{7 c}}$, for all sufficiently large $n$. Similarly we know that for

$$
\begin{equation*}
L_{n}=\underset{C_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)-\frac{1}{n^{5 c}} \quad \text { and } \quad T^{n}=\left\{w \in \mathcal{F}^{n}| |_{C_{n}}\left(G^{n} \diamond w\right) \leq L_{n}\right\} \tag{7.3}
\end{equation*}
$$

that $\operatorname{Pr}_{w \in \mathcal{F} n}\left(w \in T^{n}\right) \leq \frac{1}{n^{7 c}}$, for all sufficiently large $n$.
We will construct a family of decision-circuits $\left\{\Upsilon_{n}\right\}$, where each circuit $\Upsilon_{n}$ is of size $p_{2}\left(n^{c} \cdot c_{G}(n)\right) s_{F}(n)$, such that for infinitely many $n$ :

$$
\left|\operatorname{Pr}_{\mathrm{P}_{n}}\left(\mathrm{H}^{n}\right)-\underset{\mathrm{Y}_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| \geq \epsilon(n)+\frac{1}{n^{3 c}}
$$

proving the lemma. In the sequel we show how to construct $\Upsilon_{\boldsymbol{n}}$ for each $\boldsymbol{n}$ such that both $\left|\operatorname{Pr}_{C_{n}}\left(F^{n}\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right| \geq \epsilon(n) \delta(n)+\frac{1}{n^{c}}$., and $n$ is sufficiently large for all inequalities in the following proof to hold true.

We first give an overview of the proof. We will construct a circuit $B_{n}$ (via an intermediate circuit $A_{n}$ ) which, given as input a function $w$, almost surely approximates the
value $\operatorname{Pr}_{C_{n}}\left(G^{n} \diamond w\right)$. From Equations 7.2 and 7.3, we know that for almost all random functions $w \in \mathcal{F}^{n}$ that $\operatorname{Pr}_{C_{n}}\left(G^{n} \diamond w\right)$ will be bounded from above by the value $K_{n}$, and below by the value $L_{n}$.

We show, by means of a simple averaging argument, that it must be the case that for a large fraction of the $h \in H_{n}$ (a fraction greater than $\epsilon(n)+\frac{1}{n^{2 c}}$ ) it is the case that $\operatorname{Pr}_{C_{n}}\left(G^{n} \bigcirc h\right)$ falls outside of the range $\left[L_{n}, K_{n}\right]$. We can then distinguish between $H^{n}$ and $\mathcal{F}^{n}$ by computing $B(w)$, and accepting if $B(w)$ is greater than $K_{n}$ or less than $L_{n}$. Because the sets $S^{n}$ and $T^{n}$ contain all of the functions $w$ in $\mathcal{F}^{n}$, for which $\operatorname{Pr}_{C_{n}}\left(G^{n} \diamond w\right)$ is less than or greater than $L_{n}$ and $K_{n}$ respectively, and since $\frac{\left|S^{n}\right|}{\left|\mathcal{F}^{n}\right|}<\frac{1}{n^{\gamma^{c} c}}$ and $\frac{\left|T^{n}\right|}{\left|\mathcal{F}^{n}\right|}<\frac{1}{n^{7 c}}$; we know that about a fraction $\frac{2}{n^{\tau c} c}$ of the $w$ will be accepted by $B(n)$.

We construct a probabilistic circuit $A_{n}$ such that

$$
A_{n}(w)=\frac{1}{n^{b}} \sum_{i=1}^{n^{b}} C_{n}\left(g_{i} \diamond_{k_{i}^{4} \bullet k_{i}^{2}}^{n} w\right)
$$

where $g_{1}, \ldots, g_{n^{b}} \in G^{n}$ and $k_{1}^{1}, k_{1}^{2}, \ldots, k_{n^{b}}^{1}, k_{n^{b}}^{2} \in\{0,1\}^{n}$ are randomly chosen. Let $\kappa(n)$ be the length of the key of $\mathrm{H}^{n}$, and set $\alpha>1$ so that $n^{\alpha}>\kappa(n)$. Using the Chernoff Bound, $b$ is chosen large enough such that:
and

By Lemma 5.5 we derandomize $A_{n}$ to get the circuit $B_{n}$ for all but $\frac{1}{2^{\alpha} \alpha}$ of the $w \in \mathcal{F}^{n}$ :

$$
\begin{equation*}
\left|B_{n}(w)-\operatorname{Pr} C_{n}\left(\mathrm{G}^{n} \diamond w\right)\right|<\frac{1}{n^{7 c}} \tag{7.4}
\end{equation*}
$$

and for all of the $h_{k}^{n} \in H^{n}$ :

$$
\begin{equation*}
\left\lvert\, B_{n}\left(h^{n}\right)-\operatorname{Pr}\left(C_{n}\left(G^{n} \diamond h_{k}^{n}\right) \left\lvert\,<\frac{1}{n^{7 c}}\right.\right.\right. \tag{7.5}
\end{equation*}
$$

since for each $k \in\{0,1\}^{\kappa(n)}$ the probability of $h_{k}^{n} \in H^{n}$ is at least $\frac{1}{2^{\kappa(n)}}>\frac{1}{2^{n^{\alpha}}}$.

Let $K_{n}^{\prime}=\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)+\frac{1}{n^{2 c}}$, and let $L_{n}^{\prime}=\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)-\frac{1}{n^{7 c}}$. We now show that

$$
\operatorname{Pr}_{h^{n} \in H^{n}}\left[B_{n}\left(h^{n}\right) \geq K_{n}^{\prime} \vee B_{n}\left(h^{n}\right) \leq L_{n}^{\prime}\right] \geq \epsilon(n)+\frac{1}{n^{2 c}}
$$

The argument is an averaging argument. We first give an intuitive version of the argument, and then we give all of the details.

We now present the intuitive version of the argument. In this argument we essentially assume that any two values which differ by a value of at most $\frac{1}{\operatorname{poly(n)}}$, for some polynomial poly, are equivalent.

We remember that $B_{n}\left(h^{n}\right)$ is an approximation of $\operatorname{Pr}_{C_{n}}\left(G^{n} \diamond h^{n}\right)$, and intuitively the two values can be treated as the same. We know that $\left|\operatorname{Pr}_{C_{n}}\left(G^{n} \diamond H^{n}\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right|>$ $\epsilon(n) \delta(n)$. We partition the functions in $\mathrm{H}^{n}$ into two groups. This first group contains those $h^{n} \in H^{n}$ for which $\operatorname{Pr}_{C_{n}}\left(G^{n} \diamond h^{n}\right) \approx \operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)$. The second group contains the rest of $h^{n} \in H^{n}$ not in the first group. We then observe, by equation 7.1, that for each $h^{n} \in H^{n}$, the value $\left|\operatorname{Pr}_{C_{n}}\left(G \diamond h^{n}\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right|$ is less than $\delta(n)$. Using this fact we point out that for a random $h^{n} \in H^{n}$ it must be the case that $h^{n}$ is in the second group with a probability of at least $\epsilon(n)$, or otherwise $\left|\operatorname{Pr}_{C_{n}}\left(G^{n} \diamond H^{n}\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right|<\epsilon(n) \delta(n)$.

We know present the exact argument. Assume for contradiction that $\operatorname{Pr}_{h^{n} \in H^{n}}\left[B_{n}\left(h^{n}\right) \geq\right.$ $\left.K_{n}^{\prime} \vee B_{n}\left(h^{n}\right) \leq L_{n}^{\prime}\right]<\epsilon(n)+\frac{1}{n^{2 c}}$. Let $\mathcal{K}^{n} \subset H^{n}$ be the set of functions $h^{n} \in H^{n}$, for which $B_{n}\left(h^{n}\right) \geq K_{n}^{\prime}$ or $B_{n}\left(h^{n}\right) \leq L_{n}^{\prime}$, and let $\overline{\mathcal{K}}^{n}$ be its complement.

We now get a contradiction by a simple averaging argument.

$$
\begin{aligned}
& \left|{\underset{C}{n}}_{\operatorname{Pr}}[G \diamond H]-\operatorname{Pr}_{C_{n}}\left[\mathcal{F}^{n}\right]\right| \\
& \leq \sum_{h^{n} \in H^{n}}\left(\left|\operatorname{Pr}_{C_{n}}\left[G \diamond h^{n}\right]-\operatorname{Pr}_{C_{n}}\left[\mathcal{F}^{n}\right]\right| \operatorname{Pr}_{f \in \mathrm{H}^{n}}\left[f=h^{n}\right]\right) \\
& =\sum_{h^{n} \in \mathcal{K}^{n}}\left(\left|\operatorname{Pr}_{C_{n}}\left[G \diamond h^{n}\right]-\operatorname{Pr}_{C_{n}}\left[\mathcal{F}^{n}\right]\right| \operatorname{Pr}_{f \in H^{n}}\left[f=h^{n}\right]\right) \\
& +\sum_{h^{n} \in K^{n}}\left(\left\lfloor\left|\operatorname{Pr}_{C_{n}}\left[G \diamond h^{n}\right]-\operatorname{Pr}_{C_{n}}\left[\mathcal{F}^{n}\right]\right| \operatorname{Pr}_{f \in H^{n}}\left[f=h^{n}\right]\right)\right. \\
& \leq \sum_{h^{n} \in K^{n}}\left(\left|{\underset{C}{n}}^{\operatorname{Pr}}\left[G \diamond h^{n}\right]-\operatorname{Pr}_{C_{n}}\left[\mathcal{F}^{n}\right]\right| \operatorname{Pr}_{f \in \mathcal{H}^{n}}\left[f=h^{n}\right]\right)
\end{aligned}
$$

$$
\begin{align*}
& \quad+\sum_{h^{n} \in \mathcal{K}^{n}}\left(\left(\left|B_{n}\left(h^{n}\right)-\operatorname{Pr}_{C_{n}}\left[\mathcal{F}^{n}\right]\right|+\frac{1}{n^{7 c}}\right) \operatorname{Pr}_{f \in H^{n}}\left[f=h^{n}\right]\right) \\
& \leq \sum_{h^{n} \in \mathcal{K}^{n}}\left(\left|\underset{C_{n}}{\operatorname{Pr}}\left[G \diamond h^{n}\right]-\underset{C_{n}}{\operatorname{Pr}}\left[\mathcal{F}^{n}\right]\right| \underset{f \in \mathcal{H}^{n}}{\operatorname{Pr}_{r}}\left[f=h^{n}\right]\right)+\left(1-\epsilon(n)-\frac{1}{n^{2 c}}\right) \frac{1}{n^{4 c}}  \tag{7.6}\\
& \leq \sum_{h^{n} \in \mathcal{K}^{n}}\left(\left(\delta(n)+\frac{1}{n^{6 c}}\right)\left(\operatorname{Pr}_{f \in H^{n}}\left[f=h^{n}\right]\right)\right)+\frac{1}{n^{4 c}}  \tag{7.7}\\
& \leq\left(\epsilon(n)+\frac{1}{n^{2 c}}\right)\left(\delta(n)+\frac{1}{n^{6 c}}\right)+\frac{1}{n^{4 c}} \\
& <\epsilon(n) \delta(n)+\frac{1}{n^{c}} . \tag{7.8}
\end{align*}
$$

The inequality in line 7.6 follows from two facts. First that, by assumption, the probability that a random $h^{n} \in H^{n}$ is in $\overline{\mathcal{K}}^{n}$ is $1-\epsilon(n)-\frac{1}{n^{2 c}}$. Second, that for each $h^{n}$ in $\overline{\mathcal{K}}^{n}$, it is the case that $K_{n}>B_{n}\left(h^{n}\right)>L_{n}$, and this implies that $\left|B_{n}\left(h^{n}\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right|<\frac{1}{n^{5 c}}$ by the definitions of $K_{n}$ and $L_{n}$.

The inequality in line 7.7 follows from two facts. First, by assumption, the probability that a random $h^{n} \in H^{n}$ is in $\mathcal{K}^{n}$ is $\left(\epsilon(n)+\frac{1}{n^{2 c}}\right.$. Second, by equation 7.1, it is the case that for each $h^{n} \in H^{n}$ that $\left|\operatorname{Pr}_{C_{n}}\left(\mathbf{G} \diamond h^{n}\right)-\operatorname{Pr}_{C_{n}}\left(\mathcal{F}^{n}\right)\right|<\delta(n)+\frac{1}{n^{c c}}$.

The inequality in line 7.8 contradicts the fact that $\left|\operatorname{Pr}_{C_{n}}[G \diamond H]-\operatorname{Pr}_{C_{n}}\left[\mathcal{F}^{n}\right]\right| \geq \epsilon(n) \delta(n)+$ $\frac{1}{n^{c}}$, and therefore it must be the case that $\operatorname{Pr}_{h^{n} \in H^{n}}\left[B_{n}\left(h^{n}\right) \geq K_{n}^{\prime} \vee B_{n}\left(h^{n}\right) \leq L_{n}^{\prime}\right] \geq$ $\epsilon(n)+\frac{1}{n^{2 c}}$.

We create the decision circuit $\bar{B}_{n}(w)$ which accepts $w$ iff $B_{n}(w) \geq K_{n}^{\prime}$ or if $B_{n}(w) \leq$ $L_{n}^{\prime}$.

$$
\begin{aligned}
\left|\operatorname{PBr}_{B_{n}}\left(H^{n}\right)-\underset{B_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right)\right| & \geq \epsilon(n)+\frac{1}{n^{2 c}}-\underset{B_{n}}{\operatorname{Pr}}\left(\mathcal{F}^{n}\right) \\
& \geq \epsilon(n)+\frac{1}{n^{2 c}}-\frac{2}{n^{7 c}}-\frac{1}{2^{n^{\alpha}}} \quad \text { (Equations 7.2, 7.3 \& sampling error) } \\
& \geq \epsilon(n)+\frac{1}{n^{3 c}}
\end{aligned}
$$

Notice that the second inequality follows from constraints on the probability that a random function $w \in \mathcal{F}^{n}$ is in the set $S^{n}$ defined in equation 7.2. This is because, for all sufficiently large $n, K_{n}^{\prime}$ is greater than the value of $K_{n}$. As previously seen, the
probability that a random $w \in \mathcal{F}^{n}$ has the property that $\operatorname{Pr}_{C_{n}}\left(G^{n} \diamond w\right)>K_{n}$, is the probability that $w$ is contained in the set $S^{n}$. But, we know that the probability of a random $w \in \mathcal{F}^{n}$ being in the set $S^{n}$ is less than $\frac{1}{n^{\gamma c}}$. Clearly it follows that the probability that a random $w \in \mathcal{F}^{n}$ has the property that $\operatorname{Pr}_{C_{n}}\left(G^{n} \diamond w\right)>K_{n}^{\prime}$ is smaller than $\frac{1}{n^{\text {c }}}$, for the set of functions with this property is necessarily a subset of $S^{n}$. Clearly similar arguments apply for when $\operatorname{Pr}_{C_{n}}\left(\mathrm{G}^{n} \diamond w\right)<L_{n}$.

Finally, since $B_{n}(w)$ approximates $\operatorname{Pr}_{C_{n}}\left(G^{n} \circ w\right)$ to within a value of $\frac{1}{2^{n^{\alpha}}}$, for all but $\frac{1}{2^{n^{d}}}$ of the $w \in \mathcal{F}^{n}$, and since it still holds that for all sufficiently large $n$ that $K_{n}^{\prime}-\frac{1}{2^{n^{\alpha}}}$ is greater than $K_{n}$, there can only be a fraction $\frac{2}{n^{7^{c}}}+\frac{1}{2^{n^{\alpha}}}$ of the $w \in \mathcal{F}^{n}$ for which $B^{\prime}(w)$ accepts. The first $\frac{2}{n^{\top c}}$ of these functions are the ones in $S^{n}$ and $T_{n}$. The last $\frac{1}{2^{n^{\alpha}}}$ of these functions are the functions $w$ in which $B(w)$ 's approximation of $\operatorname{Pr}_{C_{n}}\left(G^{n} \diamond w\right)$ is off by more than a factor of $\frac{1}{n^{2 c}}$.

We let $\Upsilon_{n}=\bar{B}_{n}$, and we are finished.

### 7.3 Questioning The Model

In this section we have shown how to construct a PRFG from partially secure PRFGs, and have thus reached one of our research goals. However, the initial purpose of studying constructions which increased the security of partially secure PRFGs was to attempt to provide evidence that such constructions would amplify security when applied to completely insecure function generators (what would be considered a 0-secure PRFG in our model). Or in lay man's terms, we wanted to provide evidence that there was reason to believe that the constructions used in real world cryptographic permutation generators, such as DES and the AES candidates, actually do increase security. Unfortunately, our research program has fallen short in this respect because there is little reason to believe that the construction studied in this chapter would have any security amplifying properties on insecure function generators. This is based on the belief that, to the best
of the authors' knowledge, there exists no "real world" cryptographic system, in public use, that is based on the security amplyfiyng properties of the exlcusive-or function. Further, the author predicts that the consensus among cryptologists would be that it would be shocking to discover that the exclusive-or function can be used in the fashion described in this chapter to increase security. Therefore, while there is no formal evidence backing the claim that the $\diamond$ operator does not provide security amplification in the "realworld", there is sufficient enough evidence to have us question the model we are working in. Therefore, proving that certain constructions are security increasing when applied to partially secure PRFGs may provide little evidence that the construction will have security amplifying properties when applied to trivial generators.

## Chapter 8

## Conclusions and Open Questions

In this chapter we summarize our findings, and propose some interesting new research directions.

### 8.1 Conclusions

Our goal was to understand in which ways insecure functions generators might be combined in order to produce function generators which are conjectured to be pseudo-random. We developed a list of possible constructions which are based on "natural" operators which are commonly used in the cryptographic community.

We then considered each construction in the model of partially secure pseudo-random generators. For each proposed construction we were then able to either give evidence which suggests it is of little practical use, or classify the construction as security increasing, preserving, or decreasing. We extended the result of Luby and Rackoff [12] and showed that we could achieve a $1-\frac{1}{\log n}$ secure generator from a partially secure PRPG. We then showed that the we could achieve the same security using PRFGs and the operator, as proposed by Luby and Rackoff [13]. Finally, we have shown that we can achieve a completely secure PRFG from a partially secure PRFG based on the $\diamond$ operator generator. This last result has led us to question the belief that a construction
which is security increasing in the weakly secure pseudo-random model is also security increasing when applied to insecure generators. However, it has answered in the positive the question of whether or not it is possible to use a "natural" construction to combine partially secure PRFGs into a secure PRFG.

### 8.2 Open Questions

In this section we will first review the open questions which have appeared previously in this text. We will then describe one problem which we have spent a significant amount of time thinking about, but for which we have made little progress on. Finally, we will suggest a new direction for future research which we believe is of strong practical importance, and if successful may dramatically change the way conjectured PRPGs and PRFGs are designed.

### 8.2.1 Previously Stated Ofen Questions

We remind the reader that in Chapter 4 we could not show that the construction based on the XOR operator was security preserving when applied to generators of security greater than $\frac{1}{2}$. We then mentioned that for similar reasons we were unable to determine whether a construction based on the $\square$ and XOR operators was security increasing, preserving, or possibly diminishing. We restate the questions here, for completeness.

Open Question 8.1 (Restatement of Open Question 4.1) Is the construction $\boldsymbol{F}(\boldsymbol{\alpha})=$ $\left(G_{m(n)} \oplus \cdots \oplus G_{1}\right)(\alpha)$ security increasing, when the generators $G_{i}$ are restricted to being $1-\delta$ secure for $\delta<\frac{1}{2}$.

Open Question 8.2 (Restatement of Open Question 4.2) Is the construction $\mathrm{F}(\alpha)=$ $\left(\left(G_{2 m(n)} \square G_{2 m(n)-1}\right) \oplus \cdots \oplus\left(G_{2} \square G_{1}\right)\right)(\alpha)$ security increasing, security preserving or security decreasing.

### 8.2.2 Adaftive vs. Non-Adaftive Security

We say that a PRFG is non-adaptively secure if it is secure under a weakened version of the standard definition of secure PRFGs. The definition is weakened by disallowing the adversary from making adaptive queries to the function-oracle. Therefore, the adversary can still make a polynomial in $n$ number of queries, but the queries must all be made at the same time. This prevents the adversary from making queries based on information learned from the results of previous queries.

We give a simple example of a generator which is non-adaptively secure, but which is adaptively insecure. Let $\widehat{G}=\left\{\widehat{G^{n}}:\{0,1\}^{\kappa(n)} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ be an adaptively secure PRFG. We construct a non-adaptively secure generator $G=\{G:$ $\left.\{0,1\}^{\kappa(n)} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$, by taking for each $k \in\{0,1\}^{\kappa(n)}$ the function $\hat{g}_{k}^{n} \in \widehat{G^{n}}$, and using it to construct a new function $g_{k}^{n} \in \mathrm{G}^{n}$ : for all $\boldsymbol{x} \neq \hat{g}_{k}^{n}(\overline{0})$, for some $x \in\{0,1\}^{n}$ and $k \in\{0,1\}^{\kappa(n)}$ we set $g_{k}^{n}(x)=\hat{g}_{k}^{n}(x)$, but we set $g_{k}^{n}\left(\hat{g}_{k}^{n}(\overline{0})\right)=\overline{0}$.

The question is if there is some construction based on "natural" operators which allows us construct an adaptively secure PRFG, from non-adaptively secure generators. Specifically, we say that a construction provides adaptive security if, for all possible nonadaptively secure PRFGs used in the construction, the resulting generator is adaptively secure.

We note that using a non-adaptively secure PRFG we can easily construct a one-way function. Then using the standard constructions we can construct an adaptively secure PRFG. Therefore, when we limit our operators and constructions to be natural, we mean that the operators are ones which can easily be used to combine generators. Further, the operators are normally computable in linear time.

As we just mentioned, the existence of non-adaptively secure PRFGs implies the existence of adaptively secure PRFGs. Therefore, the point of this question is to develop an efficient construction which gives adaptively secure generators from non-adaptive ones.

We will now rule out one feasible construction, and propose several others.

## Data Dependent Re-Keying Does Not Work

Let $H=\left\{H^{n}:\{0,1\}^{n} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ and $G=\left\{G^{n}:\{0,1\}^{n} \times\{0,1\}^{n} \rightarrow\right.$ $\left.\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ be non-adaptively secure PRFGs. Let the data-dependent re-keying construction $F=\left\{F^{n} \mid \in \mathbb{N}\right\}$ be defined as follows:

$$
\mathrm{F}^{n}(k, x)=\mathrm{G}^{n}\left(\mathrm{H}^{\mathrm{n}}(k, x), x\right),
$$

for each $n$.
We will now construct specific generators $\mathrm{G}^{n}$ and $\mathrm{H}^{n}$ which are non-adaptively secure, but $F$ will not be adaptively secure. Let $\widehat{\boldsymbol{H}}=\left\{\hat{\mathbf{H}}^{n}:\{0,1\}^{n} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ and $\widehat{\mathrm{G}}=\left\{\widehat{\mathrm{G}}^{n}:\{0,1\}^{n} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ be adaptively secure PRFGs. For each $n$ we construct $\mathrm{G}^{n}$ by setting for each $x \in\{0,1\}^{n}$ and each $k \in\{0,1\}^{n} \backslash\{\overline{0}\}$ the function $g_{k}^{n}(x)=\hat{g}_{k}^{n}(x)$; and then for each $x \in\{0,1\}^{n}$ setting $g_{0}^{n}(x)=\overline{0}$. For each $n$ and each $k \in\{0,1\}^{n}$ let $\alpha_{k}^{n}=g_{i_{k}^{n}(\overline{0})}^{n}(\overline{0})$. For each $n$ we construct $H^{n}$ by setting for each $k \in\{0,1\}^{n}$ and each $x \in\{0,1\}^{n} \backslash\left\{\alpha_{k}^{n}\right\}$ the function $h_{k}^{n}(x)=\hat{h}_{k}^{n}$, and setting $h_{k}^{n}\left(\alpha_{k}^{n}\right)=\overline{0}$. Notice that $\mathrm{G}^{\boldsymbol{n}}$ is adaptively secure, and that $\mathbf{H}^{n}$ is non-adaptively secure.

Now construct $F$ from $H$ and $G$, as in the data-dependent re-keying construction given above. Notice that for each $n$ and for each $k \in\{0,1\}^{n}$ the function $f_{k}^{n} \in \mathrm{~F}^{n}$ has the property that $f_{k}^{n}\left(f_{k}^{n}(\overline{0})\right)=\overline{0}$, and therefore the generator F is clearly not adaptively secure. Further, since $G$ is adaptively secure, we see that the construction does not even preserve adaptive security.

## Proposed Constructions

We conjecture that any of the constructions presented in Chapter 4 which are security improving might also provide adaptive security. In particular, we have not been able to give examples of constructions of any non-adaptively secure PRFGs which, when used in one of the security improving constructions of Chapter 4, gives a generator which is not adaptively secure. However, we have failed to give either a proof of adaptive security or
a counterexample when we restrict the constructions so that only two constituent nonadaptively secure generators are used, and when only one adaptive query is performed.

We also conjecture that if $K$ is a 1 -wise independent permutation generator, and $G$ is a non-adaptively secure pseudo-random permutation generator then the construction $G \circ K$ is an adaptively secure pseudo-random permutation generator. We note that composition is not commutative, and the construction KoG is not conjectured to provide adaptive security, for there exist specific examples of pair-wise independent permutation generators and non-adaptively secure PRFGs which when used in the above construction do not give adaptive security.

For example if we let $K=\left\{K^{n}:\{0,1\}^{2 n} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ be a pair-wise independent permutation generator, where $K(a \bullet b, x)=a x+b ;|a|=|b|=n ; a \neq \overline{0}$; and the operations are performed over the field $G F\left(2^{n}\right)$. Let $H=\left\{H^{n}:\{0,1\}^{\kappa(n)} \times\{0,1\}^{n} \rightarrow\right.$ $\left.\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ be an adaptively secure PRPG. For each $n$, and for each $k \in\{0,1\}^{\kappa(n)}$, we will modify $h_{k}^{n} \in H$ and thereby make $H$ non-adaptively secure. Let $\alpha_{h_{k}^{n}}=\frac{h_{k}^{n}(0)-h_{k}^{n}(1)}{h_{k}^{n}(0)-h_{k}^{n}(2)}$. Modify $h_{k}^{n}$ by making the changes: $h_{k}^{n}\left(\alpha_{h_{k}^{n}}\right) \leftarrow h_{k}^{n}(0), h_{k}^{n}\left(\alpha_{h_{k}^{n}}+1\right) \leftarrow h_{k}^{n}(1)$ and $h_{k}^{n}\left(\alpha_{h_{k}^{n}}+\right.$ 2) $\leftarrow h_{k}^{n}(2)$. Clearly, $\mathrm{H}^{n}$ is no longer adaptively secure as one can easily construct an adversary which accepts a function $f^{n} \in \mathcal{F}^{n}$ iff $f^{n}(0)=f^{n}\left(\alpha_{f^{n}}\right), f^{n}(1)=f^{n}\left(\alpha_{f^{n}}+1\right)$ and $f^{n}(2)=f^{n}\left(\alpha_{f^{n}}+2\right)$. However, we can also construct an adversary which distinguishes KoH adaptively. Given a function $f^{n} \in \mathcal{F}^{n}$ our adversary will distinguish it by performing computing $\alpha_{f^{n}}$, and then computing $\beta_{f^{n}}=\frac{f^{n}\left(\alpha_{f^{n}}\right)-f\left(\alpha_{f^{n}}+1\right)}{f\left(\alpha_{f^{n}}\right)-f\left(\alpha_{\left.f^{n}+2\right)}\right.}$. The adversary accepts iff $\beta_{f n}=\alpha_{f n}$. It is a relatively simple argument to show that this will occur with probability 1 if $f^{n}$ was drawn from $\mathrm{K}^{n} \circ \mathrm{H}^{n}$, whereas it will occur with a negligible probability if $f$ was drawn from $\mathcal{F}^{n}$. This construction can be generalized to show that for any $k$-wise independent function generator $G=\left\{G^{n}:\{0,1\}^{k n} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n}: n \in \mathbb{N}\right\}$, where $g_{a_{k} \circ \cdots \circ a_{1}}^{n}(x)=a_{k} x^{k-1}+a_{k-1} x^{k-2}+\cdots+a_{1} x^{0}$ and $\left|a_{i}\right|=|x|=n$; and a generator $H^{\prime}$ which is constructed in a generalized manner of the generator $H$ presented above, that the generator $G \circ H^{\prime}$ is non-adaptively secure.

### 8.2.3 A Combinatorial Security Model

As mentioned earlier, the existence of PRFGs cannot be proven without indirectly proving that $\mathcal{P} \neq \mathcal{N} \mathcal{P}$. Given the apparent difficulty of such a task, it would seem unlikely that any modern cryptographic systems will be proven secure in the immediate future. However, this does not prevent the development of a theory which can be used to help cryptographers design cryptographic systems. Currently, block-cipher designers only seem to have intuition to guide their designs. Further, the number of designs often seem to be as numerous as the number of block-cipher designers. The result is that there are many proposed block-ciphers, none of which can be compared. We propose the development of a theory to help with this process.

This research direction is motivated by the current Advanced Encryption Standard competition. As mentioned earlier, there are many proposed function and permutation generators which are currently being considered, and there is no quantifiable method for comparing the different designs. The end result seems to be that the main criteria for choosing a winner of this competition is based on the pedigree, within the cryptographic community, of the designers. Yet many of these designs are based on composing trivial permutation generators numerous times to result in a permutation generator which appears to be secure. A theory such as the one we are proposing would allow for these generators to be compared in a quantitative, and hopefully meaningful sense. Information about the AES competition and the currently proposed ciphers is available at the AES' official web page http://csrc.nist.gov/encryption/aes/aes_home.htm.

The idea behind our proposed theory would be to combine function generators which have some trivial security properties together in some construction, such as the security increasing constructions proposed in Chapter 4. We would then conjecture that if the resulting generator has some specific combinatorial property, then it is a pseudo-random generator.

An example of how such a theory might look is the following. We would like to stress
that these examples are most likely incorrect, and are included only to help bring light to our proposed research direction.

Definition 8.1 (Trivial Permutation Generator (Example)) A function generator $T=\left\{T^{n}:\{0,1\}^{\kappa(n)} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ is considered trivial, if for each $n$, each key $k \in\{0,1\}^{\kappa(n)}$ and each $x \in\{0,1\}^{n}$, every bit of $t_{k}^{n}(x)$ is dependent on no more than $\mathcal{O}(\log n)$ bits of $x$.

An example of a construction which would be used to combine trivial function generators is repeated composition. Finally, an example of a conjecture of security based on a combinatorial principle is given below.

Conjecture 8.1 (Combinatorial Security (Example)) Let $T$ be a trivial function generator. Construct $\mathrm{F}=\mathrm{T} \circ \ldots \circ \mathrm{T}$, so that enough compositions are performed so that F is an almost pair-wise independent function generator. Then there exists a polynomial p such that $\widehat{\mathrm{F}}=\{\underbrace{\mathrm{F}^{n} \circ \ldots \circ \mathrm{~F}^{n}}_{p(n)} \mid n \in \mathbb{N}\}$ is a $P R F G$.

Note the importance of $T$ being a trivial function generator, for there are standard constructions of pair-wise independent permutation generators which are clearly not secure. For example, if we consider the generator $\mathrm{K}(a \bullet b, x)=a x+b$, where $|x|=|a|=|b|=n$, $a \neq \overline{0}$ and the operations are performed over the field $G F\left(2^{n}\right)$, then it is clearly insecure. Further, this generator would be insecure under an arbitrary number of compositions, as the resulting permutations will always be linear. Therefore, the triviality requirement would exclude the use of a generator such as $K$.

Notice that this construction is based on what people currently do to construct ciphers, and is not some notion we have invented. Both DES and almost all of the AES candidates can by described at an abstract level as being the result of composing together many completely insecure (or trivial) permutation generators. What differs in their design is the types of trivial permutation generators they use, and how many times the
generators are composed. Notice, that is exactly the types of differences which could be compared in our proposed theory.

One difference between our proposed theory, and the construction of known ciphers, such as DES and the AES candidates, is that the known ciphers are compositions of permutation generators, whereas in our construction we consider compositions of function generators. Initially, we thought that the trivial generators should be restricted to being permutation generators. However, in hindsight we realize that this restriction is probably not important. The important goal is that we are able to achieve an almost pair-wise independent function generator from some polynomial number of compositions.

We believe that using a theory such as the one proposed, designers could easily quantify and compare tradeoffs which were made during design. For example, arguments could be made for using different types of initial trivial generators. However, we could now argue tradeoffs based on how fast specific trivial generators are, and the size of the construction in which they would need to be embedded, in order for the construction to be conjectured secure. Further, it might be possible to prove security preserving reductions between different generators in this model. This would allow us to establish security classes, where certain generators would be pseudo-random only if other generators are pseudo-random. Finally, because the theory is not based on any complexity theoretic assumptions, it may be easier than in the complexity theoretic models to find counter examples, and thus disprove the theory, if the theory turned out to be incorrect.

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## Appendix A

## Proof Of Security for the Generator Described in 5.1.1

In this section we present the proof of security for the PRPG constructed in section 5.1.1. We first remind the reader of the construction presented, and then give the proof of its security.

## A. 1 Construction of the $1-\delta$ Secure Generator

To simplify the presentation we assume that $\delta$ and $\epsilon$ are of the form $\frac{1}{2^{c}}$ or $1-\frac{1}{2^{c}}$, for some constant c. Let $\widehat{\mathrm{G}}=\left\{\widehat{\mathrm{G}}^{n}:\{0,1\}^{(n)} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ be a PRPG. We construct $G=\left\{G^{n}:\{0,1\}^{\ell(n)+c} \times\{0,1\}^{n} \rightarrow\{0,1\}^{n} \mid n \in \mathbb{N}\right\}$ to be $1-\delta$ secure in two steps. We will describe the construction of $\mathrm{G}^{n}$, and note that the construction of $\mathrm{H}^{n}$ is similar.

First we set a fraction $\delta$ of the keys to correspond to the identity permutation, and the remainder to correspond to permutations chosen from $\widehat{\mathbf{G}}^{n}$. This is done in two different fashions dependent on the form of $\delta$ as described below:

Case $1\left(\delta=1-\frac{1}{2^{c}}\right)$ For each $k \in\{0,1\}^{\ell(n)+c}$ we set the permutation $g_{k}^{n} \in G^{n}$ to be the
identity permutation if the first $c$ bits of $k$ are not all 0 ; otherwise let $\bar{k}$ be the last $\ell(n)$ bits of $k$, and set $g_{k}^{n}=\hat{g}_{k}^{n}$, for $\hat{g}_{k}^{n} \in \widehat{G}^{n}$.

Case $2\left(\delta=\frac{1}{2^{c}}\right.$ ) For each $k \in\{0,1\}^{\ell(n)+c}$ we set the permutation $g_{k}^{n} \in \mathrm{G}^{n}$ to be the identity permutation if the first $c$ bits of $k$ are all 0 ; otherwise let $\bar{k}$ be the last $\ell(n)$ bits of $k$, and set $g_{k}^{n}=\hat{g}_{\hat{k}}^{n}$, for $\hat{g}_{\bar{k}}^{n} \in \widehat{G}^{n}$.

The second step in the construction of $\mathrm{G}^{n}$ is to ensure that for each $k \in\{0,1\}^{\ell(n)+c}$ we set the value of $g_{k}^{n}(\overline{0})$ in one of two manners, depending on the form of $\delta$. We describe this transformation below:

Case $1\left(\delta=1-\frac{1}{2 c}\right)$ For all $k$ set the first $c$ bits of $g_{k}^{n}(\overline{0})$ to 0 ; the last $n-c$ bits remain as they were in $g_{k}^{n}(\overline{0})$.

Case $2\left(\delta=\frac{1}{2^{c}}\right)$ For all $k$ ensure that not all of the first $c$ bits of $g_{k}^{n}(\overline{0})$ are 1 . If they are, set them to be a member of the set $\{0,1\}^{c} \backslash\left\{1^{c}\right\}$ chosen uniformly at random ${ }^{1}$; the last $n-c$ bits remain as they were in $g_{k}^{n}(\overline{0})$.

Notice that in order to maintain the permutation property of $\mathbf{G}^{\boldsymbol{n}}$ we can simply store both the initial and modified value of $g_{k}^{n}(\overline{0})$. Should a query ever be made to the initial preimage of the modified value of $g_{k}^{n}(\overline{0})$, then we respond with the initial value of $g_{k}^{n}(\overline{0})$.

## A. 2 Proof that $G$ is $1-\delta$ Secure

Theorem A. 1 The pseudo-random permutation generator $G$ is $1-\delta$ secure.

Proof: Let $A$ be an adversary such that:

$$
\begin{equation*}
\left|\operatorname{Pr}_{f^{n} \in \mathcal{F}{ }^{n}}\left(A^{f^{n}}=1\right)-\operatorname{Pr}_{g^{n} \in G^{n}}\left(A^{g^{n}}=1\right)\right| \geq \delta+\frac{1}{n^{d}}, \tag{A.1}
\end{equation*}
$$

[^10]for infinitely many $n$ and some constant $c>0$. We will show how to distinguish $\widehat{G}$ from random functions or derive a contradiction, and thereby prove the $1-\delta$ security of $G$. The intuition is as follows, we expect the adversary, $A$, to be able to distinguish the identity permutations in $\mathbf{G}^{\boldsymbol{n}}$ from random functions in $\mathcal{F}^{\boldsymbol{n}}$. However, the remaining functions in $\mathrm{G}^{n}$ should not be distinguishable from random functions $f^{n} \in \mathcal{F}^{n}$, which have the property that $f^{n}(\overline{0})$ 's first $c$ bits are either all 0 , or not all 1 , dependent on whether $\delta \geq \frac{1}{2}$ or $\delta<\frac{1}{2}$ respectively. Further, accepting any of the $f^{n} \in \mathcal{F}^{n}$ which do not conform to the above description will only decrease the adversaries distinguishing probability. Therefore, it would appear that the distinguishing probability should not be better than $\delta$, and this provides the contradiction.

We partition the set $\{0,1\}^{n}$ into two sets $X^{n}$ and $Y^{n}=\{0,1\}^{n} \backslash X^{n}$. The contents of the set $X^{n}$ is dependent on the form of $\delta$ as described below:

Case $1\left(\delta=1-\frac{1}{2^{c}}\right)$ Then $X^{n}=\left\{x \in\{0,1\}^{n} \mid\right.$ The first $c$ bits of $x$ are 0$\}$.

Case $2\left(\delta=\frac{1}{2^{c}}\right)$ Then $X^{n}=\left\{x \in\{0,1\}^{n} \mid\right.$ The first $c$ bits of $x$ are not all 1$\}$.

We use $X^{n}$ and $Y^{n}$ to distinguish between the functions $f^{n}$ which appear to have had the transformation on $f^{n}(\overline{0})$ applied to them, in which case $f^{n}(\overline{0}) \in X^{n}$, versus those which have not, in which case $f^{n}(\overline{0}) \in Y^{n}$. Finally we distinguish the identity permutations from random permutations by simply checking of $f^{n}(\overline{0})=\overline{0}$. By partitioning the functions in the distributions as in the manner described above, and assuming that $\operatorname{Pr}_{f n \in \mathcal{F} n}\left(A^{f^{n}}=\right.$ 1) $>\operatorname{Pr}_{g^{n} \in G^{n}}\left(A^{g^{n}}=1\right.$ ) (if this is not the case, then we can simply reverse the outputs of $A$ ), we can rewrite equation (A.1) as shown below:

$$
\begin{aligned}
& \operatorname{Pr}_{f^{n} \in \mathcal{F} n}\left[A^{f^{n}}=1 \mid f^{n}(\overline{0})=\overline{0}\right] \operatorname{Pr}_{f^{n} \in \mathcal{F n}}\left[f^{n}(\overline{0})=\overline{0}\right] \\
+ & \operatorname{Pr}_{f n \in \mathcal{F},}\left[A^{f^{n}}=1 \mid f^{n}(\overline{0}) \neq \overline{0} \& \& f^{n}(\overline{0}) \in X^{n}\right] \operatorname{Pr}_{f n \in \mathcal{F n}}\left[f^{n}(\overline{0}) \neq \overline{0} \& f^{n}(\overline{0}) \in X^{n}\right] \\
+ & \operatorname{Pr}_{f^{n} \in \mathcal{F} n}\left[A^{f^{n}}=1 \mid f^{n}(\overline{0}) \neq \overline{0} \& f^{n}(\overline{0}) \in Y^{n}\right] \operatorname{Pr}_{f^{n} \in \mathcal{Z}}\left[f^{n}(\overline{0}) \neq \overline{0} \& f^{n}(\overline{0}) \in Y^{n}\right]
\end{aligned}
$$

$$
\begin{aligned}
& -{\underset{g^{n} \in G^{n}}{\operatorname{Pr}}\left[A^{g^{n}}=1 \mid g^{n}(\overline{0})=\overline{0}\right] \underset{g^{n} \in G^{n}}{\operatorname{Pr}}\left[g^{n}(\overline{0})=\overline{0}\right]}_{-\operatorname{Pr}_{g^{n} \in G^{n}}\left[A^{g^{n}}=1 \mid g^{n}(\overline{0}) \neq \overline{0} \& g^{n}(\overline{0}) \in X^{n}\right]{\underset{g^{n} \in G^{n}}{ }}_{\operatorname{Pr}}\left[g^{n}(\overline{0}) \neq \overline{0} \& g^{n}(\overline{0}) \in X^{n}\right]}^{\left.-\operatorname{Pr}_{g^{n} \in G^{n}}\left[g^{g^{n}}=1 \mid g^{n}(\overline{0}) \neq \overline{0}\right) \neq \overline{0} \& g^{n}(\overline{0}) \in Y^{n}(\overline{0}) \in Y^{n}\right]} \\
& \geq \delta+\frac{1}{\boldsymbol{P}^{d}}
\end{aligned}
$$

We notice that for the set of random functions it is the case that:

$$
\begin{aligned}
& \operatorname{Pr}_{f^{n} \in \mathcal{F} n}\left[f^{n}(\overline{0})=\overline{0}\right]=\frac{1}{2^{n}} \\
& \operatorname{Pr}_{f^{n} \in \mathcal{F} n}\left[f^{n}(\overline{0}) \neq \overline{0} \& f^{n}(\overline{0}) \in X^{n}\right]=1-\delta-\frac{1}{2^{n}} \\
& \operatorname{Pr}_{f^{n} \in \mathcal{F} n}\left[f^{n}(\overline{0}) \neq \overline{0} \& f^{n}(\overline{0}) \in Y^{n}\right]=\delta .
\end{aligned}
$$

Similarly, by the construction of $\mathrm{G}^{\boldsymbol{n}}$ we know that

$$
\begin{aligned}
& \operatorname{Pr}_{g^{n} \in G^{n}}\left[g^{n}(\overline{0})=\overline{0}\right]=(\delta+\eta(n)) ; \\
& \operatorname{Pr}_{g^{n} \in G^{n}}\left[g^{n}(\overline{0}) \neq \overline{0} \& g^{n}(\overline{0}) \in X^{n}\right]=1-\delta-\eta(n) ; \\
& \operatorname{Pr}_{g^{n} \in G^{n}}\left[g^{n}(\overline{0}) \neq \overline{0} \& g^{n}(\overline{0}) \in Y^{n}\right]=0 ;
\end{aligned}
$$

for some negligible function $\eta(n)$. The quantity $\eta(n)$ comes from the observation that $\operatorname{Pr}_{\tilde{g}^{n} \in \bar{G}}\left(\hat{g}^{n}(\overline{0})=\overline{0}\right)$ must be at most $\eta(n)$, or $\widehat{G}$ is not pseudo-random.

Therefore, rewriting equation (A.1) again, using the above facts, we get:

$$
\begin{align*}
& \operatorname{Pr}_{f^{n} \in \mathcal{F}^{n}}\left[A^{f^{n}}=1 \mid f^{n}(\overline{0})=\overline{0}\right]\left(\frac{1}{2^{n}}\right)  \tag{A.2}\\
+ & \operatorname{Pr}_{f^{n} \in \mathcal{R}^{n}}\left[A^{f^{n}}=1 \mid f^{n}(\overline{0}) \neq \overline{0} \& f^{n}(\overline{0}) \in X^{n}\right]\left(1-\delta-\frac{1}{2^{n}}\right) \\
+ & \operatorname{Pr}_{f^{n} \in \mathcal{F}^{n}}\left[A^{f^{n}}=1 \mid f^{n}(\overline{0}) \neq \overline{0} \& f^{n}(\overline{0}) \in Y^{n}\right] \delta \\
- & \operatorname{Pr}_{g^{n} \in G^{n}}\left[A^{g^{n}}=1 \mid g^{n}(\overline{0})=\overline{0}\right] \delta+\eta \\
- & \operatorname{Pr}_{g^{n} \in G^{n}}\left[A^{g^{n}}=1 \mid g^{n}(\overline{0}) \neq \overline{0} \& g^{n}(\overline{0}) \in X^{n}\right](1-\delta-\eta) \\
\geq & \delta+\frac{1}{n^{d}}
\end{align*}
$$

We now use a lemma which formalizes our intuition that the adversaries should not be able to distinguish between random functions $f^{n} \in \mathcal{F}^{n}$ that have the property that $f^{n}(\overline{0}) \in X^{n}$ and those $g^{n} \in \mathrm{G}^{n}$ which are not the identity function. These have the form $g^{n}(\overline{0}) \in X^{n}$, by the construction of $\mathrm{G}^{n}$.

Lemma A. 1 For all constants $e \geq 0$ and sufficiently large $n$,

$$
\left|\operatorname{Pr}_{f^{n} \in \mathcal{F}{ }^{n}}\left[A^{f^{n}}=1 \mid f^{n}(\overline{0}) \neq \overline{0} \& f^{n}(\overline{0}) \in X^{n}\right]-\operatorname{Pr}_{g^{n} \in G^{n}}\left[A^{g^{n}}=1 \mid g^{n}(\overline{0}) \neq \overline{0} \& g^{n}(\overline{0}) \in X^{n}\right]\right|<\frac{1}{n^{e}} .
$$

We now take equation (A.2) and eliminate all negligible quantities. Further, we treat the two probabilities in Lemma A. 1 as equal because their difference is negligible. The result is:

$$
\operatorname{Pr}_{f^{n} \in \mathcal{F} \boldsymbol{n}}\left[A^{f^{n}}=1 \mid f^{n}(\overline{0}) \neq \overline{0} \& f^{n}(\overline{0}) \in Y^{n}\right] \delta-\operatorname{Pr}_{g^{n} \in G^{n}}\left[A^{g^{n}}=1 \mid g^{n}(\overline{0})=\overline{0}\right] \delta \geq \delta+\frac{1}{n^{d}} .
$$

This gives:

$$
\operatorname{Pr}_{f^{n} \in \mathcal{F} n}\left[A^{f^{n}}=1 \mid f^{n}(\overline{0}) \neq \overline{0} \& f^{n}(\overline{0}) \in Y^{n}\right]-\operatorname{Prr}_{g^{n} \in G^{n}}\left[A^{g^{n}}=1 \mid g^{n}(\overline{0})=\overline{0}\right] \geq 1+\frac{1}{\delta n^{d}},
$$

but this is a contradiction since the two probabilities are bounded to be between 0 and 1.

All that remains is to give a proof of Lemma A.1.
Proof of Lemma A.1: Suppose there exists a $e$ such that for infinitely many $\boldsymbol{n}$ the inequality stated in the lemma does not hold. We will construct an adversary $\widehat{\boldsymbol{A}}$ which can distinguish $\widehat{\mathbf{G}}^{n}$ from $\mathcal{F}^{n}$ with significant probability for infinitely many $n$.

Given a function $w, \widehat{A}$ will query $w(\overline{0})$, and if $w(\overline{0})=\overline{0}$ then $\widehat{A}$ will accept with probability $\frac{1}{2}$. Otherwise $\widehat{A}$ simulates $A$ running on $w$ exactly, except in one condition: if $A$ queries $w(\overline{0})$ and if $w(\overline{0}) \notin X^{n}$ then we modify it, dependent on $\delta$.

Case $1\left(\delta=1-\frac{1}{2^{c}}\right)$ Set the first $c$ bits of $w(\overline{0})$ to $\overline{0}$.

Case $2\left(\delta=\frac{1}{2^{c}}\right)$ Set the first $c$ bits of $w(\overline{0})$ to be a member of the set $\{0,1\}^{c} \backslash\left\{1^{c}\right\}$ chosen uniformly at random.

We maintain the permutation property of $\boldsymbol{w}$ by simply having $\widehat{A}$ store both the initial and modified value of $w(\overline{0})$. Should a query ever be made to the initial preimage of the modified value of $w(\overline{0})$, then $\widehat{A}$ responds with the initial value of $w(\overline{0})$. Finally, $\widehat{A}$ accepts iff $A$ accepts.

We notice that the probability of the event in which $w(\overline{\mathbf{0}})=\overline{\mathbf{0}}$ can be bounded in both distributions. For the set of random functions, $\operatorname{Pr}_{f \in \mathcal{P n}}(f(\overline{0})=\overline{0})=\frac{1}{2^{n}}$, and for the functions from the generator $\widehat{G}^{n}, \operatorname{Pr}_{g^{n} \in \bar{G}^{n}}\left(g^{n}(\overline{0})=\overline{0}\right) \leq \frac{1}{n^{e}}$, for all constants $e>0$. Since in both cases the event happens with negligible probability, the fact that $\widehat{A}$ accepts in both these cases with probability $\frac{1}{2}$, does not have any significant affect on $\widehat{A}$ 's distinguishing probability on the two distributions of functions. Therefore, we will ignore this factor for the remainder of the proof.

We observe that by the simulation of $A$ by $\widehat{A}$ and the construction of $G$ from $\widehat{G}$ that:

$$
\operatorname{Pr}_{g^{n} \in \mathbb{G}^{n}}\left[A^{g^{n}}=1 \mid g^{n}(\overline{0}) \neq \overline{0} \& g^{n}(\overline{0}) \in X^{n}\right]=\operatorname{Pr}_{\hat{g}^{n} \in \widetilde{\mathbf{G}}^{n}}\left[\widehat{A}^{\hat{g}^{n}}=1\right]
$$

Similarly,

$$
\operatorname{Pr}_{f^{n} \in \mathcal{F} n}\left[A^{f^{n}}=1 \mid f^{n}(\overline{0}) \neq \overline{0} \& f^{n}(\overline{0}) \in X^{n}\right]=\operatorname{Pr}_{f^{\prime} \in \mathcal{F}}\left[\widehat{A}^{f^{n}}=1\right]
$$

Therefore by assumption:

$$
\left|\operatorname{Pr}_{\hat{g}^{n} \in \widehat{\mathbf{G}}^{n}}\left[\widehat{A}^{\hat{g}^{n}}=1\right]-\operatorname{Pr}_{f^{n} \in \mathcal{P}^{n}}\left[\widehat{A}^{f^{n}}=1\right]\right| \geq \frac{1}{n^{e}}
$$

This contradicts the assumption that $\widehat{G}$ is a PRPG, and proves the claim.


[^0]:    Copyright © 1999 by Steven A. Myers

[^1]:    ${ }^{1}$ The notion of completely secure will formally be define in the sequel

[^2]:    Definition 2.1 (Pseudo-Random Number Generator - PRNG) Let $\boldsymbol{p}$ be a polynomial where $p(n)>n$ for every $n$. For each $n$ let $G^{n}:\{0,1\}^{n} \rightarrow\{0,1\}^{p(n)}$ be a function computable in a time bounded by a polynomial in $n$. Then $G=\left\{G^{n} \mid n \in \mathbb{N}\right\}$ is a pseudo-random number generator if, for all adversaries $A$, running in time bounded by a polynomial in $n$, and for all polynomials $q$, and for all sufficiently large $n$ :

[^3]:    ${ }^{1}$ Their results are actually slightly stronger, but we refer the reader to the original paper for the exact results

[^4]:    ${ }^{2}$ See [4] for a description of the birthday attack argument.

[^5]:    ${ }^{1}$ A PRFG such as G is constructed by taking a 1 -secure PRFG and modifying it so that the first $\delta$ fraction of the keys correspond to the zero function, and the remaining keys correspond to the same function as they did in the 1 -secure PRFG.

[^6]:    ${ }^{1}$ Of course we can't actually show that such generators exits, but we show that if PRPGs exist, then $H$ and $G$ exist.

[^7]:    ${ }^{2}$ In practice a PRNG $p$ would be used to set the first $c$ bits of $g_{k}^{n}(\overline{0})$. This would be done by setting the first $c$ bits of $g_{k}^{n}(\overline{0})$ to be equivalent be the first consecutive set of $c$ bits in $p(k)$ which were not all 1. This allows the first $c$ bits to be computed, and discards the need to store them, which would not be possible as it would require an exponential amount of storage.

[^8]:    ${ }^{1}$ Technically the original Isolation Lemma requires a commutative version of Lemma 5.6 and 5.7, but since the $\square$ operator is commutative the corresponding lemmas follow immediately.

[^9]:    ${ }^{2}$ In practice a PRFG $p$ would be used to set the first $c$ bits of $g_{k}^{n}\left(g_{k}^{n}(\tilde{0})\right)$. This would be done by setting the first $c$ bits of $g_{k}^{n}\left(g_{k}^{n}(\overline{0})\right.$ to be equivalent be the first consecutive set of $c$ bits in $p(k)$ which were not all 1. This allows the first $c$ bits to be computed, and discards the need to store them, which would not be possible as it would require an exponential amount of storage.

[^10]:    ${ }^{1}$ In practice a PRNG $p$ would be used to set the first $c$ bits of $g_{k}^{n}(\overline{0})$. This would be done by setting the first $c$ bits of $g_{k}^{n}(\overline{0})$ to be equivalent be the first consecutive set of $c$ bits in $p(k)$ which were not all 1. This allows the first $c$ bits to be computed, and discards the need to store them, which would not be possible as it would require an exponential amount of storage.

